

Weighted Approximation of the Renewal Spacing Processes

PHILIPPE BARBE

CREST and LSTA—Université Paris VI

In this paper, we provide a weighted approximation for the renewal spacing empirical and quantile processes. Some linear bounds for the empirical distribution and quantile functions are also given. © 1993 Academic Press, Inc.

1. INTRODUCTION

Let $\omega, \omega_1, \omega_2, \dots$, be a sequence of independent and identically distributed (i.i.d.) random variables (r.v.) with mean $0 < E\omega < \infty$ and common distribution function $F(x) := P(\omega/E\omega \leq x)$. We define the partial sums

$$S_i := \omega_1 + \omega_2 + \dots + \omega_i, \quad i = 1, 2, \dots, \quad (1.1)$$

and the renewal spacings

$$D_{i,n} := \omega_i/S_n, \quad 1 \leq i \leq n. \quad (1.2)$$

It is well known that if F is exponential, then the vector $\{S_i/S_n: 1 \leq i \leq n-1\}$ has the same distribution as the order statistics of a uniform sample of size $n-1$, and therefore, $\{D_{i,n}: 1 \leq i \leq n-1\}$ has the same distribution as the uniform spacings (see, e.g., [17]).

Now, define the quantile function

$$Q(s) := \inf\{x: F(x) \geq s\}$$

and the empirical distribution function of the normalized renewal spacings by

$$\mathbb{D}_n(s) := n^{-1} \sum_{i=1}^n \mathbb{I}\{nD_{i,n} \leq Q(s)\}, \quad (1.3)$$

Received January 22, 1992; revised July 27, 1992.

AMS 1991 subject classifications: 60F17, 60K05.

Keywords and phrases: renewal spacing, weighted approximation, Bahadur–Kiefer representation, linear bounds.

and the corresponding quantile process

$$\mathbb{D}_n^+(t) := \inf\{s: \mathbb{D}_n(s) \geq t\}.$$

Motivated by the fact that many test statistics based on renewal spacings may be expressed as a functional of \mathbb{D}_n or \mathbb{D}_n^+ when F is exponential, the limiting properties of the processes

$$a_n(s) := n^{1/2}(\mathbb{D}_n(s) - s)$$

and

$$b_n(s) := n^{1/2}(\mathbb{D}_n^+(s) - s)$$

have been subject of investigations over the years by many authors. If F is the exponential or some gamma distribution, these include Shorack [21], Rao and Sethuraman [19], Durbin [11], Beirlant [2], Einmahl and Van Zuijlen [13], Aly, Beirlant, and Horváth [1], Beirlant and Horváth [3], and Csörgő and Horváth (1986). Beirlant *et al.* [4] obtained a weighted approximation for b_n (see their Lemma 3.8) in the same spirit of M. Csörgő *et al.* (Cs–Cs–H–M) [6] or for a more general result by Einmahl and Mason [12]. However, the case of a general F is of interest since it enables us to get the limiting distribution of many statistics for testing whether a renewal process is a Poisson process or not under the null hypothesis, but also under a fixed alternative. Pyke [18] gives a weak convergence result for a_n and b_n .

The processes a_n and b_n are also related to the empirical process with estimated parameter. Since (1.1) is scale invariant, if F belongs to a parametric family with a constant coefficient of variation $E\omega/(\text{Var } \omega)^{1/2}$, \mathbb{D}_n may be viewed as the empirical d.f., where the mean or the standard deviation is estimated. For example, this situation occurs if F belongs to a family concentrated on $(0, \infty)$ and indexed by a scale parameter, like

$$\{F(x) = 1 - (k/x)^a, k > 0, a > 0, x \geq k\}$$

or if F is a Gamma distribution indexed by a scale parameter. Another example is given by

$$\{F(x) = \mathcal{N}(\theta, \theta^2): \theta > 0\}.$$

More general empirical processes with parameter estimated have been studied by Durbin [11], M. Csörgő and Révész [9], and M. Csörgő [5].

The Cs–Cs–H–M [6] weighted approximation has been a very useful tool to obtain the asymptotic distribution of a large variety of statistics based on the usual empirical distribution and quantile functions of a sample. See, for instance, the volume edited by Hahn, Mason, and

Weiner [14]. The aim of this paper is to obtain such approximations for the processes a_n and b_n and to show, as S. Csörgő and Mason [8] did for the bootstrap, that a “metatheorem” holds (see also [7, Section 17] for an implicit but less powerful metatheorem) in the sense that: all results proved using the weighted approximation and some linear bounds for the uniform empirical distribution or quantile processes hold for the renewal spacings (up to a change of the approximating process).

2. WEIGHTED APPROXIMATION AND LINEAR BOUNDS FOR THE RENEWAL SPACING PROCESS

We require the following assumptions:

A1. $0 < \text{Var } \omega < \infty$;

A2. The distribution function F of ω is twice differentiable, except maybe at a finite number of points;

A3. $\|x^2 f'(x)\| < \infty$, where $\|\cdot\|$ denotes the sup-norm and $f := F'$ is the density function of F ;

A4. $\|xf(x)F(x)^{-1/2}(1-F(x))^{-1/2}\| < \infty$.

In fact, A3 is generally satisfied under A1 if the density function f is bounded and ultimately concave or convex.

The limiting behavior of the processes a_n and b_n are different according to the expectation of ω is zero or not. If $E\omega \neq 0$, then there is no loss of generality in assuming

A5. $E\omega = 1$.

Finally, with B denoting a Brownian bridge, we define the Gaussian process

$$\Gamma_B(s) := B(s) - Q(s)f(Q(s)) \int_0^1 B(t) dQ(t).$$

If g is a function from $(0, 1)$ into \mathbb{R} , we define

$$\|g\|_n := \sup_{1/n \leq s \leq (n-1)/n} |g(s)|.$$

Our first theorem gives the weighted approximation of the process a_n :

THEOREM 2.1. *Assume A1–A5. If for some $v \in [0, 1/4]$ we have*

$$\int [F(x)(1-F(x))]^{(1/2)-v} dx < \infty, \quad (2.1)$$

then one can construct a sequence of Brownian bridges $(B_n)_{n \geq 1}$ such that

$$n^v \left\| \frac{a_n(s) - \Gamma_{B_n}(s)}{(s(1-s))^{(1/2)-v}} \right\|_n = O_p(1). \quad (2.2)$$

To obtain an analogous result for b_n , we shall establish a Bahadur–Kiefer type representation. For this, recall that a function g is regularly varying at $a \pm$ (with index $\rho \in \mathbb{R}$) iff $\lim_{\lambda \rightarrow 0} g(a \pm \lambda s)/g(a \pm s) = \lambda^\rho$ for any $\lambda > 0$. We need the following assumptions:

- B1. $\|F(x)^{1/2}(1-F(x))^{1/2}xf'(x)/f(x)\| < \infty$;
- B2. $\|xf(x)\| < \infty$;
- B3. $Q(u)f'(Q(u))/f(Q(u))$ and its derivative are regularly varying at $0+$ and $1-$.

If $\max_{1 \leq i \leq n} \omega_i$ and $\min_{1 \leq i \leq n} \omega_i$ normalized by a sequence of affine transformations admit a nondegenerate limiting distribution function and f' is ultimately concave or convex, then (see [20, pp. 85–93]) B1 and B3 hold.

THEOREM 2.2. *Assume A1–A5 and B1–B3. If for some $v \in [0, 1/4)$ the condition (2.1) holds and*

$$\|(\log|\log(F(1-F))|)^{1/2} (F(1-F))^{2v-1/2} (1+(f'/f))\| < \infty, \quad (2.3)$$

then one can construct a sequence $(\omega_i)_{i \geq 1}$ of the desired type on a probability space such that

$$n^{v+(1/2)} \left\| \frac{a_n(s) + b_n(s)}{(s(1-s))^{(1/2)-v}} \right\|_n = O_p(1). \quad (2.4)$$

Combining Theorems 2.1 and 2.2, we obtain the approximation for the quantile process:

THEOREM 2.3. *Under the assumptions of Theorem 2.2 and on the probability space of Theorem 2.2,*

$$n^v \left\| \frac{b_n(s) + \Gamma_{B_n}(s)}{(s(1-s))^{(1/2)-v}} \right\|_n = O_p(1). \quad (2.5)$$

Remark 2.1. The weighted approximation of Cs–Cs–H–M [6] and Mason and Van Zwet [16] suggest that (2.2) and (2.5) hold also for $1/4 < v < 1/2$ but for different Brownian motions. However, our proof (see Lemma 3.1) allows only $0 \leq v < 1/4$.

Our next goal is to provide linear bounds for \mathbb{D}_n and \mathbb{D}_n^- . For this, we denote

$$D_{(1),n} \leq \dots \leq D_{(n),n}$$

the order statistics of the renewal spacings $D_{1,n}, \dots, D_{n,n}$.

THEOREM 2.4. (i) *Under the assumptions of Theorem 2.1,*

$$\sup_{0 < s \leq 1} \frac{\mathbb{D}_n(s)}{s} + \sup_{0 \leq s < 1} \frac{1 - \mathbb{D}_n(s)}{1 - s} = O_p(1),$$

$$\sup_{F(nD_{(1),n}) \leq s \leq 1} \frac{s}{\mathbb{D}_n(s)} + \sup_{0 \leq s < F(nD_{(n),n})} \frac{1 - s}{1 - \mathbb{D}_n(s)} = O_p(1).$$

(ii) *Under the assumptions of Theorem 2.3, for any $\lambda > 0$, we have*

$$\sup_{\lambda/n \leq s \leq 1} \frac{\mathbb{D}_n^-(s)}{s} + \sup_{0 \leq s \leq 1 - (\lambda/n)} \frac{1 + \mathbb{D}_n^-(s)}{1 - s} = O_p(1)$$

$$\sup_{0 < s < 1} \frac{s}{\mathbb{D}_n(s)} + \sup_{0 < s < 1} \frac{1 - s}{1 - \mathbb{D}_n^-(s)} = O_p(1).$$

3. PROOFS

We first prove a lemma which will be used later.

LEMMA 3.1. *Under the assumptions of Theorem 2.1, we have*

$$n^v \left\| \frac{B_n(F((S_n/n) Q(s))) - B_n(s)}{(s(1-s))^{(1/2)-v}} \right\|_n = O_p(1), \quad (3.1)$$

Proof of Lemma 3.1. Our proof follows exactly Mason [15] and can also be derived equivalently following M. Csörgő and Horváth [8]: Let $\varepsilon > 0$ and denote $\Delta_n := (S_n - n)/n$. Since $n^{1/2}\Delta_n$ is asymptotically $\mathcal{N}(0, \sigma^2)$, we have

$$P \left(n^v \left\| \frac{B_n(F((S_n/n) Q(s))) - B_n(s)}{(s(1-s))^{(1/2)-v}} \right\|_n > u \right)$$

$$\leq P \left(n^v \left\| \frac{B_n(F((S_n/n) Q(s))) - B_n(s)}{(s(1-s))^{(1/2)-v}} \right\|_n > u; \Delta_n \leq \lambda/n^{1/2} \right) + \varepsilon \quad (3.2)$$

provided λ is large enough. From Mason [15] we have

$$P\left(n^v \max_{1 \leq i \leq n} \sup_{|s - (i/n)| \leq 1/(2n)} \frac{|B_n(s) - B_n(i/n)|}{(s(1-s))^{(1/2)-v}} \geq u\right) \leq \varepsilon$$

for u sufficiently large. Using the same technique, we obtain

$$\begin{aligned} & P\left(\sup_{|s - (i/n)| \leq 1/(2n)} n^{1/2} \frac{|B_n(F((S_n/n) Q(s))) - B(i/n)|}{i^{(1/2)-v}} \geq u; \Delta_n \leq \lambda/n^{1/2}\right) \\ & \leq P\left(\sup_{|t - (i/n)| \leq \mu_{i,n}} |B((i/n) + t) - B(i/n)| \geq u \mu_{i,n}^{1/2} \frac{i^{(1/2)-v}}{n^{1/2}} \mu_{i,n}^{-1/2}\right), \end{aligned} \quad (3.3)$$

where

$$\begin{aligned} \mu_{i,n} &:= \sup_{\substack{|s - (i/n)| \leq 1/(2n) \\ |t| \leq \lambda}} |F((1 + n^{-1/2}t) Q(s)) - (i/n)| \\ &\geq \sup_{|s - (i/n)| \leq 1/(2n)} |F((S_n/n) Q(s)) - (i/n)| \quad \text{if } |\Delta_n| \leq \lambda/n^{1/2}. \end{aligned}$$

As in Mason [15], we bound (3.3) by

$$A u^{-1} i^{v - (1/2)} n^{1/2} \mu_{i,n}^{1/2} \exp\left(-\frac{u^2 i^{1 - (2v)}}{8 n \mu_{i,n}}\right). \quad (3.4)$$

To bound $\mu_{i,n}$, expand $F((1 + t n^{-1/2}) Q(s))$ to obtain

$$\begin{aligned} & F((1 + t n^{-1/2}) Q(s)) - (i/n) \\ &= s - (i/n) + t n^{-1/2} Q(s) f(Q(s)) \\ & \quad + \frac{t^2}{2n} Q(s)^2 f'(Q((1 + \theta(n, s, t) n^{-1/2}) Q(s))) \end{aligned}$$

for some $\theta(n, s, t) \in [0, 1]$.

Since $|s - (i/n)| \leq 1/(2n)$ and $|t| \leq \lambda$, we obtain from A3 and A4 that for some constants c_1 and c_2 (independent of i and n) we have

$$\begin{aligned} \mu_{i,n} &\leq n^{-1} c_1 + n^{-1/2} \lambda \sup_{|s - (i/n)| \leq 1/(2n)} |Q(s) f(Q(s))| \\ &\leq n^{-1} c_1 + \lambda n^{-1} i^{1/2} \leq c_2 i^{1/2}/n. \end{aligned}$$

Hence, (3.4) is upper bounded by

$$A u^{-1} i^{v - (1/4)} \exp(-u^2 i^{(1/2) - 2v}/8),$$

which is the term of a convergent series for any fixed $u > 0$. Therefore (3.2) may be taken arbitrarily small, provided u is sufficiently large and ε is small enough, and this proves the lemma. ■

Proof of Theorem 2.1. We can assume that $\omega_i = Q(U_i)$ where U_1, U_2, \dots is a sequence of uniform r.v. over $(0, 1)$. Let us define the uniform empirical distribution function

$$\mathbb{F}_n(x) := n^{-1} \sum_{i=1}^n \mathbb{I}\{U_i \leq x\}. \quad (3.5)$$

Using (1.2) and (1.3), we can write \mathbb{D}_n as

$$\mathbb{D}_n(s) = \mathbb{F}_n\left(F\left(\frac{S_n}{n} Q(s)\right)\right). \quad (3.6)$$

We will first approximate $F((S_n/n) Q(s))$ and then use a weighted approximation for \mathbb{F}_n . Let $\Delta_n := (S_n - n)/n$. Then

$$\begin{aligned} F\left(\frac{S_n}{n} Q(s)\right) &= s + \Delta_n Q(s) f(Q(s)) \\ &\quad + (\Delta_n^2/2) Q(s)^2 f'[(1 + \theta_n(s) \Delta_n) Q(s)], \end{aligned} \quad (3.7)$$

where $\theta_n(s) \in [0, 1]$. Observe the following facts:

- (a) $\lim_{n \rightarrow \infty} \sup_{0 \leq s \leq 1} |\theta_n(s) \Delta_n| = 0$ a.s. under A1,
- (b) $\|Q^2 f'\| = \|x^2 f'(x)\| < \infty$ under A3,
- (c) $\left\| \frac{Q(s) f(Q(s))}{s^{1/2}(1-s)^{1/2}} \right\| = \left\| \frac{xf(x)}{F(x)^{1/2}(1-F(x))^{1/2}} \right\| < \infty$ under A4,

and by the central limit theorem

$$(d) \quad n^v \left\| \frac{\Delta_n}{(s(1-s))^{(1/2)-v}} \right\|_n = O_p(1).$$

These facts with (3.3) imply that

$$\left(F\left(\frac{S_n}{n} Q(s)\right) \left[1 - F\left(\frac{S_n}{n} Q(s)\right) \right] \right) / s(1-s) \quad (3.8)$$

is bounded away from 0 and ∞ in probability, for s on $[1/n, 1 - (1/n)]$ and, moreover

$$n^{v+(1/2)} \left\| \frac{F((S_n/n) Q(s)) - (s + \Delta_n Q(s) f(Q(s)))}{(s(1-s))^{(1/2)-v}} \right\|_n = O_p(1). \quad (3.9)$$

We may assume that the random variable U_1, U_2, \dots , are defined on the probability space where the weighted approximations by Cs-Cs-H-M [6] or Mason and Van Zwet [16] hold. Therefore, there exists a sequence of Brownian bridges $(B_n)_{n \geq 1}$, such that for any $0 \leq v < 1/4$

$$n^v \left\| \frac{n^{1/2}(\mathbb{F}_n(x) - x) - B_n(x)}{(x(1-x))^{(1/2)-v}} \right\|_n = O_p(1). \quad (3.10)$$

Combining (3.6)–(3.10), we get

$$n^v \left\| \frac{a_n(s) - n^{1/2} \Delta_n Q(s) f(Q(s)) - B_n(F(S_n/n) Q(s))}{(s(1-s))^{(1/2)-v}} \right\|_n = O_p(1). \quad (3.11)$$

From (3.11) and (3.1) we deduce

$$n^v \left\| \frac{a_n(s) - n^{1/2} \Delta_n Q(s) f(Q(s)) - B_n(s)}{(s(1-s))^{(1/2)-v}} \right\|_n = O_p(1). \quad (3.12)$$

To end the proof, note that

$$n^{1/2} \Delta_n = n^{1/2} \int_0^1 Q(t) d(\mathbb{F}_n(t) - t),$$

which after integration by parts becomes

$$n^{1/2} \Delta_n = - \int_0^1 n^{1/2} (\mathbb{F}_n(t) - t) dQ(t).$$

This, in conjunction with (3.10) gives

$$n^{1/2} \Delta_n = - \int_0^1 (B_n(t) + O_p(n^{-v})(t(1-t))^{(1/2)-v}) dQ(t), \quad (3.13)$$

where the $O_p(\cdot)$ is uniform in t . Assumption (2.1) ensures that

$$\int_0^1 (t(1-t))^{(1/2)-v} dQ(t) = \int [F(x)(1-F(x))]^{(1/2)-v} dx < \infty. \quad (3.14)$$

Assumption A4 implies that

$$\left\| \frac{Q(s) f(Q(s))}{(s(1-s))^{(1/2)-v}} \right\|_n \leq \left\| \frac{xf(x)}{(F(x)(1-F(x)))^{1/2}} \right\| < \infty. \quad (3.15)$$

Equations (3.13)–(3.15) imply

$$n^v \left\| Q(s) f(Q(s)) \frac{n^{1/2} \Delta_n + \int_0^1 B_n(t) dQ(t)}{(s(1-s))^{(1/2)-v}} \right\|_n = O_p(1), \quad (3.16)$$

which with (3.12) gives (2.2). ■

Proof of Theorem 2.2. We first define the inverse

$$\mathbb{G}_n(y) := \inf\{x: \mathbb{F}_n(x) \geq y\}$$

where \mathbb{F}_n is defined in (3.1). Then, an easy calculation shows that

$$\mathbb{D}_n^+(s) = F\left(\frac{n}{S_n} Q(\mathbb{G}_n(s))\right) = F((1 - \Delta_n n/S_n) Q(\mathbb{G}_n(s))). \quad (3.17)$$

Therefore,

$$\begin{aligned} \mathbb{D}_n^+(s) &= \mathbb{G}_n(s) - \Delta_n \frac{n}{S_n} Q(\mathbb{G}_n(s)) f(Q(\mathbb{G}_n(s))) \\ &\quad + (\Delta_n^2/2)(n/S_n)^2 Q(\mathbb{G}_n(s))^2 f' \left(\left(1 - \theta_n(s) \Delta_n \frac{n}{S_n}\right) Q(s) \right) \end{aligned} \quad (3.18)$$

with $0 \leq \theta_n(s) \leq 1$. Since we assume $E\omega = 1$, there exists a random n_0 a.s. finite such that for all $n \geq n_0$,

$$1 - \theta_n(s) \Delta_n \frac{n}{S_n} \in [1/2, 3/2]. \quad (3.19)$$

Moreover, the central limit theorem yields

$$n\Delta_n^2 = O_p(1) \quad (3.20)$$

and A3 implies

$$\|Q(u)^2 f'(\lambda Q(u))\| = \lambda^{-1} \|x^2 f'(x)\| < \infty. \quad (3.21)$$

From (3.18)–(3.19) we deduce

$$n^v \left\| \frac{n^{1/2} \mathbb{D}_n^+(s) - \mathbb{G}_n(s) + \Delta_n(n/S_n) Q(\mathbb{G}_n(s)) f(Q(\mathbb{G}_n(s)))}{(s(1-s))^{(1/2)-v}} \right\|_n = O_p(1). \quad (3.22)$$

We recall here Corollary 2.3 in Cs–Cs–H–M [6], which asserts

$$n^{v+(1/2)} \left\| \frac{\mathbb{F}_n(s) + \mathbb{G}_n(s) - 2s}{(s(1-s))^{(1/2)-v}} \right\|_n = O_p(1). \quad (3.23)$$

Combining (3.22) and (3.23) gives

$$\begin{aligned} n^{v+(1/2)} \left\| \frac{\mathbb{D}_n^-(s) + \mathbb{F}_n(s) + \Delta_n(n/S_n) Q(\mathbb{G}_n(s)) f(Q(\mathbb{G}_n(s)))}{(s(1-s))^{(1/2)-v}} \right\|_n \\ = O_p(1). \end{aligned} \quad (3.24)$$

Now, we use (3.10), (3.12), and (3.24) to obtain

$$\begin{aligned} n^{v+(1/2)} \left\| \frac{\mathbb{D}_n^-(s) + \mathbb{D}_n(s) - 2s - \Delta_n Q(s) f \circ Q(s) \\ + \Delta_n(n/S_n) Q(\mathbb{G}_n(s)) f(Q(\mathbb{G}_n(s)))}{(s(1-s))^{(1/2)-v}} \right\|_n \\ = O_p(1). \end{aligned} \quad (3.25)$$

It remains for us to study

$$\begin{aligned} (n/S_n) Q(\mathbb{G}_n(s)) f(Q(\mathbb{G}_n(s))) - Q(s) f(Q(s)) \\ = -\Delta_n(n/S_n) Q(\mathbb{G}_n(s)) f(Q(\mathbb{G}_n(s))) \\ + (\mathbb{G}_n(s) - s)(Qf(Q))'(s + \theta(n, s)(\mathbb{G}_n(s) - s)) \end{aligned}$$

for some $0 \leq \theta(n, s) \leq 1$. Observe that

$$n^v \Delta_n \left\| \frac{Q(\mathbb{G}_n(s)) f(Q(\mathbb{G}_n(s)))}{(s(1-s))^{(1/2)-s}} \right\|_n = O_p(1),$$

since $\|Q(s) f(Q(s))\| = \|xf(x)\| < \infty$ by B2 and $n^{1/2}\Delta_n = O_p(1)$ by A1 and the central limit theorem. Moreover, using (3.10), B3, and the linear bounds for \mathbb{G}_n given in Wellner [22], we have

$$\begin{aligned} n^v \left\| \frac{(\mathbb{G}_n(s) - s)(Qf(Q))'(s + \theta(n, s)(\mathbb{G}_n(s) - s))}{(s(1-s))^{(1/2)-v}} \right\|_n \\ \leq n^{v-(1/2)} \left\| \frac{n^{1/2}(\mathbb{G}_n(s) - s) - B_n(s)}{(s(1-s))^{(1/2)-v}} (Qf(Q))'(s + \theta(n, s)(\mathbb{G}_n(s) - s)) \right\|_n \\ + n^{v-(1/2)} \left\| \frac{B_n(s)}{(s(1-s))^{(1/2)-v}} ((Qf(Q))'(s + \theta(n, s)(\mathbb{G}_n(s) - s))) \right\|_n \end{aligned} \quad (3.26)$$

$$\begin{aligned} \leq n^{v-1/2} O_p(1) \|(Qf(Q))'(s)\|_n \\ + n^{v-(1/2)} \left\| \frac{B_n(s)}{(s(1-s))^{(1/2)-v}} (Qf(Q))'(s) \right\|_n O_p(1). \end{aligned} \quad (3.27)$$

Since by Darling and Erdős [10],

$$\left\| \frac{B_n(s)}{(s(1-s))^{1/2}} \right\|_n = O_p(\log \log n)^{1/2}, \quad (3.28)$$

both (3.27) and (3.28) are $O_p(1)$ provided

$$n^{-1/2} \|(Qf(Q))'(s)\|_n = O(1) \quad (3.29)$$

and

$$(\log \log n)^{1/2} n^{v-(1/2)} \|(s(1-s))^v (Qf(Q))'(s)\|_n = O(1). \quad (3.30)$$

Assumptions B1 and (2.3) imply that (3.29) and (3.30) hold. Hence (3.26) is $O_p(1)$ and with (3.25), this gives (2.4). ■

Proof of Theorem 2.3. Equations (2.2) and (2.4) give (2.5). ■

Proof of Theorem 2.5. (i) Using (3.6), we have

$$\mathbb{D}_n(s) = \frac{\mathbb{F}_n(F((S_n/n) Q(s)))}{F((S_n/n) Q(s))} F\left(\frac{S_n}{n} Q(s)\right).$$

Now, Wellner [22] provides the needed linear bounds on \mathbb{F}_n , and (3.4) gives the conclusion.

(ii) Using (3.18) and A3 we have

$$\left\| \mathbb{D}_n^+(s) - G_n(s) + \Delta_n \frac{n}{S_n} Qf(Q(G_n(s))) \right\|_n = O_p(n^{-1}).$$

Next, the linear bounds on G_n given in Wellner [22], combined with the fact that $Qf(Q)$ is regularly varying at $0+$ and $1-$ (under B3), leads easily to linear bounds for $Qf(Q(G_n(s)))$, and the results follow. ■

REFERENCES

- [1] ALY, E. E. A. A., BEIRLANT, J., AND HORVÁTH, L. (1984). Strong and weak approximation of k -spacings processes. *Zeit. Warsch. Verw. Geb.* **66** 461–484.
- [2] BEIRLANT, J. (1984). Strong approximation of the empirical and quantile processes of uniform spacings. In *Proc. Coll. Math. Soc. J. Bolyai 36, Limit Theorems in Probability and Statistics*, pp. 77–90.
- [3] BEIRLANT, J., AND HORVÁTH, L. (1984). Approximation of m -overlapping spacings processes. *Scand. J. Statist.* **11** 225–245.
- [4] BEIRLANT, J., DEHEUVELS, P., EINMAHL, J. H. J., AND MASON, D. M. (1991). Bahadur–Kieffer theorems for uniform spacings processes. *Theory Probab. Appl.* **36** 724–743.

- [5] CSÖRGÖ, M. (1983). *Quantile Processes with Statistical Applications*, CBMS-NSF Reg. Conf. Series 42. SIAM, Philadelphia.
- [6] CSÖRGÖ, M., CSÖRGÖ, S., HORVÁTH, L., AND MASON, D. M. (1986). Weighted empirical and quantile processes. *Ann. Probab.* **14** 31–85.
- [7] CSÖRGÖ, M., CSÖRGÖ, S., AND HORVÁTH, L. (1986). *An Asymptotic Theory for Empirical Reliability and Concentration Processes*, Lecture Notes in Statist., Vol. 33. Springer, Berlin.
- [8"] CSÖRGÖ, M., AND HORVÁTH, L. (1986). Approximations of weighted empirical and quantile processes. *Statist. Probab. Letters* **4** 275–280.
- [8] CSÖRGÖ, S., AND MASON, D. M. (1989). Bootstrapping empirical functions. *Ann. Statist.* **17** 1447–1471.
- [9] CSÖRGÖ, M., AND RÉVÉSZ, P. (1981). *Strong Approximation in Probability and Statistics*. Academic Press, New York.
- [10] DARLING, D. A., AND ERDŐS, P. (1956). A limit theorem for the maximum of normalized sums of independent random variables. *Duke Math. J.* **23** 143–145.
- [11] DURBIN, J. (1975). Kolmogorov–Smirnov test when parameter are estimated with applications to tests of exponentiality and tests on spacings. *Biometrika* **62** 5–22.
- [12] EINMAHL, U., AND MASON, D. M. (1992). Approximations to permutation and exchangeable processes. *J. Theory Probab.* **5** 101–126.
- [13] EINMAHL, J. H. J., AND VAN ZUIJLEN, M. C. A. (1988). Strong bounds for weighted empirical distribution functions based on uniform spacings. *Ann. Probab.* **16** 108–125.
- [14] HAHN, M. G., MASON, D. M., AND WEINER, D. C. (Eds.) (1991). *Sums, Trimmed Sums and Extremes*. Birkhäuser, Boston.
- [15] MASON, D. M. (1991). A note on weighted approximation to the uniform empirical and quantile processes. In *Sums, Trimmed Sums and Extremes* (M. G. Hahn, D. M. Mason, and D. C. Weiner, Eds.). Birkhäuser, Boston.
- [16] MASON, D. M., AND VAN ZWET, W. R. (1987). A refinement of the KMT inequality for the uniform empirical process. *Ann. Probab.* **15** 871–884.
- [17] PYKE, R. (1965). Spacings (with discussion). *J. Roy. Statist. Soc. B* **27** 395–449.
- [18] PYKE, R. (1970). Spacings revisited. In *Proc. 6th Berkeley Symp.* (L. Le Cam, J. Neymann, and E. L. Scotte, Eds.), Vol. 1, pp. 417–427. Univ. of California Press, Berkeley.
- [19] RAO, J. S., AND SETHURAMAN, J. (1975). Weak convergence of empirical distribution functions of random variables subject to perturbations and scale factors. *Ann. Statist.* **3** 299–313.
- [20] RESNICK, S. I. (1987). *Extreme Values, Regular Variation and Point Processes*. Springer, New York.
- [21] SHORACK, G. R. (1972). Convergence of quantile and spacings processes with applications. *Ann. Math. Statist.* **43** 1400–1411.
- [22] WELLNER, J. A. (1978). Limit theorem for the ratio of the empirical distribution function to the true distribution. *Z. Wahrsch. Verw. Geb.* **45** 73–88.