

# CONVERGENCE RATES IN THE COMPLETE MOMENT OF MOVING-AVERAGE PROCESSES

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(*Communicated by Gejza Wimmer*)

**ABSTRACT.** In this paper, we discuss precise asymptotics for a new kind of moment convergence of the moving-average process  $X_k = \sum_{i=-\infty}^{\infty} a_{i+k} \varepsilon_i$ ,  $k \geq 1$ , where  $\{\varepsilon_i : -\infty < i < \infty\}$  is a doubly infinite sequence of independent identically distributed random variables with mean zero and the finiteness of variance,  $\{a_i : -\infty < i < \infty\}$  is an absolutely summable sequence of real numbers, i.e.,  $\sum_{i=-\infty}^{\infty} |a_i| < \infty$ .

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## 1. Introduction and main results

We assume that  $\{\varepsilon_i : -\infty < i < \infty\}$  is a doubly infinite sequence of identically distributed variables. Let  $\{a_i : -\infty < i < \infty\}$  be an absolutely summable sequence of real numbers and  $X_k = \sum_{i=-\infty}^{\infty} a_{i+k} \varepsilon_i$ ,  $k \geq 1$ , also set  $S_n = \sum_{k=1}^n X_k$ .

Many limiting results have been obtained for moving-average processes  $\{X_k : k \geq 1\}$ . For example, Burton and Dehling [1] have obtained a large deviation principle for  $\{X_k : k \geq 1\}$  assuming  $E \exp t \varepsilon_1 < \infty$  for all  $t$ , Ibragimov [5] has established the central limit theorem for  $\{X_k : k \geq 1\}$ , Li, et al. [6]

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2010 Mathematics Subject Classification: Primary 60F15.

Keywords: Rosenthal type inequality, precise asymptotics, complete moment, moving-average processes.

This work was supported by the Natural Science Research Project of Ordinary Universities in Jiangsu Province, PR China. Grant No. 12KJB110003.

derived convergence rates of moderate deviations and the precise asymptotics in the law of the iterated logarithm, Dong, et al. [3] also obtained moderate deviation principles for moving-average processes of real stationary sequences.

On the other hand, Gut and Spătaru [4] proved the precise asymptotics of i.i.d random variables. One of their results is as follows.

**THEOREM 1.1.** *Suppose that  $\{Y_k : k \geq 1\}$  is a sequence of i.i.d random variables with  $EY_1 = 0$  and  $EY_1^2 = \sigma^2 < \infty$ . Then*

$$\lim_{\varepsilon \searrow 0} \varepsilon^2 \sum_{n=1}^{\infty} \frac{1}{n \log n} P\left(\left|\sum_{k=1}^n Y_k\right| \geq \varepsilon \sqrt{n \log \log n}\right) = \sigma^2.$$

Chow [2] discussed the complete moment convergence of i.i.d random variables. He got the following result:

**THEOREM 1.2.** *Let  $\{Y, Y_k : k \geq 1\}$  be a sequence of i.i.d random variables with  $EY_1 = 0$ . Suppose that  $p \geq 1$ ,  $\alpha > \frac{1}{2}$ ,  $p\alpha > 1$ ,  $E\{|Y|^p + |Y| \log(1 + |Y|)\} < \infty$ . Then for any  $\varepsilon > 0$ , we have*

$$\sum_{n=1}^{\infty} n^{p\alpha-2-\alpha} E\left\{\max_{j \leq n} \left|\sum_{k=1}^j Y_k\right| - \varepsilon n^\alpha\right\}_+ < \infty.$$

Recently, Liu and Lin [9] derived precise asymptotics for a new kind of complete moment convergence, one of their results is as follows:

**THEOREM 1.3.** *Suppose that  $\{X, X_n : n \geq 1\}$  is a sequence of i.i.d. random variables,  $0 < \delta \leq 1$ ,*

$$EX = 0, \quad EX^2 = \sigma^2 \quad \text{and} \quad EX^2 (\log^+ |X|)^\delta < \infty.$$

*Then we have*

$$\lim_{\varepsilon \searrow 0} \frac{1}{-\log \varepsilon} \sum_{n=1}^{\infty} \frac{1}{n^2} ES_n^2 I\{|S_n| \geq \varepsilon n\} = 2\sigma^2.$$

Inspired by the above results, we will extend this kind of results to moving-average processes. Now we state our results as follows.

**THEOREM 1.4.** *Suppose that  $\{\varepsilon, \varepsilon_n : n \geq 1\}$  is defined as above, and*

$$E\varepsilon = 0, \quad E\varepsilon^2 = \sigma^2, \quad E\varepsilon^3 < \infty. \quad (1.1)$$

*If  $\sum_{i=-\infty}^{\infty} a_i \neq 0$ , we have*

$$\lim_{\varepsilon \searrow 0} \frac{1}{-\log \varepsilon} \sum_{n=1}^{\infty} \frac{1}{n^2} ES_n^2 I\{|S_n| \geq \varepsilon n\} = 2\tau^2, \quad (1.2)$$

*where  $\tau^2 = \sigma^2 \left(\sum_{i=-\infty}^{\infty} a_i\right)^2$ .*

**THEOREM 1.5.** Suppose that  $\{\varepsilon, \varepsilon_n : n \geq 1\}$  is defined as above, and

$$E\varepsilon = 0, \quad E\varepsilon^2 = \sigma^2, \quad E\varepsilon^3 < \infty. \quad (1.3)$$

If  $\sum_{i=-\infty}^{\infty} a_i \neq 0$ , then, for  $1 \leq p < 2$ ,  $1 + \frac{p}{2} < r < 2$ ,  $0 \leq q < \frac{2(r-p)}{2-p}$ , we have

$$\begin{aligned} \lim_{\varepsilon \searrow 0} \varepsilon^{\frac{2(r-p)}{2-p}-q} \sum_{n=1}^{\infty} n^{\frac{r}{p}-2-\frac{q}{p}} E|S_n|^q I \left\{ |S_n| \geq \varepsilon n^{\frac{1}{p}} \right\} \\ = \frac{2p}{2r-2p-2q+pq} E|Z|^{\frac{2(r-p)}{2-p}}, \end{aligned} \quad (1.4)$$

where  $Z$  has a normal distribution with mean 0 and variance  $\tau^2 = \sigma^2 \left( \sum_{i=-\infty}^{\infty} a_i \right)^2$ .

## 2. Some lemmas

First, we give some lemmas which will be used in the proofs. Lemmas 3.1 and 3.2 are from Burton and Dehling [1], Yang [10] respectively.

**LEMMA 2.1.** Let  $\sum_{i=-\infty}^{\infty} a_i$  be an absolutely convergent series of real numbers with  $a = \sum_{i=-\infty}^{\infty} a_i$  and  $k \geq 1$ . Then

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{i=-\infty}^{\infty} \left| \sum_{j=i+1}^{i+n} a_j \right|^k = |a|^k.$$

**LEMMA 2.2.** Let  $\{\varepsilon_i : -\infty < i < \infty\}$  be a sequence of random variables with  $E\varepsilon_i = 0$ ,  $0 < E\varepsilon_i^2 < \infty$ , and  $\{X_i : i \geq 1\}$  is defined as above, where  $\{\varepsilon_i : -\infty < i < \infty\}$  is a sequence of real numbers with  $\sum_{i=-\infty}^{\infty} |a_i| < \infty$ . Then the moving-average process  $\{X_k\}$  fulfills the CLT, that is,

$$\frac{S_n}{\tau\sqrt{n}} \xrightarrow{\mathcal{D}} N(0, 1) \quad \text{where} \quad \tau = \sigma \sum_{i=-\infty}^{\infty} a_i.$$

**LEMMA 2.3 (Rosenthal inequality).** ([8]) Assume  $\{Y_i : i \geq 1\}$  is a sequence of independent random variables,  $EY_i = 0$ ,  $E|Y_i|^p < \infty$ , for some  $p \geq 2$  and every  $i \in \mathbb{R}$ . Then there exists  $C = C(p)$ , such that

$$E \left| \sum_{i \in \mathbb{R}} Y_i \right|^p \leq C \left\{ \sum_{i \in \mathbb{R}} E|Y_i|^p + \left( \sum_{i \in \mathbb{R}} |EY_i|^2 \right)^{\frac{p}{2}} \right\}.$$

Throughout the sequel,  $N$  represents standard normal variable and  $Z$  has a normal distribution with mean 0 and variance  $\tau^2 = \sigma^2 \left( \sum_{i=-\infty}^{\infty} a_i \right)^2$ .  $C$  will denote a positive constant although its value may change from one appearance to the next and let  $[x]$  indicate the maximum integer not larger than  $x$ . Without loss of generality, we assume  $\tau = 1$  in the sequel.

### 3. Proof of Theorem 1.4

In this section, we set  $a(\varepsilon) = [\varepsilon^{-2}]$  for  $0 < \varepsilon < 1$ . Note that

$$\sum_{n=1}^{\infty} \frac{1}{n^2} E S_n^2 I \{|S_n| \geq \varepsilon n\} = \varepsilon^2 \sum_{n=1}^{\infty} P(|S_n| \geq \varepsilon n) + \sum_{n=1}^{\infty} \frac{1}{n^2} \int_{\varepsilon n}^{\infty} 2x P(|S_n| \geq x) dx.$$

When taking  $p = 1$  and  $r = 2$  in Li [7], we have

$$\lim_{\varepsilon \searrow 0} \varepsilon^2 \sum_{n=1}^{\infty} P(|S_n| \geq \varepsilon n) = 1.$$

Thus, in order to prove (1.2), it suffices to show

$$\lim_{\varepsilon \searrow 0} \frac{1}{-\log \varepsilon} \sum_{n=1}^{\infty} \frac{1}{n^2} \int_{\varepsilon n}^{\infty} 2x P(|S_n| \geq x) dx = 2. \quad (3.1)$$

This will be proved by the following propositions.

**PROPOSITION 3.1.** *One has*

$$\lim_{\varepsilon \searrow 0} \frac{1}{-\log \varepsilon} \sum_{n=1}^{\infty} \frac{1}{n^2} \int_{\varepsilon n}^{\infty} 2x P(|N| \geq x/\sqrt{n}) dx = 2.$$

*Proof.* See the proof of [9, Proposition 3.1]. □

**PROPOSITION 3.2.** *One has*

$$\lim_{\varepsilon \searrow 0} \frac{1}{-\log \varepsilon} \sum_{n=1}^{a(\varepsilon)} \frac{1}{n^2} \left| \int_{\varepsilon n}^{\infty} 2x P(|S_n| \geq x) dx - \int_{\varepsilon n}^{\infty} 2x P(|N| \geq x/\sqrt{n}) dx \right| = 0. \quad (3.2)$$

*Proof.* Denote

$$\Delta_n = \sup_x \left| P\left(\frac{|S_n|}{\sqrt{n}} \geq x\right) - P(|N| \geq x) \right|,$$

it follows from Lemma 2 that  $\Delta_n \rightarrow 0$  as  $n \rightarrow \infty$ . Obviously,

$$\begin{aligned}
 & \left| \sum_{n=1}^{a(\varepsilon)} \frac{1}{n^2} \left| \int_{\varepsilon n}^{\infty} 2xP(|S_n| \geq x) \, dx - \int_{\varepsilon n}^{\infty} 2xP(|N| \geq x/\sqrt{n}) \, dx \right| \right| \\
 &= \left| \sum_{n=1}^{a(\varepsilon)} \left| \int_0^{\infty} 2(x+\varepsilon)P(|S_n| \geq n(x+\varepsilon)) \, dx - \int_0^{\infty} 2(x+\varepsilon)P(|N| \geq \sqrt{n}(x+\varepsilon)) \, dx \right| \right| \\
 &\leq \sum_{n=1}^{a(\varepsilon)} \frac{1}{n} \int_0^{\infty} 2n(x+\varepsilon) |P(|S_n| \geq n(x+\varepsilon)) - P(|N| \geq \sqrt{n}(x+\varepsilon))| \, dx \\
 &\leq \sum_{n=1}^{a(\varepsilon)} \frac{1}{n} (\Delta_{n1} + \Delta_{n2} + \Delta_{n3}),
 \end{aligned}$$

where

$$\begin{aligned}
 \Delta_{n1} &= \int_0^{\frac{1}{\sqrt{n}\Delta_n^{1/4}}} 2n(x+\varepsilon) |P(|S_n| \geq n(x+\varepsilon)) - P(|N| \geq \sqrt{n}(x+\varepsilon))| \, dx; \\
 \Delta_{n2} &= \int_0^{\infty} 2n(x+\varepsilon)P(|S_n| \geq n(x+\varepsilon)) \, dx; \\
 \Delta_{n3} &= \int_{\frac{1}{\sqrt{n}\Delta_n^{1/4}}}^{\infty} 2n(x+\varepsilon)P(|N| \geq \sqrt{n}(x+\varepsilon)) \, dx.
 \end{aligned}$$

Since  $n \leq b(\varepsilon)$  implies  $\varepsilon\sqrt{n} \leq 1$ , we have

$$\begin{aligned}
 \Delta_{n1} &\leq \int_0^{\frac{1}{\sqrt{n}\Delta_n^{1/4}}} 2n(x+\varepsilon)\Delta_n \, dx \leq n\Delta_n \left( \frac{1}{\sqrt{n}\Delta_n^{1/4}} + \varepsilon \right)^2 \\
 &\leq \left( \Delta_n^{1/4} + \Delta_n^{1/2} \right)^2 \rightarrow 0, \quad \text{as } n \rightarrow \infty.
 \end{aligned} \tag{3.3}$$

Next, observe that

$$\sum_{k=1}^n X_k = \sum_{i=-\infty}^{\infty} \sum_{k=1}^n a_{k+i\varepsilon_i}.$$

Set  $a_{ni} = \sum_{k=1}^n a_{k+i}$ . Then

$$\sum_{k=1}^n X_k = \sum_{i=-\infty}^{\infty} a_{ni} \varepsilon_i =: \sum_{i=-\infty}^{\infty} Y_i.$$

From Lemma 1, we can assume, without loss of generality, that

$$\sum_{i=-\infty}^{\infty} |a_{ni}|^k \leq Cn, \quad n \geq 1, \quad k \geq 1 \quad \text{and} \quad \sum_{i=-\infty}^{\infty} |a_i| \leq 1.$$

And then, by Lemma 3 (rosenthal's inequality), we get

$$E|S_n|^3 \leq C \left( \sum_{i=-\infty}^{\infty} E|a_{ni} \varepsilon_i|^3 + \left( \sum_{i=-\infty}^{\infty} E|a_{ni} \varepsilon_i|^2 \right)^{\frac{3}{2}} \right) \leq Cn^{\frac{3}{2}}. \quad (3.4)$$

Thus, via Markov's inequality, we have

$$\begin{aligned} \Delta_{n2} &\leq C \int_{\frac{1}{\sqrt{n}\Delta_n^{1/4}}}^{\infty} n(x+\varepsilon) \frac{n^{\frac{3}{2}}}{n^3(x+\varepsilon)^3} dx \\ &\leq C \int_{\frac{1}{\sqrt{n}\Delta_n^{1/4}}}^{\infty} \frac{1}{\sqrt{n}(x+\varepsilon)^2} dx \\ &\leq C\Delta_n^{\frac{1}{4}} \rightarrow 0, \quad \text{as } n \rightarrow \infty. \end{aligned} \quad (3.5)$$

Now, we estimate  $\Delta_{n3}$ . By Markov's inequality, we have

$$\Delta_{n3} \leq Cn \int_{\frac{1}{\sqrt{n}\Delta_n^{1/4}}}^{\infty} (x+\varepsilon) \frac{1}{n^2(x+\varepsilon)^4} dx \leq C\Delta_n^{\frac{1}{2}} \rightarrow 0, \quad \text{as } n \rightarrow \infty. \quad (3.6)$$

From (3.3), (3.5), (3.6), the proof of this proposition is derived.  $\square$

**PROPOSITION 3.3.** *One has*

$$\lim_{\varepsilon \searrow 0} \frac{1}{-\log \varepsilon} \sum_{n=a(\varepsilon)+1}^{\infty} \frac{1}{n^2} \left| \int_{\varepsilon n}^{\infty} 2xP(|S_n| \geq x) dx - \int_{\varepsilon n}^{\infty} 2xP(|N| \geq x/\sqrt{n}) dx \right| = 0. \quad (3.7)$$

P r o o f.

$$\begin{aligned}
 & \lim_{\varepsilon \searrow 0} \frac{1}{-\log \varepsilon} \sum_{n=a(\varepsilon)+1}^{\infty} \frac{1}{n^2} \left| \int_{\varepsilon n}^{\infty} 2xP(|S_n| \geq x) \, dx - \int_{\varepsilon n}^{\infty} 2xP(|N| \geq x/\sqrt{n}) \, dx \right| \\
 & \leq \lim_{\varepsilon \searrow 0} \frac{1}{-\log \varepsilon} \sum_{n=a(\varepsilon)+1}^{\infty} \int_0^{\infty} 2(x+\varepsilon)P(|S_n| \geq n(x+\varepsilon)) \, dx \\
 & \quad + \frac{1}{-\log \varepsilon} \sum_{n=b(\varepsilon)+1}^{\infty} \int_0^{\infty} 2(x+\varepsilon)P(|N| \geq \sqrt{n}(x+\varepsilon)) \, dx \\
 & =: I_1 + I_2.
 \end{aligned}$$

For  $I_1$ , by (3.2) and Markov's inequality, we have

$$\begin{aligned}
 I_1 & \leq \lim_{\varepsilon \searrow 0} \frac{1}{-\log \varepsilon} \sum_{n=a(\varepsilon)+1}^{\infty} \frac{1}{n^2} \int_{\varepsilon n}^{\infty} \frac{n^{\frac{3}{2}}}{x^2} \, dx \\
 & \leq \frac{C}{-\log \varepsilon} \sum_{n=a(\varepsilon)+1}^{\infty} \frac{1}{\varepsilon n^{\frac{3}{2}}} \\
 & \leq \frac{C}{-\log \varepsilon} \frac{1}{\sqrt{a(\varepsilon)+1}} \rightarrow 0, \quad \text{as } \varepsilon \searrow 0.
 \end{aligned}$$

For  $I_2$ , in view of Markov's inequality, we have

$$\begin{aligned}
 I_2 & \leq \lim_{\varepsilon \searrow 0} \frac{C}{-\log \varepsilon} \sum_{n=a(\varepsilon)+1}^{\infty} \frac{1}{n^2} \int_{\varepsilon n}^{\infty} \frac{n^2}{x^3} \, dx \\
 & \leq \frac{C}{-\log \varepsilon} \sum_{n=a(\varepsilon)+1}^{\infty} \frac{1}{\varepsilon^2 n^2} \\
 & \leq \frac{C}{-\varepsilon^2 \log \varepsilon} \frac{1}{a(\varepsilon)+1} \rightarrow 0, \quad \text{as } \varepsilon \searrow 0.
 \end{aligned}$$

Thus, we complete the proof of the proposition.  $\square$

Now, Theorem 1.4 follows from the propositions.

#### 4. Proof of Theorem 1.5

In this section, we set  $b(\varepsilon) = [\varepsilon^{\frac{-2p}{2-p}}]$  for  $0 < \varepsilon < 1$ , and  $1 \leq p < 2$ . When taking  $q = 0$ , we have

$$\lim_{\varepsilon \searrow 0} \varepsilon^{\frac{2(r-p)}{2-p}} \sum_{n \geq 1} n^{\frac{r}{p}-2} P(|S_n| \geq \varepsilon n^{\frac{1}{p}}) = \frac{p}{r-p} E|Z|^{\frac{2(r-p)}{2-p}}.$$

This is the result of Li [7]. Thus we only discuss the case  $0 < p < 2$ . Note that

$$\begin{aligned} & \sum_{n=1}^{\infty} n^{\frac{r}{p}-2-\frac{q}{p}} E|S_n|^q I\left\{|S_n| \geq \varepsilon n^{\frac{1}{p}}\right\} \\ &= \varepsilon^q \sum_{n=1}^{\infty} n^{\frac{r}{p}-2} P\left(|S_n| \geq \varepsilon n^{\frac{1}{p}}\right) + \sum_{n=1}^{\infty} n^{\frac{r}{p}-2-\frac{q}{p}} \int_{\varepsilon n^{\frac{1}{p}}}^{\infty} qx^{q-1} P(|S_n| \geq x) dx. \end{aligned}$$

Via the result of Li [7]. Thus, in order to prove (1.4), it suffices to show

$$\begin{aligned} & \lim_{\varepsilon \searrow 0} \varepsilon^{\frac{2(r-p)}{2-p}-q} \sum_{n=1}^{\infty} n^{\frac{r}{p}-2-\frac{q}{p}} \int_{\varepsilon n^{\frac{1}{p}}}^{\infty} qx^{q-1} P(|S_n| \geq x) dx \\ &= \frac{pq(2-p)}{(r-p)(2r-2p-2q+pq)} E|N|^{\frac{2(r-p)}{2-p}}. \end{aligned} \quad (4.1)$$

This will be proved by the following propositions.

**PROPOSITION 4.1.** *One has*

$$\begin{aligned} & \lim_{\varepsilon \searrow 0} \varepsilon^{\frac{2(r-p)}{2-p}-q} \sum_{n=1}^{\infty} n^{\frac{r}{p}-2-\frac{q}{p}} \int_{\varepsilon n^{\frac{1}{p}}}^{\infty} qx^{q-1} P\left(|N| \geq \frac{x}{\sqrt{n}}\right) dx \\ &= \frac{2p}{2r-2p-2q+pq} E|Z|^{\frac{2(r-p)}{2-p}} E|N|^{\frac{2(r-p)}{2-p}}. \end{aligned}$$

**Proof.** Via the change of variation, we have

$$\begin{aligned} & \lim_{\varepsilon \searrow 0} \varepsilon^{\frac{2(r-p)}{2-p}-q} \sum_{n=1}^{\infty} n^{\frac{r}{p}-2-\frac{q}{p}} \int_{\varepsilon n^{\frac{1}{p}}}^{\infty} qx^{q-1} P\left(|N| \geq \frac{x}{\sqrt{n}}\right) dx \\ &= \lim_{\varepsilon \searrow 0} \varepsilon^{\frac{2(r-p)}{2-p}-q} \sum_{n=1}^{\infty} n^{\frac{r}{p}-2-\frac{q}{p}+\frac{q}{2}} \int_{\varepsilon n^{\frac{1}{p}-\frac{1}{2}}}^{\infty} qx^{q-1} P(|N| \geq t) dt \end{aligned}$$



$$\begin{aligned}
 &= \lim_{\varepsilon \searrow 0} \varepsilon^{\frac{2(r-p)}{2-p}-q} \int_1^\infty x^{\frac{r}{p}-2-\frac{q}{p}+\frac{q}{2}} dx \int_{\varepsilon x^{\frac{1}{p}-\frac{1}{2}}}^\infty qx^{q-1} P(|N| \geq t) dt \\
 &= \frac{2pq}{2-p} \int_0^\infty y^{\frac{2(r-p)}{2-p}-q-1} dy \int_y^\infty t^{q-1} P(|N| \geq t) dt \\
 &= \frac{2p}{2-p} \frac{q}{\frac{2(r-p)}{2-p}-q} \int_0^\infty t^{\frac{2(r-p)}{2-p}-1} P(|N| \geq t) dt \\
 &= \frac{2p}{2r-2p-2q+pq} E|Z|^{\frac{2(r-p)}{2-p}} E|N|^{\frac{2(r-p)}{2-p}}.
 \end{aligned}$$

This completes the proof of Proposition 4.1.  $\square$

**PROPOSITION 4.2.** *For  $M > 1$ , one has*

$$\begin{aligned}
 \lim_{\varepsilon \searrow 0} \varepsilon^{\frac{2(r-p)}{2-p}-q} \sum_{n=1}^{Mb(\varepsilon)} n^{\frac{r}{p}-2-\frac{q}{p}} \left| \int_{\varepsilon n^{\frac{1}{p}}}^\infty qx^{q-1} P(|S_n| \geq x) dx \right. \\
 \left. - \int_{\varepsilon n^{\frac{1}{p}}}^\infty qx^{q-1} P(|N| \geq x/\sqrt{n}) dx \right| = 0.
 \end{aligned}$$

**Proof.** It is easy to see that

$$\begin{aligned}
 &\sum_{n=1}^{Mb(\varepsilon)} n^{\frac{r}{p}-2-\frac{q}{p}} \left| \int_{\varepsilon n^{\frac{1}{p}}}^\infty qx^{q-1} P(|S_n| \geq x) dx - \int_{\varepsilon n^{\frac{1}{p}}}^\infty qx^{q-1} p(|N| \geq x/\sqrt{n}) dx \right| \\
 &\leq \sum_{n=1}^{Mb(\varepsilon)} n^{\frac{r}{p}-2} \int_0^\infty q(x+\varepsilon)^{q-1} \left| p(|S_n| \geq (x+\varepsilon)n^{\frac{1}{p}}) \right. \\
 &\quad \left. - p(|N| \geq (x+\varepsilon)n^{\frac{1}{p}-\frac{1}{2}}) \right| dx \\
 &\leq \sum_{n=1}^{Mb(\varepsilon)} n^{\frac{r}{p}-2-\frac{q(2-p)}{2p}} (\Delta'_{n1} + \Delta'_{n2}),
 \end{aligned}$$

where

$$\begin{aligned}\Delta'_{n1} &= n^{\frac{q(2-p)}{2p}} \int_0^{\frac{1}{n^{\frac{2-p}{2p}} \Delta_n^{1/2q}}} q(x+\varepsilon)^{q-1} \left| p \left( |S_n| \geq (x+\varepsilon)n^{\frac{1}{p}} \right) \right. \\ &\quad \left. - p \left( |N| \geq (x+\varepsilon)n^{\frac{1}{p}-\frac{1}{2}} \right) \right| dx, \\ \Delta'_{n2} &= n^{\frac{q(2-p)}{2p}} \int_{\frac{1}{n^{\frac{2-p}{2p}} \Delta_n^{1/2q}}}^{\infty} q(x+\varepsilon)^{q-1} \left| P \left( |S_n| \geq (x+\varepsilon)n^{\frac{1}{p}} \right) \right. \\ &\quad \left. - p \left( |N| \geq (x+\varepsilon)n^{\frac{1}{p}-\frac{1}{2}} \right) \right| dx.\end{aligned}$$

Since  $n \leq Mb(\varepsilon)$  implies  $\varepsilon n^{\frac{2-p}{2p}} \leq M^{\frac{2-p}{2p}}$ , one can easily obtain that

$$\Delta'_{n1} \leq \Delta_n n^{\frac{q(2-p)}{2p}} \left( \frac{1}{n^{\frac{2-p}{2p}} \Delta_n^{1/2q}} + \varepsilon \right)^q \leq \left( \Delta_n^{\frac{1}{2q}} + M^{\frac{2-p}{2p}} \Delta_n^{\frac{1}{q}} \right)^p. \quad (4.2)$$

By Markov's inequality, we have

$$\Delta'_{n2} \leq C n^{\frac{q(2-p)}{2p}} \int_{\frac{1}{n^{\frac{2-p}{2p}} \Delta_n^{1/2q}}}^{\infty} \frac{1}{(x+\varepsilon)^{3-q} n^{\frac{2}{p}}} dx \leq C \Delta_n^{\frac{1}{q}-\frac{1}{2}}. \quad (4.3)$$

Denote  $\Delta'_n = \Delta'_{n1} + \Delta'_{n2}$ , it follows that

$$m^{-\frac{r}{p}+1+\frac{q(2-p)}{2p}} \sum_{n=1}^m n^{\frac{r}{p}-2-\frac{q(2-p)}{2p}} \Delta'_n \rightarrow 0, \quad \text{as } m \rightarrow \infty.$$

We have

$$\begin{aligned}\lim_{\varepsilon \searrow 0} \varepsilon^{\frac{2(r-p)}{2-p}-q} \sum_{n=1}^{Mb(\varepsilon)} n^{\frac{r}{p}-2-\frac{q}{p}} &\left| \int_{\varepsilon n^{\frac{1}{p}}}^{\infty} q x^{q-1} P(|S_n| \geq x) dx \right. \\ &\left. - \int_{\varepsilon n^{\frac{1}{p}}}^{\infty} q x^{q-1} P(|N| \geq x/\sqrt{n}) dx \right|\end{aligned}$$

$$\begin{aligned}
 &= \lim_{\varepsilon \searrow 0} \varepsilon^{\frac{2(r-p)}{2-p}-q} \sum_{n=1}^{Mb(\varepsilon)} n^{\frac{r}{p}-2-\frac{q(2-p)}{2p}} \Delta'_n \\
 &\leq \lim_{\varepsilon \searrow 0} \varepsilon^{\frac{2(r-p)}{2-p}-q} [Mb(\varepsilon)]^{\frac{r}{p}-1-\frac{q(2-p)}{2p}} [Mb(\varepsilon)]^{-\frac{r}{p}+1+\frac{q(2-p)}{2p}} \sum_{n=1}^{Mb(\varepsilon)} n^{\frac{r}{p}-2-\frac{q(2-p)}{2p}} \Delta'_n = 0.
 \end{aligned}$$

Then this proposition is proved.  $\square$

**PROPOSITION 4.3.** *For  $M > 1$ , one has*

$$\begin{aligned}
 \lim_{M \rightarrow \infty} \lim_{\varepsilon \searrow 0} \varepsilon^{\frac{2(r-p)}{2-p}-q} \sum_{n > Mb(\varepsilon)} n^{\frac{r}{p}-2-\frac{q}{p}} \left| \int_{\varepsilon n^{\frac{1}{p}}}^{\infty} qx^{q-1} P(|S_n| \geq x) dx \right. \\
 \left. - \int_{\varepsilon n^{\frac{1}{p}}}^{\infty} qx^{q-1} P(|N| \geq x/\sqrt{n}) dx \right| = 0.
 \end{aligned}$$

*Proof.* Observe that

$$E|S_n|^2 \leq CE \left( \sum_{i=-\infty}^{\infty} a_{ni}^2 \varepsilon_i^2 \right) \leq Cn.$$

Thus, by Markov's inequality, we have

$$\begin{aligned}
 &\varepsilon^{\frac{2(r-p)}{2-p}-q} \sum_{n > Mb(\varepsilon)} n^{\frac{r}{p}-2-\frac{q}{p}} \left| \int_{\varepsilon n^{\frac{1}{p}}}^{\infty} qx^{q-1} P(|S_n| \geq x) dx \right. \\
 &\quad \left. - \int_{\varepsilon n^{\frac{1}{p}}}^{\infty} qx^{q-1} P(|N| \geq x/\sqrt{n}) dx \right| \\
 &\leq C\varepsilon^{\frac{2(r-p)}{2-p}-q} \sum_{n > Mb(\varepsilon)} n^{\frac{r}{p}-2-\frac{q}{p}} \int_{\varepsilon n^{\frac{1}{p}}}^{\infty} \frac{n}{x^{3-q}} dx \\
 &\leq C\varepsilon^{\frac{2(r-p)}{2-p}-2} \sum_{n > Mb(\varepsilon)} n^{\frac{r}{p}-\frac{2}{p}-1} \\
 &\leq CM^{\frac{r-2}{p}} \rightarrow 0, \quad \text{as } M \rightarrow \infty.
 \end{aligned}$$

We complete the proof of this proposition.  $\square$

Our main result now follows from the propositions via triangle inequality.

**Acknowledgement.** The authors would like to thank professor Zhengyan Lin of Zhejiang University for his help.

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Received 16. 5. 2010

Accepted 12. 7. 2011

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