

BISPECTRAL ANALYSIS OF SCALP ELECTROENCEPHALOGRAMS: QUADRATIC PHASECOUPLING PHENOMENON IN DETECTING BRAIN TUMOR

Salai Selvam, V. and S. Shenbagadevi

Department of Electronics and Communication Engineering,
Faculty of Information and Communication Engineering, Anna University, Chennai, 600 025, India

Received 2012-11-20, Revised 2013-02-16; Accepted 2013-04-22

ABSTRACT

Since the power spectral analysis of a non-Gaussian process generated by a nonlinear mechanism e.g., EEG, does not provide much information on the underlying nonlinear dynamics due to the lack of phase information, the higher-order statistics such as the bispectra are used to better understand the underlying nonlinear dynamics e.g., the quadratic phase coupling phenomena. The quadratic phase couplings have been observed in the EEG by the researchers over a decade for many diagnostic applications such as epilepsy, sleep, mental states. This study discusses the use of bispectral analysis of the EEG recorded from the posterior region of the head of the brain tumor patient in quantifying the quadratic phase couplings to indicate the presence of the tumor. The Bicoherence Index (BCI) or simply the Bicoherence (BIC) has been used for the purpose. Self-couplings (around 27-52%) in the [8-13] Hz (alpha) band and phase couplings (around 23-42%) in the [1-8] Hz (delta-theta) band have been observed for the normal subjects while only self-couplings (around <6.5% and around 40-53%) have been seen in both bands for the brain tumor patients. Significant lowering of coupling strengths (from 38.15% ($\pm 12.76\%$) to 3.51% ($\pm 3.28\%$)) in the alpha band and mild increase of them (from 32.76% ($\pm 18.73\%$) to 45.49% ($\pm 17.49\%$)) in the delta band have been observed for the brain tumor patients. The Power Ratio Index (PRI) based on the power spectrum is only statistically inferior ($p > 0.05$) to the BIC in discriminating the brain tumor case from the normal one.

Keywords: Quadratic Nonlinearity, Bispectrum, Bicoherence Index, Phase Synchrony, Brain Tumor, PRI

1. INTRODUCTION

Diagnosis and subsequent (early) treatment are either missed or delayed in 69% of the brain tumor cases due to the fact that the most of the brain tumor symptoms are highly misleading and around 26% of these cases suffer a delay of more than a year before proper diagnosis (MFBTRI, 2013). Once the brain tumor symptoms are found, the advanced neuroimaging techniques such as MRI and CT or biopsy are not immediately suggested due to the following facts: They are either costly or invasive or do involve risks like

hazardous radiation, especially in case of children, pregnant women and patients with implant devices (Black, 2010). Since an early treatment increases the survival rate, a better method that does not involve much cost, risks or complexity is required to reduce the delay in the diagnosis of the brain tumors (Black, 2010). One such option is the use of the scalp Electroencephalograms (EEGs) (Fattal-Valevski *et al.*, 2012). Qualitative investigations on the use of scalp EEG for the diagnosis of brain tumors are numerous (Accolla *et al.*, 2011; Bagchi *et al.*, 1961; Hartman and Lesser, 2012; Walter, 1936; Watemberg *et al.*, 2002). However the quantitative

Corresponding Author: V. Salai Selvam, Department of Electronics and Communication Engineering

Faculty of Information and Communication Engineering, Anna University, Chennai, 600 025, India

Tel: +914427689029, Fax: +914427689364

works such as automated classification or localization using the quantitative (scalp) EEG (qEEG) features are very few (Chetty and Venayagamoorthy, 2002; Hahl *et al.*, 2000; Karamah and Dahleh, 2000; Murugesan and Sukanesh, 2009; Nagata *et al.*, 1985; Selvam and Shenbagadevi, 2011; Silipo *et al.*, 1999). Some of these literatures (Selvam and Shenbagadevi, 2011; Silipo *et al.*, 1999) are interesting and they suggest the use of nonlinear analysis of the scalp EEG for the purpose.

The power spectrum does not provide much information about the non-Gaussian processes and the processes generated by the nonlinear mechanisms due to the lack of phase information (Nikias and Raghuveer, 1987). The conditions of non-Gaussianity and nonlinearity are often met in many practical applications, e.g., the biological signals and the biological systems generating them. Hence the analysis of such processes requires the use of the higher-order spectra such as the bispectra as they provide more information such as the deviation from Gaussianity and linearity, phase couplings. In this study the application of bispectrum, which unveils the quadratic phase coupling phenomena arising from the second-order nonlinearity, has been considered.

The bispectrum has been extensively used to analyze the EEG time series to discriminate various pathological conditions, mental tasks, physical (sleepy, anesthetic) states (Hosseini *et al.*, 2010; Saikia and Hazarika, 2011; Swarnkar *et al.*, 2010; Venkatakrishnan *et al.*, 2011; Yufune *et al.*, 2011). It has been shown in these papers that the quadratic phase couplings in EEG do change in accordance with the physical and physiological conditions of the brain. As the self-emitting oscillations of the complicated neural network of the brain (Buzsaki and Draguhn, 2004) can be considered non-Gaussian (Nikias and Raghuveer, 1987), the application bispectrum to EEG is quite intuitive. As the presence of a brain tumor affects the brain both physically and physiologically, the expectation of changes in the (nonlinear) dynamics of the brain and subsequently the use of bispectrum for quantifying them are intuitively understandable.

In this study, the bispectral analysis of the scalp EEG recorded from the posterior (parietal and occipital) regions of the heads of both the brain tumor patients and normal subjects has been presented. The Bicoherence Index (BCI) or simply Bicoherence (BIC) is used to quantify the quadratic phase couplings in the EEG records. The results obtained have been tested for their statistical significance in characterizing the presence of the brain tumor. A power spectral feature, namely the Power Ratio Index (PRI) has also been analyzed for comparison.

2. MATERIALS AND METHODS

2.1. Acquisition of Data

Nineteen-channel Common Average Reference (CAR) montage EEG records, namely Fz-REF, Cz- REF, Pz-REF, FP1-REF, FP2-REF, F3-REF, F4-REF, C3-REF, C4-REF, P3-REF, P4-REF, O1-REF, O2-REF, F7-REF, F8-REF, T3-REF, T4-REF, T5-REF and T6-REF in the standard 10-20 electrode system, with eyes closed, from 67 brain tumor patients and 42 normal subjects are obtained at a sampling rate of 256 Hz for about 15-20 min. The age is approximately uniformly distributed ranging from 14 to 53 years with a mean of 31.36 years and a median of 30 years for the brain tumor patients and ranging from 16 to 53 years with a mean of 33.34 years and a median of 34 years for the normal subjects. The majority of the brain tumor cases fall in the age group between 20 and 40 years.

2.2. Preprocessing

A 50 Hz FIR notch filter and an IIR EMG filter are used to eliminate the 50 Hz power line interference and the EMG artifacts, respectively. Certain artifacts such as eye movements, eye rolling and essential tremors are still present in the EEG records. Generally the removal of these artifacts by the traditional method of filtering is not possible since they fall in the useful bandwidth of cerebral EEG.

The method of Independent Component Analysis (ICA) is the best solution for this situation (Mammone *et al.*, 2012; Selvam *et al.*, 2011). A recently proposed ICA technique known as the Modified Wavelet ICA (MwICA) (Selvam *et al.*, 2011) was used to remove these artifacts.

From this set of “clean” EEG records, ten min ($10 \times 60 \times 256 = 1,53,600$ data points) of only the seven posterior channels, namely, Pz, P3, P4, O1, O2, T5 and T6, (‘REF’ has been dropped for convenience), of each EEG record are retained for the proposed analysis. All these 7-channel, 10 min EEG records are then bandpass-filtered to 1-40 Hz as this is the band of interest generally for most of the clinical applications (including this proposed analysis) (Picot *et al.*, 2009).

2.3. Basics of Bispectrum and Bicoherence

The concept of the bispectrum and its properties are extensively presented in the literature (Nikias and Raghuveer, 1987).

The bispectrum, $B^x(\omega_1, \omega_2)$ of a discrete, real, stationary (univariate) random process, $X(t) = \{x(t)\}$, $t = 0, \pm 1, \pm 2, \dots$ with zero mean is defined as the two-

dimensional Fourier transform of its third-order cumulant sequence i.e., $B^x(\omega_1, \omega_2) = F_2 \{ \gamma_3^x(\tau_1, \tau_2) \}$ where $F_2 \{ \cdot \}$ represents the 2-dimensional Fourier transform and $\gamma_3^x(\tau_1, \tau_2)$ is the third-order cumulant sequence. The third-order cumulant sequence is defined as $\gamma_3^x(\tau_1, \tau_2) = E \{ x(t)x(t+\tau_1)x(t+\tau_2) \}$ where $E \{ \cdot \}$ represents the statistical expectation. It can be shown that $B^x(\omega_1, \omega_2) = E \{ X(\omega_1)X(\omega_2)X^*(\omega_1, \omega_2) \}$ where $X(\omega)$ is the Fourier transform of $X(t)$ and $*$ represents the complex conjugate.

The properties that are worth mentioned here are: (i) the bispectrum is generally complex and has the magnitude and phase, (ii) the complex plane of the bispectrum has 12 symmetric regions, (iii) the bispectrum of a (linear) Gaussian process is everywhere zero, (iv) the bispectrum of a linear non-Gaussian process is everywhere constant and (v) the bispectrum of a nonlinear process is varying and peaking due to the phase couplings arising from the nonlinearity. The property (ii) implies that there is a "principal" region only in which the bispectrum needs to be computed. This principal region is the triangular region, $\omega_2 \geq 0, \omega_1$ and $\omega_1 + \omega_2 \leq \pi$.

A quadratically nonlinear system is the one with second-order nonlinearity e.g., a Volterra-Wiener system of order 2. The phenomenon, where the phases have the same relations as the frequencies, is called the quadratic phase coupling. Generally, in case of a frequency triplet in the output of a system, there are three possible cases: (i) when they are not harmonically related and are with random phases (i.e., no phase-coupling), the bispectrum is identically zero everywhere, (ii) when they are harmonically related with same phase relationship as frequency (i.e., phase-coupling), the bispectrum peaks due to the quadratic phase coupling and (iii) when they are harmonically related with random phases, a situation known as the frequency-coupling, the bispectrum may still peak under certain frequency and amplitude combinations. In all the three cases, the power spectrum invariably shows three peaks at these three frequencies. Due to this ambiguity, the estimated bispectral peak, especially from a finite data set, needs to be statistically tested for its significance in quantifying the presence of phase coupling. This ambiguity is overcome in the normalized bispectrum i.e., bicoherence (Elgar and Guza, 1988).

2.4. Estimation of Bispectrum and Bicoherence

The bispectral estimation methods are broadly classified into two: The conventional or non-parametric

methods and parametric methods. Two conventional or non-parametric methods, namely, the direct and indirect methods are generally found in the literatures. The parametric methods provide higher frequency resolution than the conventional methods and however suffer from the difficulty in determining the model order parameter. The conventional methods are popular for their ease of implementation. In this study, the direct class of the conventional methods is used. The direct method, which is based on the Fast Fourier Transform (FFT), is extensively described in (Kim and Powers, 1979; Nikias and Raghuveer, 1987). The bicoherence is the normalized version of the bispectrum (Kim and Powers, 1979). Given a discrete-time sequence, $x(n)$, $n = 0, 1, 2, \dots, N-1$, where N is the cardinality of the sequence, the sequence is segmented into K short sequences, $x_i(n) = x[n + iM - (i-1)(D-1)]$, $n = 0, 1, 2, \dots, M-1$ and $i = 0, 1, 2, \dots, K-1$ of M samples each with an overlap of D samples and each segment, $x_i(n)$ is centered as $x_i(n) - \bar{x}_i$ where \bar{x}_i is the arithmetic mean of $x_i(n)$. The bicoherence is then given by Equation 1 and 2:

$$b^x(k_1, k_2) = \frac{|B^x(k_1, k_2)|}{\left(\frac{1}{K} \sum_{i=1}^K |X_i(k_1)|^2 |X_i(k_2)|^2 \right)^{1/2} \left(\frac{1}{K} \sum_{i=1}^K |X_i(k_1 + k_2)|^2 \right)^{1/2}} \quad (1)$$

Where:

$$B^x(k_1, k_2) = \frac{1}{K} \sum_{i=0}^{K-1} \left(\frac{1}{M} \right) [X_i(k_1)X_i(k_2)X_i^*(k_1 + k_2)] \quad (2)$$

$$X_i(k) = \sum_{n=0}^{M-1} x_i(n) e^{-j \left[\frac{2\pi kn}{M} \right]} \quad (3)$$

The $X_i(k)$, $i = 0, 1, 2, \dots, K-1$ in Equation 3 are the discrete Fourier transforms of $x_i(n)$. Each segment, $x_i(n)$ can be windowed, if desired, in order to reduce the spectral leakage. The possible window functions include Hamming, Hann and Blackman. For a given N , the overlapping not only reduces the variance in the estimated bispectrum, as it increases the number of segments (i.e., K) for the ensemble averaging, but also provides a chance for a larger segment length (i.e., M) thereby improving the frequency resolution. A segment size of M samples gives a frequency resolution of $\Delta f = f_s/M$ where f_s is the sampling rate of the time series in

Hz and an ensemble average over K segments yields a variance proportional to N/K^2 . Hence, given a set of f_s and N the values of M and D must be properly selected so that the desired frequency resolution is achieved and at the same time the variance in the estimated bispectrum is as low as possible.

The estimation of the bispectrum requires the EEG signal to be stationary (Siu *et al.*, 2008). Since the EEG signals are non-stationary, they are often analyzed in segments in order to ensure the stationarity using a criterion of weak stationarity, which requires that the statistical parameters up to certain order remain (practically) constant over the entire period of the segment (Blanco *et al.*, 1995). The most popularly used weak stationarity is the second-order stationarity which requires the second-order statistics, mean and standard deviation (std), to remain constant at some prescribed significant level (e.g., 5%). The test for detecting the stationary segment length was carried out as follows. Each channel record was split into overlapping segments of length 1s (256 samples) called the bins. The overlap was 0.5s (128 samples). The means (bin means) and stds (bin stds) of each of these bins were computed. Then the means and stds of the bin means and bin stds were computed and the constancy of the bin means and bin stds was tested by checking whether their values lied within the 95% confidence interval of their means i.e., within $\text{mean} \pm 2\text{std}$ for various lengths of consecutive bins. The percentage of number of consecutive bins that met this criterion was calculated. All the consecutive bins of lengths, 2s, 3s and 4s remained (weakly) stationary. For the bin length of 5s and above, the number of consecutive bins that failed the stationarity test was above 6%. The bin length of 4s is chosen for this analysis as this ensures both the quasi-stationarity and the sufficiency of the data points to estimate the bispectrum (Miller *et al.*, 2004). Thus there are totally 10,050 (67 cases \times 10 min \times 60 sec / 4 sec) epochs of seven posterior channels of brain tumor EEG and 6,300 (42 cases \times 10 min \times 60 sec / 4 sec) epochs of seven posterior channels of normal EEG for the analysis.

From each 4-sec (1024 data points i.e., $N = 1024$) epoch of each posterior channel EEG record, the bicoherence is estimated with $K = 30$ overlapping segments of size $M = 256$ (1 sec) and overlap $D = 230$ (0.8984 sec i.e., 90% overlap). The largest (peak) Bicoherence (BIC) values in the [1-8] Hz (delta-theta)

band and the [8-13] Hz (alpha) band from each epoch of each posterior channel EEG record are then selected and averaged over the all the epochs (10,050 from brain tumor case and 6,300 from normal case) to obtain the final channelwise value for the proposed analysis.

2.5. Test for Gaussianity and Linearity

The Hinich test is the statistical hypothetical testing against the null hypothesis that the time series under consideration is a Gaussian process generated by a linear mechanism. The Hinich test is extensively formulated in (Hinich, 1982). In this test, two statistics, namely the S statistic and the R statistic derived from the routine bicoherence and the non-central Chi-square distribution, are used as the measure of deviation from the Gaussianity and the measure of degree of linearity, respectively. The test makes use of the facts that the bicoherence is zero for a Gaussian process and constant for a linear process. The hypothetical test for the linearity is considered only if the null hypothesis on the Gaussianity is rejected. The HOSA toolbox (freely available at <http://www.mathworks.com/matlabcentral/fileexchange/3013>) contains a MATLAB routine called `glstat` to calculate these statistics.

The S and R statistics required for the Hinich test are estimated using the bicoherence for each 4-sec epoch. It should be noted that though the Hinich test is performed on the epoch basis i.e., 10,050 epochs of brain tumor EEG and 6,300 epochs of normal EEG are analyzed individually, the results are used, on the average, to statistically determine the Gaussianity and linearity of the entire channel EEG record.

2.6. Comparison of Bispectra and Power Spectra

Nagata *et al.* (1985) used a parameter, called the power ratio index (PRI), to correlate the epicenter of the power dysfunction to the locality of the tumors in 15 patients with malignant brain tumors. He defined the PRI as the ratio of absolute power in the [1-8] Hz (delta-theta) band to that in the [8-30] Hz (alpha-beta) band. This parameter was also calculated from all the 4-second EEG epochs for the sake of comparison.

2.7. Statistical Analysis of BIC and PRI Values

The objective of the statistical analysis of the BIC and PRI values obtained is to assess the ability of the

proposed parameter in characterizing the presence of the brain tumor. Since the size of the data set is large enough (>30) in both the brain tumor and normal cases for the assumption of the normality of the estimated BIC and PRI values according to the central limit theorem (Montgomery and Runger, 2010), the two-tailed z-test is used to statistically test the channel-wise and overall mean BIC values in discriminating the brain tumor patient from the normal subject at 5% significance level and to estimate the 95% confidence intervals of these mean differences. The confidence interval is here rather than the p-values following the suggestions from (Gardner and Altman, 2000).

2.8. Implementation

The entire analysis was carried out using the MATLAB® (The MathWorks Inc., USA) software package on a personal computer.

3. RESULTS

The results of the test for determining the stationary segments using the second-order weak stationarity criterion are shown in **Fig. 1a and b** from two exemplary EEG records, one being the brain tumor and the other, normal. The dotted line shows the 95% confidence interval of the respective second-order statistic (mean or std) for two sets of 7 consecutive bins i.e., two 4-sec segments. The markers (asterisks) indicate the values of the respective statistic.

All the epochs of the seven posterior channel EEG records from both the brain tumor patients and normal subjects fail the Hinich test for the Gaussianity. Around 94% of the epochs fail the linearity test in the brain tumor case while around 81% of the epochs fail the linearity test in the normal case.

The contour plots of the bispectra of seven posterior-channel scalp EEG records from a normal subject and a brain tumor patient are shown respectively in **Fig. 2 and 3** where f_1 and f_2 are the bifrequency values. Strong self-couplings (Raghuveer, 1988) are seen in the [8-13] Hz (alpha) band in the normal case while strong self- (some phase-) couplings are seen in the [1-8] Hz (delta-theta) band in the brain tumor cases.

The channel-wise distributions of all the (largest) bicoherence (BIC) values (circles) and their mean values (triangles) in the [1-8] Hz (delta-theta) band and in the

[8-13] Hz (alpha) band across all the normal subjects and across all the brain tumor patients are shown respectively in **Fig. 4 and 5**. One of the two mean (largest) BIC values for the normal case falls in the (axial) middle of the [8-13] Hz (alpha) band and the other, in the (middle) middle of the [1-8] Hz (delta-theta) band whereas for the brain tumor case one falls in the left (axial) edge of the [8-13] Hz (alpha) band and the other, in the left (axial) middle of the [1-4] Hz (delta) band. In both cases, the [1-8] Hz (delta-theta) band and [8-13] Hz (alpha) band show some phase couplings and some self-couplings. On the average, for the normal case, a self-coupling is detected in the [8-13] Hz (alpha) band while a phase coupling is observed in the [1-8] Hz (delta-theta) band. In contrast, only self-couplings are seen in both the bands for the brain tumor case.

The 95% confidence intervals of the channel-wise mean (largest) BIC values for the normal and brain tumor cases in the [1-8] Hz (delta-theta) band and the [8-13] Hz (alpha) band are shown, respectively in **Fig. 6 and 7** where a value, 0.0 equals 0% coupling strength and 1.0 equals 100% coupling strength. The coupling strength in the [1-8] Hz (delta-theta band) varies from about 40% to 53% for the brain tumor patients while it is in the range around 23-42% for the normal subjects. Meanwhile, the coupling strength in the [8-13] Hz (alpha) band is very low ($<6.5\%$) for the brain tumor case and very high (about 27% to 52%) for the normal case, especially in and around the occipital region.

The statistics of the overall mean (largest) BIC values in the [1-8] Hz (delta-theta) band and the [8-13] Hz (alpha) band are shown in **Table 1**. The mean difference in the [8-13] Hz (alpha) band is statistically more significant ($p < 0.0001$) than that in the [1-8] Hz (delta-theta) band ($p < 0.001$). The 95% confidence interval on the mean difference of the largest BIC values in the [1-8] Hz (delta-theta) band is quite large, almost 11% more than the observed difference. However, this quantity in the [8-13] Hz (alpha) band is only approximately onequarter (more precisely 23%) of the observed difference.

Table 2 shows the statistical analysis of the Power Ratio Index (PRI) using the two-tailed z-test. The 95% confidence interval on the mean difference between the PRI for the brain tumor case and that for the normal case is statistically insignificant ($p < 0.05$).

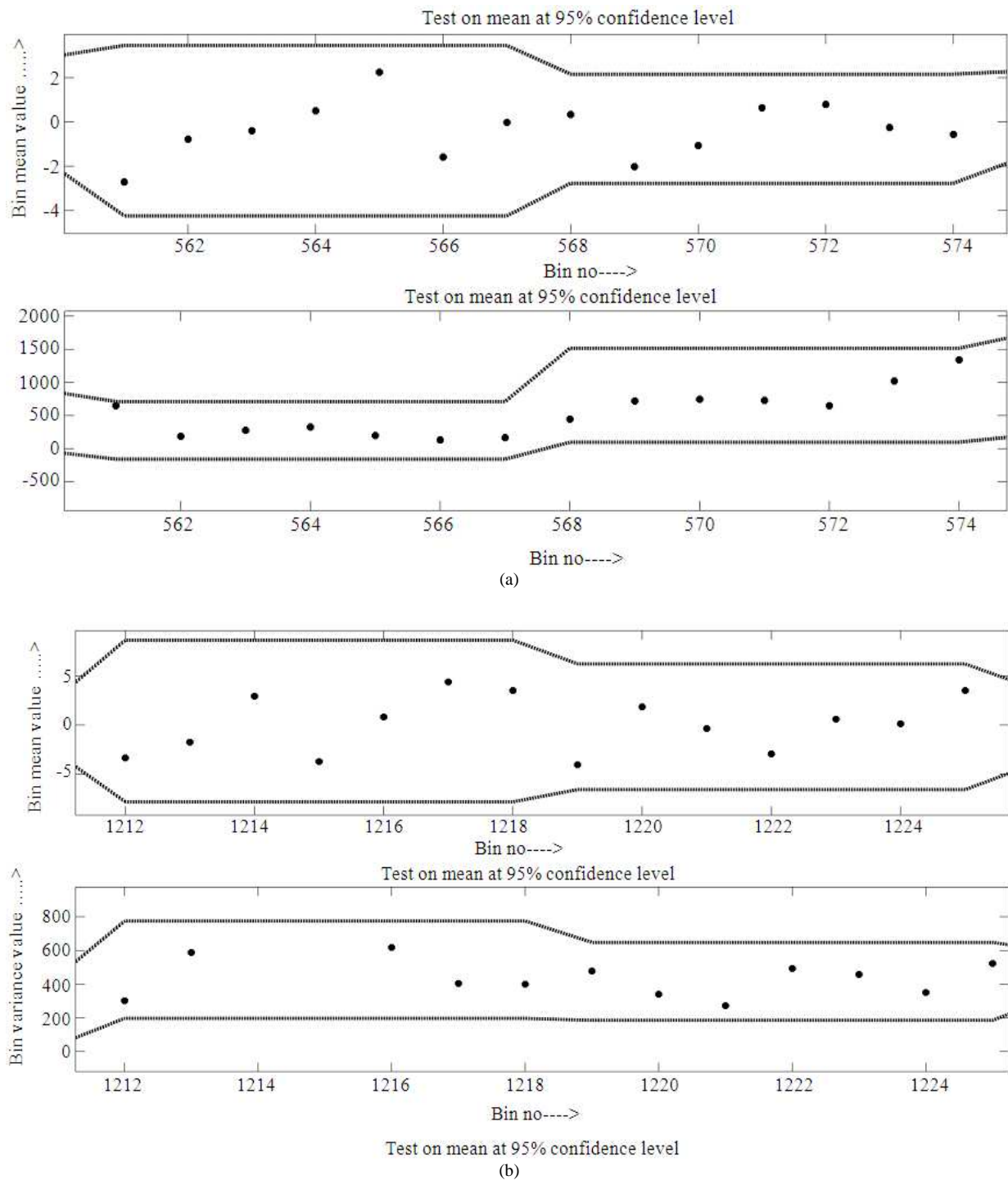


Fig. 1. Quasi Stationarity Test for Signal Segmentation of (a) an exemplary normal EEG record and (b) an exemplary brain tumor EEG record. The dotted line shows the 95% confidence interval (mean \pm 2std) of the second-order statistic (mean or std) for 7 consecutive bins i.e., a 4-sec segment and the markers (asterisks), the values of the statistic for the bins

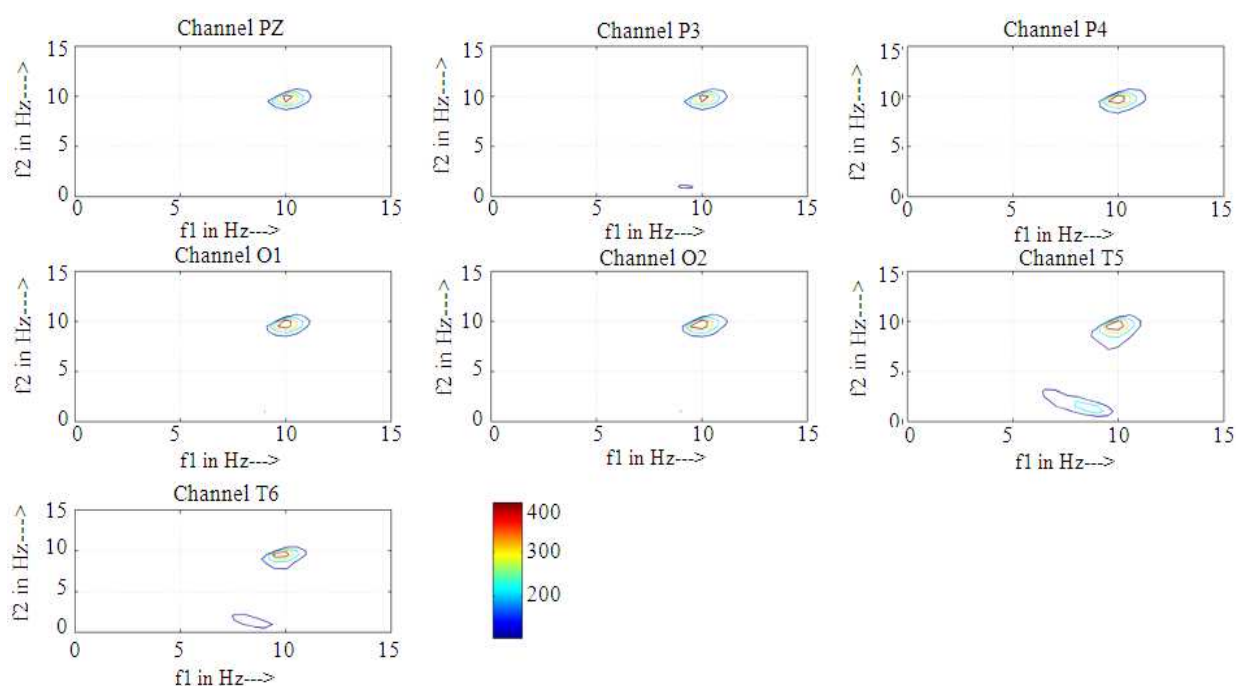


Fig. 2. The bispectra of seven posterior-channel EEG records of a normal subject

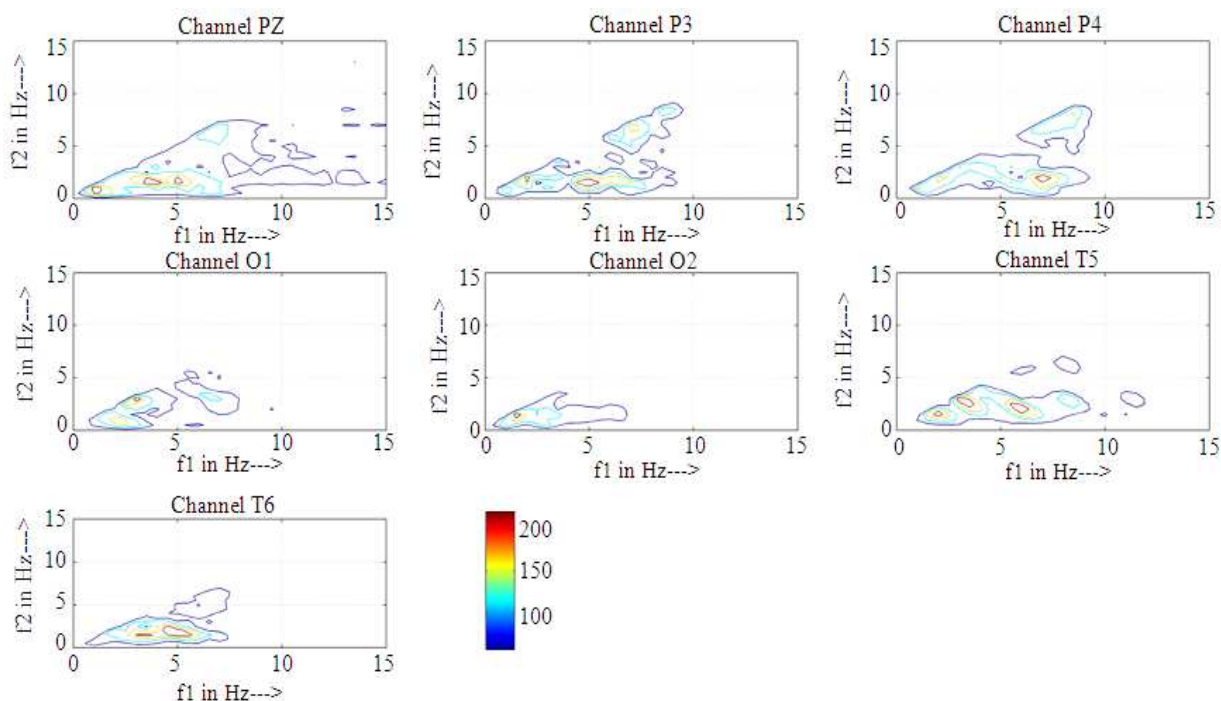


Fig. 3. The bispectra of seven posterior-channel EEG records of a brain tumor patient

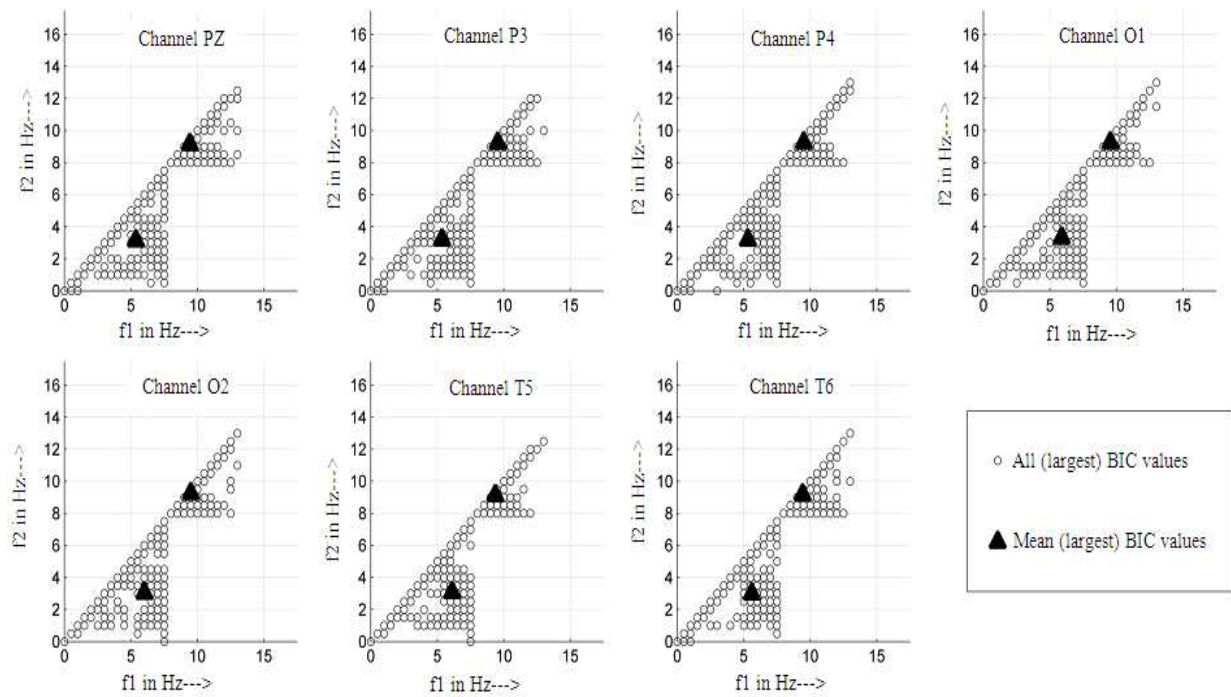


Fig. 4. The channel-wise distribution of the largest BIC values across all the normal cases

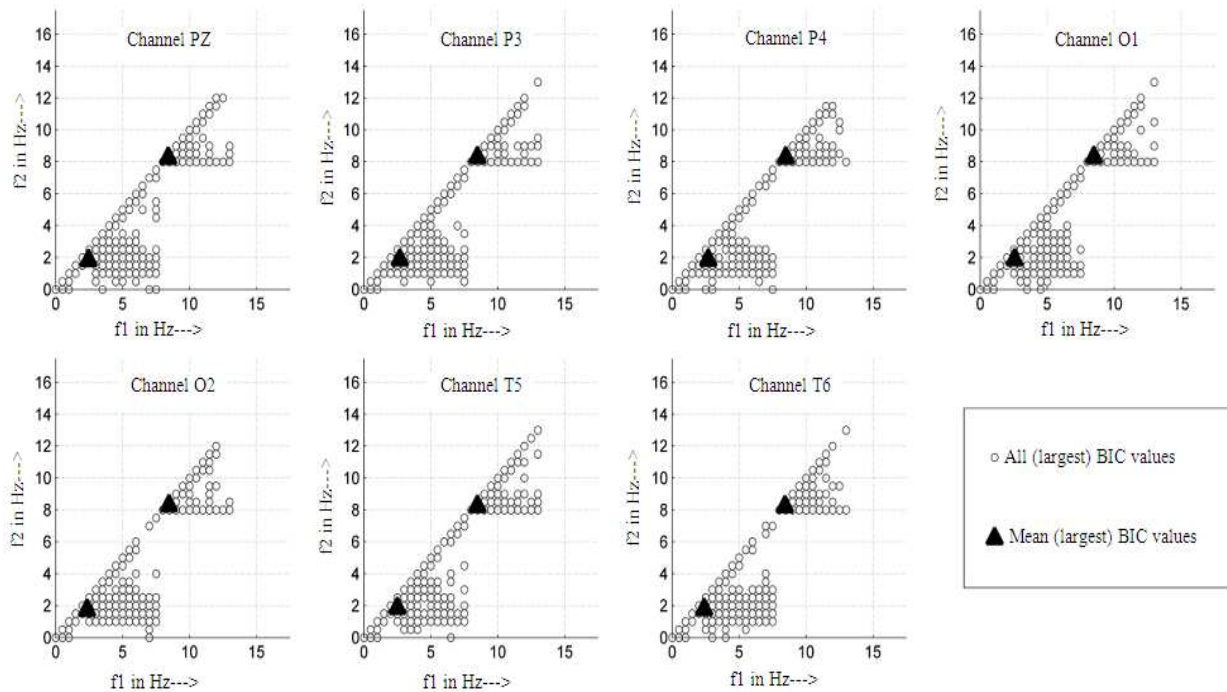


Fig. 5. The channel-wise distribution of the largest BIC values across all the brain tumor case

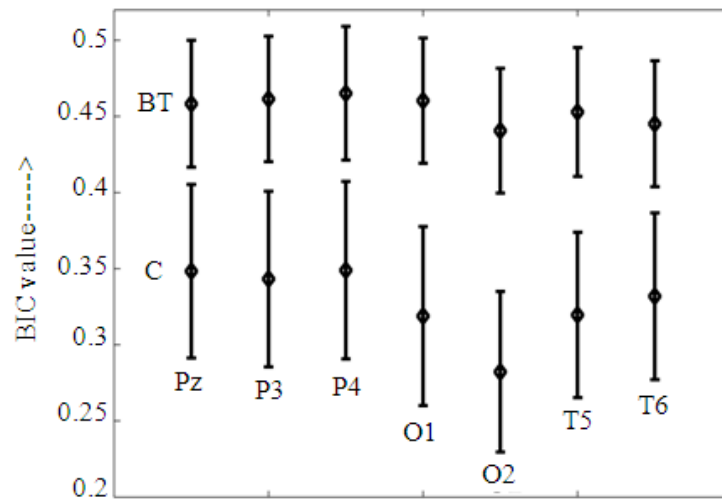


Fig. 6. The 95% confident intervals of the channel-wise mean (largest) BIC values from the normal (C) and brain tumor (BT) cases in the [1-8] Hz (delta-theta) band

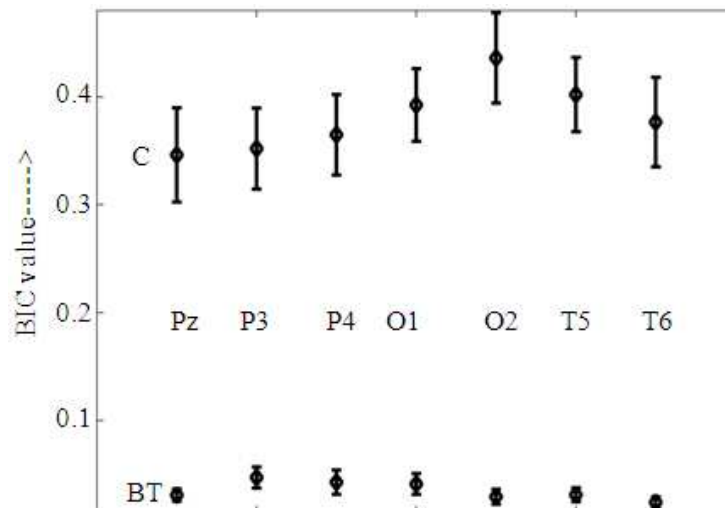


Fig. 7. The 95% confident intervals of the channel-wise mean (largest) BIC values from the normal (C) And Brain Tumor (BT) cases in the [8-13] Hz (alpha) band

Table 1. Statistics of overall mean (largest) BCI values

Table 1: Statistics of bifrequency mean (range), BIC values							
Band	Mean BIC value (std)		Mean bifrequency values (std)				95% confidence interval on mean Normal difference
	Normal	Brain tumor	Normal		Brain tumor		
			f ₁ (Hz)	f ₂ (Hz)	f ₁ (Hz)	f ₂ (Hz)	
[1-8] Hz (delta-theta) band	0.3276 (0.1873)	0.4549 (0.1749)	5.6342 (0.84)	3.4894 (0.82)	2.5288 (0.64)	1.8165 (0.3)	(0.0569, 0.1977)*
[8-13] Hz (alpha) band	0.3815 (0.1276)	0.0351 (0.0328)	9.6445 (0.31)	9.3432 (0.35)	8.4744 (0.15)	8.2586 (0.09)	(0.3070,0.3858)**

*p<0.001; **p<0.0001; std-standard deviation

Table 2. Statistics of overall mean PRI values

Mean BIC value (std)		95% confidence interval on mean difference
Normal	Brain Tumor	
0.7720 (0.4532)	3.1595 (10.0000)	(0.0000, 4.7859)*

*p>0.05; std-standard deviation

4. DISCUSSION

The normal cases are selected with an age distribution similar to the age distribution of the brain tumor cases since the EEG characteristics are generally dependent on the age (Nunez and Srinivasan, 2006). This avoids any ambiguity on the results obtained due to the age-dependency of the scalp EEG.

The results of the Hinich test for the Gaussianity and linearity suggest that an EEG record with eyes closed can most often be considered as a non-Gaussian process generated by a (possibly quadratically) nonlinear mechanism at a 95% confidence level (or 5% significance level). This complies with the result presented in (Pradhan *et al.*, 2012) and concludes that the attempt to quantify the phase couplings in these records is meaningful. The EEG time series remains non-Gaussian all the time and nonlinear most of the time. It may also be noted that the nonlinear structure in the EEG record is more dominant in the brain tumor case than in the normal case. Such dominance of the nonlinear character of the scalp EEG has been already observed under certain other pathological conditions as well (Diambra *et al.*, 2001). This might be due to the fact that the presence of a tumor adds inhomogeneity (spatial non-uniformity) to the locality of the tumor (O'Connor and Robinson, 2005).

The reason for the selection of the posterior channels is the fact that the alpha activity is dominant in the frequency range 8-13 Hz during wakefulness in the posterior region of the head, particularly over the occipital region with eyes closed (Nunez and Srinivasan, 2006). The strong self-couplings in the [8-13] Hz band for a normal case indicate the presence of strong alpha rhythms. The presence of strong alpha rhythms in the frequency range 8-13 Hz is generally an indication of the healthy brains in 95% of the cases. Such strong self-couplings in the alpha rhythms for normal subjects with eyes closed have already observed by Barnett *et al.* (1971). These self-couplings are almost lost in the brain tumor cases. The noteworthy association of the alpha rhythms

with a variety of physiological, cognitive and behavioral states of the brain is once again evident from these results.

Since the phase coupling is considered as a sort of (nonlinear) phase synchrony (Kim and Powers, 1979; Siu *et al.*, 2008), the strong phase (self-) coupling in the alpha band suggests a strong phase synchrony in the alpha band for the normal case. Such a strong phase synchrony is seen in the delta-theta band for both the normal and brain tumor cases.

For the brain tumor patients the [1-4] Hz (delta) band exhibits slightly stronger phase synchrony while the phase synchrony in the [8-13] Hz (alpha) band seems to be significantly lost. Decreased coupling strengths in the high-frequency band and increased ones in the low-frequency band are often indicative of low brain activity (e.g., levels of sedation, depth of sleep) (Barnett *et al.*, 1971; Roustan *et al.*, 2005; Venkatakrishnan *et al.*, 2011; Yufune *et al.*, 2011). Thus the presence of tumor reduces the brain activity.

The absence of phase coupling, which is found in the [1-8] Hz (delta-theta) band for the normal case, under the presence of a brain tumor is an indication of loss of quadratically nonlinear interaction in this band. These findings support the nonlinear interdependency between the brain regions and its contribution to the alpha rhythms (Breakspear and Terry, 2002).

The physiological model proposed for the corticothalamic dynamics in assessing the characteristics of the EEG associated with the thalamic tumors in (O'Connor and Robinson, 2005) and subsequent discussions therein are the evidences for the effect of the supratentorial tumors on the corticothalamic inputs and corticocortical networks which are responsible for the generation of the alpha rhythms (Schaul, 1998). The finding presented in this study not only supports this model but might also be helpful in extending this model to other supratentorial tumors on the assumption of non-Gaussianity for the model response, the EEG.

The final note is that the use of the bispectral index in the assessment of the level of consciousness during the administration of the anesthetic agent for the brain tumor patients has become questionable and is under serious study (e.g., <http://clinicaltrials.gov/show/NCT01060631>). The anticipated correlation between the quadratic phase coupling phenomenon and the presence of the brain tumor is possibly the outcome of the presented analysis.

The Power Ratio Index (PRI) is not as significant in discriminating the brain tumor EEG from the normal one as in assessing the tumor locality (Nagata *et al.*, 1985). The statistical results on the Power Ratio Index (PRI) based on the power spectrum once again prove that the bispectrum is far more informative than the power spectrum.

The strengths of the proposed analysis are the robustness of the bispectrum against the observational noises such as (white or colored) Gaussian noises (Nikias and Raghuveer, 1987) and the fact that the discriminating factor is the coupling strength in the alpha band which is less or not at all corrupted by the noises whereas the delta-theta and beta-gamma bands are always sensitive, respectively to low-frequency artifacts such as movements, EKG and high-frequency artifacts such as EMG (Nunez and Srinivasan, 2006). The major setback of the method is its computational expenses. However, with latest advancements in the processor technology, the computational speed is not an issue.

5. CONCLUSION

From the Results and Discussion, it can easily be concluded that the bispectral analysis of the posterior channel EEG records can be useful for the diagnosis of (supratentorial) brain tumors. The absence of phase couplings in the [1-8] Hz (delta-theta) band and subsequently self-couplings in the 8-13 Hz (alpha) band at the posterior, especially occipital region of the head is indicative of a structural lesion in the supratentorial region of the brain.

6. ACKNOWLEDGEMENT

The researchers thank Dr. J. Mohanasundaram MD, Dean, Madras Medical College (MMC), Chennai and Dr. A. Sundaram MD, Vice-Principal, MMC, Chennai for having approved the collection of the data required for this research study from MMC and Dr. V. Sundar, Prof., Dept. of Neurosurgery, MMC, Chennai for having rendered his help in getting the approval from the Ethical Committee of MMC. The authors also thank Dr. S. Arunan, Asst. Prof., Dept. of Neurology and Dr. K. Thirumaran, Asst. Prof., Dept. of Neurosurgery for having rendered their help in selecting the cases (brain tumor patients and normal subjects) and recording their scalp EEGs. The first author is indebted to Dr. N. Kumaravel, Prof and Head, Dept. of ECE, Anna University, Chennai who has been a consistent encouragement to all his research activities.

7. REFERENCES

- Accolla, E.A., P.W. Kaplan, M. Maeder-Ingvar, S. Jukopila and A.O. Rossetti, 2011. Clinical correlates of Frontal Intermittent Rhythmic Delta Activity (FIRDA). *Clin. Neurophysiol.*, 122: 27-31. PMID: 20673647
- Bagchi, B.K., K.A. Kooh, B.T. Selving and H.D. Calhoun, 1961. Subtentorial tumors and other lesions: An electroencephalographic study of 121 cases. *Electr. Clin. Neurophysiol.*, 13: 180-192. DOI: 10.1016/0013-4694(61)90134-1
- Barnett, T.P., L.C. Johnson, P. Naitoh, N. Hicks and C. Nute, 1971. Bispectrum analysis of electroencephalogram signals during waking and sleeping. *Science*, 172: 401-402. DOI: 10.1126/science.172.3981.401
- Black, P.M., 2010. Symptoms and diagnosis. National Brain Tumor Society.
- Blanco, S., H. Garcia, R.Q. Quiroga, L. Romanelli and O.A. Rosso, 1995. Stationarity of the EEG series. *IEEE Engg. Med. Biol.*, 14: 395-399. DOI: 10.1109/51.395321
- Breakspear, M. and J.R. Terry, 2002. Detection and description of non-linear interdependence in normal multichannel human EEG data. *Clin. Neurophysiol.*, 113: 735-753. DOI: 10.1016/S1388-2457(02)00051-2
- Buzsaki, G. and A. Draguhn, 2004. Neuronal oscillations in cortical networks. *Science*, 304: 1926-1929. DOI: 10.1126/science.1099745
- Chetty, S. and G.K. Venayagamoorthy, 2002. A neural network based detection of brain tumours using electroencephalography. *Proceedings of the IASTED International Conference Artificial Intelligence and Soft Computing, (ICAISC' 02)*, Jul. 17-19, Banff, Canada, pp: 391-396.
- Diambra, L., C.P. Malta, A. Capurro and J. Fernandez, 2001. Nonlinear structures in electroencephalogram signals. *Physica A: Stat. Mech. Applic.*, 300: 505-520. DOI: 10.1016/S0378-4371(01)00352-1
- Elgar, S. and R.T. Guza, 1988. Statistics of bicoherence. *IEEE Trans. Acoust. Speech Signal Process*, 36: 1667-1668. DOI: 10.1109/29.7555
- Fattal-Valevski, A., N. Nissan, U. Kramer and S. Constantini, 2012. Seizures as the clinical presenting symptom in children with brain tumors. *J. Child. Neurol.*, 28: 292-296. DOI: 10.1177/0883073812445786

- Gardner, M.J. and D.G. Altman, 2000. Confidence Intervals Rather than P Values. In: Statistics with confidence, Altman, D., D. Machin, T. Bryant and S. Gardner (Eds.), Wiley-Blackwell, ISBN-10: 0727913751, pp: 15-27.
- Habl, M., C. Bauer, C. Ziegeus, E.W. Lang and F. Schulmeyer, 2000. Can ICA help identify brain tumor related EEG signals? University of Regenburh.
- Hartman, A.L. and R.P. Lesser, 2012. Brain Tumors and Other Space-Occupying Lesions. In: Electroencephalography: Basic Principles, Clinical Applications and Related Fields, Niedermeyer, E. and F.L.D. Silva (Eds.), Lippincott Williams and Wilkins, Philadelphia, ISBN-10: 1469801752, pp: 321-327.
- Hinich, M.J., 1982. Testing for Gaussianity and linearity of a stationary time series. J. Time Ser. Anal., 3: 169-176. DOI: 10.1111/j.1467-9892.1982.tb00339.x
- Hosseini, P.T., R. Shalbaf, A.M. Nasrabadi, 2010. Extracting a seizure intensity index from one-channel EEG Signal using bispectral and detrended fluctuation analysis. J. Biomed. Sci. Eng., 3: 253-261. DOI: 10.4236/jbise.2010.33034
- Karamah, F.N. and M.A. Dahleh, 2000. Automated classification of EEG signals in brain tumor diagnostics. Proceedings of the American Control Conference, Jun. 28-30, IEEE Xplore Press, Chicago, IL., pp: 4169-4173. DOI: 10.1109/ACC.2000.877006
- Kim, Y.C. and E.J. Powers, 1979. Digital bispectral analysis and its applications to nonlinear wave interactions. IEEE Trans. Plasma Sci., 7: 120-131. DOI: 10.1109/TPS.1979.4317207
- Mammone, N., F.L. Foresta and F.C. Morabito, 2012. Automatic artifact rejection from multichannel Scalp EEG by wavelet ICA. IEEE Sens. J., 12: 533-542. DOI: 10.1109/JSEN.2011.2115236
- MFBTRI, 2013. Brain tumor symptoms. The Musella Foundation for Brain Tumor Research and Information.
- Miller, A., J.W. Sleigh, J. Barnard and D.A. Steyn-Ross, 2004. Does bispectral analysis of the electroencephalogram add anything but complexity? Br. J. Anaesth., 92: 8-13. DOI: 10.1093/bja/ae003
- Montgomery, D.C. and G.C. Runger, 2010. Applied Statistics and Probability for Engineers. 5th Edn., John Wiley and Sons, ISBN-10: 0470505788, pp: 776.
- Murugesan, M. and R. Sukanesh, 2009. Automated detection of brain tumor in EEG signals using artificial neural networks. Proceedings of the International Conference on Advances in Computing, Control and Telecommunication Technologies, Dec. 28-29, IEEE Xplore Press, Trivandrum, Kerala, pp: 284-288. DOI: 10.1109/ACT.2009.77
- Nagata, K., C.E. Gross, G.W. Kindt, M.J. Geier and G.R. Adey, 1985. Topographic electroencephalographic study with power ratio index mapping in patients with malignant brain tumors. Neurosurgery, 17: 613-619. PMID: 4058699
- Nikias, C.L. and M.R. Raghuveer, 1987. Bispectrum estimation: A digital signal processing framework. IEEE Proc., 75: 869-891. DOI: 10.1109/PROC.1987.13824
- Nunez, P.L. and R. Srinivasan, 2006. Electric Fields of the Brain: The Neurophysics of EEG. 2nd Edn., Oxford University Press, New York, ISBN-10: 019505038X, pp: 611.
- O'Connor, S.C. and P.A. Robinson, 2005. Analysis of the electroencephalographic activity associated with thalamic tumors. J. Theor. Biol., 233: 271-286. DOI: 10.1016/j.jtbi.2004.10.009
- Picot, A., S. Charbonnier and A. Caplier, 2009. Monitoring Drowsiness On-Line using a Single Encephalographic Channel. In: Biomedical Engineering, De Mello, B.C.A. (Ed.), In-Tech, Croatia, pp: 145-164.
- Pradhan, C., S.K. Jena, S.R. Nadar and N. Pradhan, 2012. Higher-order spectrum in understanding nonlinearity in EEG rhythms. Comput. Math. Meth. Med., 2012: 1-8. DOI: 10.1155/2012/206857
- Raghuveer, M.R., 1988. High resolution estimation of quadratic phase coupling in nonlinear systems. Proceedings of American Control Conference, Jun. 15-17, IEEE Xplore Press, Atlanta, Ga, USA., pp: 2124-2128.
- Roustan, J.P., S. Valette, P. Aubas, G. Rondouin and X. Capdevila, 2005. Can electroencephalographic analysis be used to determine sedation levels in critically ill patients? Anesth Analg, 101: 1141-51. DOI: 10.1213/01.ane.0000167782.47957.e1
- Saikia, A. and S.M. Hazarika, 2011. Bispectrum analysis of EEG during observation and imagination of hand movement. Proceedings of IEEE Students' Technical Symposium, Jan. 14-16, IEEE Xplore Press, Kharagpur, pp: 128-133. DOI: 10.1109/TECHSYM.2011.5783840

- Schaul, N., 1998. The fundamental neural mechanisms of electroencephalography. *Elect. Clin. Neuro.*, 106: 101-107. DOI: 10.1016/S0013-4694(97)00111-9
- Selvam, V.S. and S. Shenbagadevi, 2011. Brain tumor detection using scalp EEG with modified wavelet-ICA and multi layer feed forward neural network. *Proceedings of the Annual International Conference of the IEEE Engineering in Medicine and Biology Society*, Aug. 30-Sept. 3, Boston, MA., pp: 6104-6109. DOI: 10.1109/IEMBS.2011.6091508
- Selvam, V.S., S. Shenbagadevi, V. Padhma, D. Sujatha and R. Sharmila, 2011. Two new approaches of independent component analysis. *Proceedings of the 4th International Conference, ObCom*, Dec. 9-11, CCIS, Springer, pp: 503-513. DOI: 10.1007/978-3-642-29216-3_55
- Silipo, R., G. Deco and H. Bartsch, 1999. Brain tumor classification based on EEG hidden dynamics. *Intell. Data Anal.*, 3: 287-306. DOI: 10.1.1.37.5882
- Siu, K.L., J.M. Ahn, K. Ju, J.H. Lee and K. Shin *et al.*, 2008. Statistical approach to quantify the presence of phase coupling using the bispectrum. *IEEE Trans. Biomed. Eng.*, 55: 1512-1520. DOI: 10.1109/TBME.2007.913418
- Swarnkar, V., U. Abeyratne and C. Hukins, 2010. Objective measure of sleepiness and sleep latency via bispectrum analysis of EEG. *Med. Biol. Eng. Comput.*, 48: 1203-1213. DOI: 10.1007/s11517-010-0715-x
- Venkatakrisnan, P., R. Suganesh and S. Sangeetha, 2011. Detection of quadratic phase coupling from human EEG signals using higher order statistics and spectra. *Signal Image Video Process.*, 5: 217-229. DOI: 10.1007/s11760-010-0156-x
- Walter, G., 1936. The location of cerebral tumors by electroencephalography. *Lancet*, 228: 305-308. DOI: 10.1016/S0140-6736(01)05173-x
- Waternberg, N., F. Alehan, R. Dabby, T. Lerman-Sagie and P. Pavot *et al.*, 2002. Clinical and radiologic correlates of frontal intermittent rhythmic delta activity. *J. Clin. Neurophysiol.*, 19: 535-539. PMID: 12488784
- Yufune, S., I. Takamatsu, K. Masui and T. Kazama, 2011. Effect of remifentanyl on plasma propofol concentration and bispectral index during propofol anaesthesia. *Br. J. Anaesth.*, 106: 208-214. DOI: 10.1093/bja/aeq334