

# Stein's method for nonconventional sums

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## Abstract

We obtain almost optimal convergence rate in the central limit theorem for (appropriately normalized) "nonconventional" sums of the form  $S_N = \sum_{n=1}^N (F(\xi_n, \xi_{2n}, \dots, \xi_{\ell n}) - \bar{F})$ . Here  $\{\xi_n : n \geq 0\}$  is a sufficiently fast mixing vector process with some stationarity conditions,  $F$  is bounded Hölder continuous function and  $\bar{F}$  is a certain centralizing constant. Extensions to more general functions  $F$  will be discusses, as well. Our approach here is based on the so called Stein's method, and the rates obtained in this paper significantly improve the rates in [7]. Our results hold true, for instance, when  $\xi_n = (T^n f_i)_{i=1}^p$  where  $T$  is a topologically mixing subshift of finite type, a hyperbolic diffeomorphism or an expanding transformation taken with a Gibbs invariant measure, as well as in the case when  $\{\xi_n : n \geq 0\}$  forms a stationary and exponentially fast  $\phi$ -mixing sequence, which, for instance, holds true when  $\xi_n = (f_i(\Upsilon_n))_{i=1}^p$  where  $\Upsilon_n$  is a Markov chain satisfying the Doeblin condition considered as a stationary process with respect to its invariant measure.

**Keywords:** central limit theorem; Berry-Esseen theorem; mixing; nonconventional setup; Stein's method.

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## 1 Introduction

Let  $\Phi$  be the standard normal distribution function and let  $X_1, X_2, X_3 \dots$  be a sequence of independent and identically distributed random variables such that  $\mathbb{E}X_1 = 0$  and  $0 < \mathbb{E}X_1^2 = \sigma^2 < \infty$ . The classical Berry-Esseen theorem provides a uniform approximation of the error term in the central limit theorem (CLT) for the sums  $\hat{S}_n = \frac{1}{\sqrt{n}\sigma} \sum_{k=1}^n X_k$ , stating that for any  $n \in \mathbb{N}$ ,

$$\sup_{x \in \mathbb{R}} |F_n(x) - \Phi(x)| \leq \frac{C\mathbb{E}|X_1|^3}{\sqrt{n}} \quad (1.1)$$

where  $F_n$  is the distribution function of  $\hat{S}_n$  (see Section 6 of Ch. III in [17]) and  $C > 0$  is an absolute constant which by efforts of many researchers was optimized by now to a number a bit less than 1/2.

During the last 50 years there were several extensions of the CLT for sums of weakly dependent random variables and for martingales, including many estimates of error terms. Among the most used methods in the case of weak dependence are Gordin's

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method for martingale approximation (see [5], [12] and [6]) and Stein's method (see [15]). While Stein's method can yield close to optimal convergence rate (see [15] and [13]), the martingale approximation method can not, since Berry-Esseen type estimates for martingales do not yield (in general) optimal convergence rates even for sums of independent random variables (see, for instance [6] and [1]).

Partially motivated by the research on nonconventional ergodic theorems (the term "nonconventional" comes from [4]), probabilistic limit theorems for sums of the form  $S_N = \sum_{n=1}^N F(\xi_{q_1(n)}, \xi_{q_2(n)}, \dots, \xi_{q_\ell(n)})$  have become a well studied topic. Here  $\{\xi_n, n \geq 0\}$  is a sufficiently fast mixing processes with some stationarity properties and  $F$  is a function satisfying some regularity conditions. The summands here are nonstationary and long range dependent which makes it difficult to apply standard methods. This line of research started in [9], where a functional CLT was proved for the normalized sums  $N^{-\frac{1}{2}} S_{[Nt]}$  using characteristic function estimates. In [11] a functional CLT was established for more general  $q_i$ 's than in [9], where one of the main parts of the proof was showing that the martingale approximation approach is applicable. These results included the case when  $q_i(n) = in$ , which was the original motivation for the study of nonconventional averages (see [4]). In [7] the authors estimated the convergence rate of  $\mathcal{Z}_N = N^{-\frac{1}{2}} S_N$  in the Kolmogorov (uniform) metric towards its weak limit under the assumptions of [11]. The proof relied on Berry-Esseen type results for martingales, which led to estimates of order  $N^{-\frac{1}{10}} \ln(N+1)$ , which is far from optimal. In the special case when  $\xi_n$ 's are independent the authors provided optimal rate of order  $N^{-\frac{1}{2}}$  relying on Stein's method for sums of locally dependent random variables (see [3]).

The goal of this paper is to show that Stein's method is applicable for nonconventional sums when  $\xi_n$ 's are weakly dependent, and to significantly improve the rates obtained in [7]. We first consider the case when  $F$  is a bounded Hölder continuous function and  $q_i(n) = in$  for any  $1 \leq i \leq \ell$  and  $n \in \mathbb{N}$ , and (in the self normalized case) provide almost optimal upper bound of the form

$$\sup_{x \in \mathbb{R}} |P(S_N \leq x \sqrt{\mathbb{E} S_N^2}) - \Phi(x)| \leq C N^{-\frac{1}{2}} \ln^2(N+1) \quad (1.2)$$

assuming that  $D^2 > 0$ , where  $D^2 = \lim_{N \rightarrow \infty} N^{-1} \mathbb{E} S_N^2$ . We also obtain rates of the form

$$\sup_{x \in \mathbb{R}} |P(N^{-\frac{1}{2}} S_N \leq x) - \Phi(x D^{-1})| \leq C_\epsilon N^{-\frac{1}{2} + \epsilon} \quad (1.3)$$

where  $\epsilon > 0$  is an arbitrary positive constant and  $C_\epsilon$  is a constant which in general depends on  $\epsilon$ . When  $\{\xi_n : n \geq 0\}$  forms a stationary and exponentially fast  $\phi$ -mixing sequence then, in fact, we show that (1.2) and (1.3) hold true for any bounded function  $F$  which is not necessarily continuous. Convergence rates for more general functions and more general indexes  $q_i(n)$ 's will be discussed, as well.

As in [7], our results hold true when, for instance,  $\xi_n = T^n f$  where  $f = (f_1, \dots, f_d)$ ,  $T$  is a topologically mixing subshift of finite type, a hyperbolic diffeomorphism or an expanding transformation taken with a Gibbs invariant measure, as well as in the case when  $\xi_n = f(\Upsilon_n)$ ,  $f = (f_1, \dots, f_d)$  where  $\Upsilon_n$  is a Markov chain satisfying the Doeblin condition considered as a stationary process with respect to its invariant measure. In fact, any stationary and exponentially fast  $\phi$ -mixing sequence  $\{\xi_n, n \geq 0\}$  can be considered. In the dynamical systems case each  $f_i$  should be either Hölder continuous or piecewise constant on elements of Markov partitions. As an application we can consider  $\xi_n = ((\xi_n)_1, \dots, (\xi_n)_\ell)$ ,  $(\xi_n)_j = \mathbb{I}_{A_j}(T^n x)$  in the dynamical systems case and  $(\xi_n)_j = \mathbb{I}_{A_j}(\Upsilon_n)$  in the Markov chain case, where  $\mathbb{I}_A$  is the indicator of an appropriate set  $A$ . Let  $F = F(x_1, \dots, x_\ell)$ ,  $x_j = (x_j^{(1)}, \dots, x_j^{(\ell)})$  be a bounded Hölder continuous function which identifies with the function  $G(x_1, \dots, x_\ell) = x_1^{(1)} \cdot x_2^{(2)} \cdots x_\ell^{(\ell)}$  on the cube  $([0, 1]^\varphi)^\ell$ .

Let  $N(n)$  be the number of  $l$ 's between 0 and  $n$  for which  $T^{q_j(l)}x \in A_j$  for  $j = 0, 1, \dots, \ell$  (or  $\Upsilon_{q_j(l)} \in A_j$  in the Markov chains case), where we set  $q_0 = 0$ , namely the number of  $\ell$ -tuples of return times to  $A_j$ 's (either by  $T^{q_j(l)}$  or by  $\Upsilon_{q_j(l)}$ ). Then our results yield a central limit theorem with almost optimal convergence rate for the numbers  $N(n)$ .

## 2 Preliminaries and main results

Our setup consists of a  $\wp$ -dimensional stochastic process  $\{\xi_n, n \geq 0\}$  on a probability space  $(\Omega, \mathcal{F}, P)$  and a family of sub- $\sigma$ -algebras  $\mathcal{F}_{k,l}$ ,  $-\infty \leq k \leq l \leq \infty$  such that  $\mathcal{F}_{k,l} \subset \mathcal{F}_{k',l'} \subset \mathcal{F}$  if  $k' \leq k$  and  $l' \geq l$ . We will impose restrictions of the mixing coefficients

$$\phi(n) = \sup\{\phi(\mathcal{F}_{-\infty,k}, \mathcal{F}_{k+n,\infty}) : k \in \mathbb{Z}\} \quad (2.1)$$

where we recall that for any two sub- $\sigma$ -algebras  $\mathcal{G}, \mathcal{H} \subset \mathcal{F}$ ,

$$\phi(\mathcal{G}, \mathcal{H}) = \sup\left\{\left|\frac{P(A \cap B)}{P(A)} - P(B)\right| : A \in \mathcal{G}, B \in \mathcal{H}, P(A) > 0\right\}.$$

In order to ensure some applications, in particular, to dynamical systems we will not assume that  $\xi_n$  is measurable with respect to  $\mathcal{F}_{n,n}$  but instead impose conditions on the approximation rates

$$\beta_\infty(r) = \sup_{k \geq 0} \|\xi_k - \mathbb{E}[\xi_k | \mathcal{F}_{k-r, k+r}]\|_{L^\infty} \quad (2.2)$$

where  $\|X\|_{L^\infty}$  denotes the essential supremum of the absolute value of a random variable  $X$ .

We do not require stationarity of the process  $\{\xi_n, n \geq 0\}$ , assuming only that the distribution of  $\xi_n$  does not depend on  $n$  and that the joint distribution of  $(\xi_n, \xi_m)$  depends only on  $n - m$ , which we write for further reference by

$$\xi_n \sim \mu \quad \text{and} \quad (\xi_n, \xi_m) \sim \mu_{m-n} \quad (2.3)$$

where  $Y \sim \mu$  means that  $Y$  has  $\mu$  for its distribution.

Let  $F = F(x_1, \dots, x_\ell) : (\mathbb{R}^\wp)^\ell \rightarrow \mathbb{R}$ ,  $\ell \geq 1$  be a bounded Hölder function and let  $M > 0$  and  $\kappa \in (0, 1]$  be such that

$$|F(x)| \leq M \quad \text{and} \quad (2.4)$$

$$|F(x) - F(y)| \leq M \sum_{i=1}^{\ell} |x_i - y_i|^\kappa \quad (2.5)$$

for any  $x = (x_1, \dots, x_\ell)$  and  $y = (y_1, \dots, y_\ell)$  in  $(\mathbb{R}^\wp)^\ell$ . To simplify formulas we assume the centering condition

$$\bar{F} = \int F(x_1, x_2, \dots, x_\ell) d\mu(x_1) d\mu(x_2) \dots d\mu(x_\ell) = 0 \quad (2.6)$$

which is not really a restriction since we can always replace  $F$  by  $F - \bar{F}$ . The main goal of this paper is to prove a central limit theorem with convergence rate for the normalized sums  $(c_N)^{-1} S_N$ , where

$$S_N = \sum_{n=1}^N F(\xi_n, \xi_{2n}, \dots, \xi_{\ell n})$$

and either  $c_N = N^{-\frac{1}{2}}$  or  $c_N = \sqrt{\mathbb{E} S_N^2}$ . Our results will rely on

**Assumption 2.1.** *There exist  $d > 0$  and  $c \in (0, 1)$  such that*

$$\phi(n) + (\beta_\infty(n))^\kappa \leq dc^n \quad (2.7)$$

for any  $n \in \mathbb{N}$ .

Before formulating our main results we will first state the following

**Theorem 2.2.** *Suppose that Assumption (2.1) is satisfied. Then the limit  $D^2 = \lim_{N \rightarrow \infty} N^{-1} \mathbb{E} S_N^2$  exists and there exists  $C_1 > 0$  which depends only on  $\ell, c$  and  $d$  such that*

$$|\mathbb{E} S_N^2 - D^2 N| \leq C_1 M^2 N^{\frac{1}{2}} \quad (2.8)$$

for any  $N \in \mathbb{N}$ . Moreover,  $D^2 > 0$  if and only if there exists no stationary in the wide sense process  $\{V_n : n \geq 1\}$  such that

$$F(\xi_n^{(1)}, \xi_{2n}^{(n)}, \dots, \xi_{\ell n}^{(\ell)}) = V_{n+1} - V_n, \quad P\text{-almost surely}$$

for any  $n \in \mathbb{N}$ , where  $\xi^{(i)}, i = 1, \dots, \ell$  are independent copies of  $\xi = \{\xi_n : n \geq 1\}$ .

This theorem is a consequence of the arguments in [11], [10] and [7] and is formulated here for readers' convenience.

Next, recall that the Kolmogorov (uniform) metric is defined for each pairs of distributions  $\mathcal{L}_1$  and  $\mathcal{L}_2$  on  $\mathbb{R}$  with distribution functions  $G_1$  and  $G_2$  by

$$d_K(\mathcal{L}_1, \mathcal{L}_2) = \sup_{x \in \mathbb{R}} |G_1(x) - G_2(x)|.$$

For any random variable  $X$  we denote its law by  $\mathcal{L}(X)$ . Our main result is the following

**Theorem 2.3.** *Suppose that Assumption (2.1) holds true and that  $D^2 > 0$ . Set  $s_N = \sqrt{\mathbb{E} S_N^2}$  and  $Z_N = (s_N)^{-1} S_N$  when  $s_N > 0$ , while when  $s_N = 0$  we set  $Z_N = N^{-\frac{1}{2}} S_N$ . Let  $\mathcal{N}(0, 1)$  be the zero mean normal distribution with variance 1. Then there exists a constant  $C > 0$  which depends only on  $\ell, d$  and  $c$  such that*

$$d_K(\mathcal{L}(Z_N), \mathcal{N}(0, 1)) \leq C \max(1, \rho^3) N^{-\frac{1}{2}} \ln^2(N + 1) \quad (2.9)$$

for any  $N \in \mathbb{N}$ , where  $\rho = MD^{-1}$ . Moreover, for any  $\epsilon > 0$  there exists a constant  $c_\epsilon > 0$  which depends only on  $\epsilon, c, d$  and  $\ell$  so that for any  $N \geq 1$ ,

$$d_K(\mathcal{L}(N^{-\frac{1}{2}} S_N), \mathcal{N}(0, D^2)) \leq c_\epsilon \max(1, \rho^3) N^{-\frac{1}{2} + \epsilon} \quad (2.10)$$

where  $\mathcal{N}(0, D^2)$  is the zero mean normal distribution with variance  $D^2$ . When  $\beta_\infty(r_0) = 0$  for some  $r_0$  then (2.9) and (2.10) hold true with constants  $C$  and  $c_\epsilon$  which depend also on  $r_0$ , assuming only that  $F$  is a bounded function satisfying (2.4).

Note that  $\beta_\infty(0) = 0$  when  $\mathcal{F}_{m,n} = \sigma\{\xi_{\max(0,m)}, \dots, \xi_{\max(0,n)}\}$  and therefore when the processes  $\{\xi_n : n \geq 0\}$  itself is exponentially fast  $\phi$ -mixing (i.e. when (2.7) holds true with these  $\sigma$ -algebras) we obtain (2.9) for any bounded function  $F$ .

The outline of the proof is as follows. Relying on [13], Stein's method becomes effective for the sum  $S_N$  when  $\{F_n : 1 \leq n \leq N\}$ ,  $F_n = F(\xi_n, \xi_{2n}, \dots, \xi_{\ell n})$  are locally weak dependent in the sense that there exist sets  $A_n$  and nonnegative integers  $d_{n,m}$ ,  $1 \leq n, m \leq N$  so that  $n \in A_n$ ,  $a_n = |A_n|$  and  $b_n(k) = |\{1 \leq m \leq N : d_{n,m} = k\}|$ ,  $k \geq 0$  are small relatively to  $N$ ,  $F_n$  and  $\{F_s : s \notin A_n\}$  are weakly dependent and the random vectors  $\mathbf{F}_n = \{F_k : k \in A_n\}$  and  $\mathbf{F}_m = \{F_s : s \in A_m\}$  are weakly dependent when  $d_{n,m}$  is sufficiently large. We first reduce the problem of approximation of the left hand side of (2.9) to the case when  $\xi = \{\xi_n : n \geq 0\}$  forms a sufficiently fast  $\phi$ -mixing process. Then we consider the sets

$$A_n = A_{n,N,\ell} = \{1 \leq m \leq N : \min_{1 \leq i,j \leq \ell} |in - jm| \leq \ell\}$$

and the numbers  $d_{n,m} = \min\{|ia - jb| : a \in A_n, b \in A_m, 1 \leq i, j \leq \ell\}$  and show that  $a_n$  and  $b_n(k)$  defined above are of order  $l$ . In Section 3 we provide estimates which will show that the required type of the above weak dependence is satisfied, and then we take  $l$  of the form  $l = A \ln(N+1)$  to complete the proof. In fact, existing estimates on the left hand side of (2.9) using Stein's method become effective only after using the expectation estimates obtained in Section 3 even for "conventional" sums of  $\phi$ -mixing sequences (i.e. in the case  $\ell = 1$ ), which is a particular case of our setup, and so, in particular, we show that Stein's method is effective for such sums and yields almost optimal convergence rate.

### 3 Auxiliary results

**Lemma 3.1.** *Let  $X$  and  $Y$  be two random variables defined on the same probability space. Let  $Z$  be a random variable with density  $\rho$  bounded from above by some constant  $c > 0$ . Then,*

$$d_K(\mathcal{L}(Y), \mathcal{L}(Z)) \leq 3d_K(\mathcal{L}(X), \mathcal{L}(Z)) + 4c\|X - Y\|_{L^\infty} \quad \text{and for any } b \geq 1,$$

$$d_K(\mathcal{L}(Y), \mathcal{L}(Z)) \leq 3d_K(\mathcal{L}(X), \mathcal{L}(Z)) + (1 + 4c)\|X - Y\|_{L^b}^{1 - \frac{1}{b+1}}.$$

The second inequality is proved in Lemma 3.3 in [8], while the proof of the first inequality goes in the same way as the proof of that Lemma 3.3, taking in (3.2) from there  $\delta = \|X - Y\|_{L^\infty}$ .

Next, we recall that (see [2], Ch. 4) for any two sub- $\sigma$ -algebras  $\mathcal{G}, \mathcal{H} \subset \mathcal{F}$ ,

$$2\phi(\mathcal{G}, \mathcal{H}) = \sup\{\|\mathbb{E}[g|\mathcal{G}] - \mathbb{E}g\|_{L^\infty} : g \in L^\infty(\Omega, \mathcal{H}, P), \|g\|_{L^\infty} \leq 1\}. \quad (3.1)$$

The following lemma does not seem to be new but for readers' convenience and completeness we will sketch its proof here.

**Lemma 3.2.** *Let  $\mathcal{G}_1, \mathcal{G}_2 \subset \mathcal{F}$  be two sub- $\sigma$ -algebras of  $\mathcal{F}$  and for  $i = 1, 2$  let  $V_i$  be a  $\mathbb{R}^{d_i}$ -valued random  $\mathcal{G}_i$ -measurable vector with distribution  $\mu_i$ . Set  $d = d_1 + d_2$ ,  $\mu = \mu_1 \times \mu_2$ , denote by  $\kappa$  the distribution of the random vector  $(V_1, V_2)$  and consider the measure  $\nu = \frac{1}{2}(\kappa + \mu)$ . Let  $\mathcal{B}$  be the Borel  $\sigma$ -algebra on  $\mathbb{R}^d$  and  $H \in L^\infty(\mathbb{R}^d, \mathcal{B}, \nu)$ . Then  $\mathbb{E}[H(V_1, V_2)|\mathcal{G}_1]$  and  $\mathbb{E}H(v, V_2)$  exist for  $\mu_1$ -almost any  $v \in \mathbb{R}^{d_1}$  and*

$$|\mathbb{E}[H(V_1, V_2)|\mathcal{G}_1] - h(V_1)| \leq 2\|H\|_{L^\infty(\mathbb{R}^d, \mathcal{B}, \nu)}\phi(\mathcal{G}_1, \mathcal{G}_2), \quad P - a.s. \quad (3.2)$$

where  $h(v) = \mathbb{E}H(v, V_2)$  and a.s. stands for almost surely.

*Proof.* Clearly  $H$  is bounded  $\mu$  and  $\kappa$  a.s.. Thus  $\mathbb{E}[H(V_1, V_2)|\mathcal{G}_1]$  exists and existence of  $\mathbb{E}H(v, V_2)$  ( $\mu_1$ -a.s.) follows from the Fubini theorem. Relying on (3.1), inequality (3.2) follows easily for functions of the form  $G(v_1, v_2) = \sum_i I(v_1 \in A_i)g_i(v_2)$  where  $\{A_i\}$  is a measurable partition of the support of  $\mu_1$ . Any uniformly continuous function  $H$  is a uniform limit of functions of the above form, which implies that (3.2) holds true for uniformly continuous functions. Finally, by Lusin's theorem (see [14]), any function  $H \in L^\infty(\mathbb{R}^d, \mathcal{B}, \nu)$  is an  $L^1$  (and a.s.) limit of a sequence  $\{H_n\}$  of continuous functions with compact support satisfying  $\|H_n\|_{L^\infty(\mathbb{R}^d, \mathcal{B}, \nu)} \leq \|H\|_{L^\infty(\mathbb{R}^d, \mathcal{B}, \nu)}$  and (3.2) follows for any  $H \in L^\infty(\mathbb{R}^d, \mathcal{B}, \nu)$ .  $\square$

**Corollary 3.3.** *Let  $U_i$  be a  $d_i$ -dimensional random vector,  $i = 1, \dots, k$  defined on the probability space  $(\Omega, \mathcal{F}, P)$  from Section 2. Suppose that each  $U_i$  is  $\mathcal{F}_{m_i, n_i}$ -measurable, where  $n_{i-1} < m_i \leq n_i < m_{i+1}$ ,  $i = 1, \dots, k$ ,  $n_0 = -\infty$  and  $m_{k+1} = \infty$ . Let  $\{C_i : 1 \leq i \leq s\}$*

be a partition of  $\{1, 2, \dots, k\}$ . Denote by  $\mu_i$  the distribution of the random vector  $U(\mathcal{C}_i) = \{U_j : j \in \mathcal{C}_i\}$ ,  $i = 1, \dots, s$ . Then, for any bounded Borel function  $H : \mathbb{R}^{d_1+d_2+\dots+d_k} \rightarrow \mathbb{R}$ ,

$$\begin{aligned} & \left| \mathbb{E}H(U_1, U_2, \dots, U_k) - \int H(u_1, u_2, \dots, u_k) d\mu_1(u^{(\mathcal{C}_1)}) d\mu_2(u^{(\mathcal{C}_2)}) \dots d\mu_s(u^{(\mathcal{C}_s)}) \right| \\ & \leq 4\|H\|_\infty \sum_{i=2}^k \phi(m_i - n_{i-1}) \end{aligned} \quad (3.3)$$

where  $u^{(\mathcal{C}_i)} = \{u_j : j \in \mathcal{C}_i\}$ ,  $i = 1, \dots, s$  and  $\|H\|_\infty$  stands for the supremum of  $|H|$ . Namely, let  $U^{(i)}(\mathcal{C}_i)$  be independent copies of the processes  $U(\mathcal{C}_i)$ ,  $i = 1, \dots, s$ . Then

$$\left| \mathbb{E}H(U_1, U_2, \dots, U_k) - \mathbb{E}H(U_1^{(j_1)}, U_2^{(j_2)}, \dots, U_k^{(j_k)}) \right| \leq 4\|H\|_\infty \sum_{i=2}^k \phi(m_i - n_{i-1})$$

where  $j_i$  satisfies that  $i \in \mathcal{C}_{j_i}$ , for any  $1 \leq i \leq k$ .

*Proof.* Denote by  $\nu_i$  the distribution of  $U_i$ ,  $i = 1, \dots, k$ . We first prove by induction on  $k$  that for any choice of  $H$  and  $U_i$ 's with the required properties,

$$\begin{aligned} & \left| \mathbb{E}H(U_1, U_2, \dots, U_k) - \int H(u_1, u_2, \dots, u_k) d\nu_1(u_1) d\nu_2(u_2) \dots d\nu_k(u_k) \right| \\ & \leq 2\|H\|_\infty \sum_{i=2}^k \phi(m_i - n_{i-1}). \end{aligned} \quad (3.4)$$

Indeed, suppose that  $k = 2$  and set  $V_1 = U_1$ ,  $V_2 = U_2$ ,  $h(u_1) = E[H(u_1, U_2)]$ ,  $\mathcal{G}_1 = \mathcal{F}_{-\infty, n_1}$  and  $\mathcal{G}_2 = \mathcal{F}_{m_2, \infty}$ . Taking expectation in (3.2) yields

$$|\mathbb{E}H(U_1, U_2) - \mathbb{E}h(U_1)| \leq 2\|H\|_\infty \phi(m_2 - n_1)$$

which means that (3.4) holds true when  $k = 2$ . Now, suppose that (3.4) holds true for any  $k \leq j - 1$ ,  $U_1, \dots, U_k$  with the required properties and any bounded Borel function  $H : \mathbb{R}^{e_1+\dots+e_{k-1}} \rightarrow \mathbb{R}$ , where  $e_1, \dots, e_{k-1} \in \mathbb{N}$ . In order to deduce (3.4) for  $k = j$ , set  $V_1 = (U_1, \dots, U_{j-1})$ ,  $V_2 = U_j$ ,  $h(v_1) = \mathbb{E}H(v_1, U_j)$ ,  $v_1 = (u_1, \dots, u_{j-1})$ ,  $\mathcal{G}_1 = \mathcal{F}_{-\infty, n_{j-1}}$  and  $\mathcal{G}_2 = \mathcal{F}_{m_j, \infty}$ . Taking expectation in (3.2) yields

$$|\mathbb{E}H(U_1, U_2, \dots, U_j) - \mathbb{E}h(U_1, U_2, \dots, U_{j-1})| \leq 2\|H\|_\infty \phi(m_j - n_{j-1}).$$

Applying the induction hypothesis with the function  $h$  completes the proof of (3.4), since  $\|h\|_\infty \leq \|H\|_\infty$ . Next, we prove by induction on  $s$  that for any choice of  $k$ ,  $H$ ,  $U_i$ 's with the required properties and  $\mathcal{C}_1, \dots, \mathcal{C}_s$ ,

$$\begin{aligned} & \left| \int H(u_1, u_2, \dots, u_k) d\mu_1(u^{(\mathcal{C}_1)}) d\mu_2(u^{(\mathcal{C}_2)}) \dots d\mu_s(u^{(\mathcal{C}_s)}) - \right. \\ & \left. \int H(u_1, u_2, \dots, u_k) d\nu_1(u_1) d\nu_2(u_2) \dots d\nu_k(u_k) \right| \leq 2\|H\|_\infty \sum_{i=2}^k \phi(m_i - n_{i-1}). \end{aligned} \quad (3.5)$$

For  $s = 1$  this is just (3.4). Now suppose that (3.5) holds true for any  $s \leq j - 1$ , and any real valued bounded Borel function  $H$  defined on  $\mathbb{R}^{d_1+\dots+d_k}$ , where  $k$  and  $d_1, \dots, d_k$  are some natural numbers. In order to prove (3.5) for  $s = j$ , set  $u^{(I)} = (u^{(\mathcal{C}_1)}, u^{(\mathcal{C}_2)}, \dots, u^{(\mathcal{C}_{s-1})})$  and let the function  $I$  be defined by

$$I(u^{(I)}) = \int H(u_1, u_2, \dots, u_k) \prod_{j \in \mathcal{C}_s} d\nu_j(u_j). \quad (3.6)$$

Then

$$\int H(u_1, u_2, \dots, u_k) d\nu_1(u_1) d\nu_2(u_2) \dots d\nu_k(u_k) = \int I(u^{(I)}) \prod_{j \notin \mathcal{C}_s} d\nu_j(u_j). \quad (3.7)$$

Let the function  $J$  be defined by

$$J(u^{(I)}) = \int H(u_1, u_2, \dots, u_k) d\mu_s(u^{(C_s)}). \quad (3.8)$$

Then by (3.4), for any  $u^{(C_1)}, \dots, u^{(C_{s-1})}$ ,

$$|I(u^{(I)}) - J(u^{(I)})| \leq 2\|H\|_\infty \sum_{i \in C_s} \phi(m_i - n_{i-1}). \quad (3.9)$$

It is clear that  $\|J\|_\infty \leq \|H\|_\infty$ . Applying the induction hypothesis with the function  $J$  (considered as a function of the variable  $u$ ) and taking into account (3.7) and (3.9) we obtain (3.5) with  $s = j$ . We have completed the induction. Inequality (3.3) follows by (3.4) and (3.5), and the proof of Corollary 3.3 is complete.  $\square$

**Remark 3.4.** In the notations of Corollary 3.3, let  $Z_i, i = 1, \dots, s$  be a bounded  $\sigma\{U(C_i)\}$ -measurable random variable. Then each  $Z_i$  has the form  $Z_i = H_i(U(C_i))$  for some function  $H_i$  which satisfies  $\|H_i\|_\infty \leq \|Z_i\|_{L^\infty}$ . By considering the function  $H(u) = \prod_{i=1}^s H_i(u^{(C_i)})$ , we obtain from (3.3) that,

$$|\mathbb{E}[\prod_{i=1}^s Z_i] - \prod_{i=1}^s \mathbb{E}Z_i| \leq 4(\prod_{i=1}^s \|Z_i\|_{L^\infty}) \sum_{j=2}^k \phi(m_j - n_{j-1}). \quad (3.10)$$

In general we can replace  $\|H\|_\infty$  appearing in the right hand side of (3.3) by some essential supremum norm of  $H$  with respect to some measure which has a similar but more complicated form as  $\kappa$  defined in Lemma 3.2.

## 4 Nonconventional CLT with almost optimal convergence rate via Stein's method

First, the proof of Theorem 2.2 follows from arguments in [11], [10], and [7]. Indeed, relying on (2.25) in [11], the conditions of [7] and [10] hold true in our circumstances. Existence of  $D^2$  follows from Theorem 2.2 in [11], inequality (2.8) follows from the arguments in [10] (first by considering the case when  $M = 1$ ) and the condition for positivity follows from Theorem 2.3 in [7].

Before proving Theorem 2.3 we introduce the following notations. For any  $a, b \in \mathbb{R}$  set

$$d_\ell(a, b) = \min_{1 \leq i, j \leq \ell} |ia - jb|$$

and for any  $A, B \subset \mathbb{R}$  set

$$\text{dist}(A, B) = \inf\{|a - b| : a \in A, b \in B\} \text{ and } d_\ell(A, B) = \inf\{d_\ell(a, b) : a \in A, b \in B\}.$$

Finally, for any  $A_1, A_2, \dots, A_L \subset \mathbb{R}$ , we will write  $A_1 < A_2 < \dots < A_L$  if  $a_1 < a_2 < \dots < a_L$  for any  $a_i \in A_i, i = 1, 2, \dots, L$ .

### Proof of Theorem 2.3

Suppose that  $D^2 > 0$ . We consider first the self normalized case. Clearly, in the proof of (2.9) we can assume that  $M = 1$ . For any  $N \geq 1$  set  $s_N = \sqrt{\mathbb{E}(S_N)^2}$ . Then by (2.8),

$$(s_N)^2 \geq D^2 N - N^{\frac{1}{2}} C_1. \quad (4.1)$$

Let  $N$  be so that  $D^2 N > N^{\frac{1}{2}} C_1$ . Then  $s_N > 0$  and we set  $Z_N = \frac{S_N}{s_N}$ . Let  $l$  be of the form  $l = 4A \ln(N + 1)$  where  $A \geq 1$  is a positive constant considered here as a parameter which will be chosen later. Set  $r = \lfloor \frac{l}{4} \rfloor$  and

$$S_{N,r} = \sum_{n=1}^N F(\xi_{n,r}, \xi_{2n,r}, \dots, \xi_{\ell n,r})$$

where  $\xi_{m,r} = \mathbb{E}[\xi_m | \mathcal{F}_{m-r, m+r}]$  for any  $m \in \mathbb{N}$ . Then by (2.5) and (2.7),

$$\|S_N - S_{N,r}\|_{L^\infty} \leq \ell N (\beta_\infty(r))^\kappa \leq d\ell c^{-1} N c^{\frac{1}{4}} = d\ell c^{-1} N c^{A \ln(N+1)} \leq c_0 (N+1)^{1+A \ln c} \quad (4.2)$$

where  $c_0 = d\ell c^{-1}$  and we also used our assumption that  $M = 1$ . Next, let  $n > l$ , consider the random vectors  $U_i = \xi_{in,r}$  and set  $m_i = in - r$  and  $n_i = in + r$ ,  $i = 1, \dots, \ell$ . Then each  $U_i$  is  $\mathcal{F}_{in-r, in+r}$ -measurable and  $m_i - n_{i-1} = n - 2r \geq l - 2r \geq \frac{l}{2}$ . Applying Corollary 3.3 with the sets  $\mathcal{C}_i = \{in\}$ ,  $i = 1, \dots, \ell$  we obtain

$$\begin{aligned} & |\mathbb{E}F(\xi_{n,r}, \xi_{2n,r}, \dots, \xi_{\ell n,r}) - \mathbb{E}F(\xi_{n,r}^{(1)}, \xi_{2n,r}^{(2)}, \dots, \xi_{\ell n,r}^{(\ell)})| \\ & \leq 4\ell \phi\left(\frac{l}{2}\right) \leq 4\ell d c^{\frac{1}{2}} = c_1 (N+1)^{2A \ln c} \end{aligned}$$

where  $c_1 = 4\ell d$  and  $\xi_{in,r}^{(i)}$ 's are independent copies of  $\xi_{in,r}$ 's. Considering the product measure of the laws of the vectors  $(\xi_{in,r}, \xi_{in})$ ,  $i = 1, \dots, \ell$ , we can always assume that there (on a larger probability space) exist independent copies  $\xi_{in}^{(i)}$  of the  $\xi_{in}$ 's such that  $\|\xi_{in}^{(i)} - \xi_{in,r}^{(i)}\|_{L^\infty} \leq \beta_\infty(r)$  for any  $i = 1, 2, \dots, \ell$ . Thus by (2.5) and (2.7),

$$\begin{aligned} & |\mathbb{E}F(\xi_{n,r}^{(1)}, \xi_{2n,r}^{(2)}, \dots, \xi_{\ell n,r}^{(\ell)}) - \mathbb{E}F(\xi_n^{(1)}, \xi_{2n}^{(2)}, \dots, \xi_{\ell n}^{(\ell)})| \\ & \leq \ell (\beta_\infty(r))^\kappa \leq \ell d c^{-1} c^{\frac{1}{4}} = c_0 (N+1)^{A \ln c} \end{aligned}$$

and notice that  $\mathbb{E}F(\xi_n^{(1)}, \xi_{2n}^{(2)}, \dots, \xi_{\ell n}^{(\ell)}) = \bar{F} = 0$ . We conclude from (2.4) and the above estimates that

$$\begin{aligned} |\mathbb{E}S_{N,r}| & \leq |\mathbb{E}S_{l,r}| + N(4\ell d c^{\frac{1}{2}} + d\ell c^{-1} c^{\frac{1}{4}}) \\ & \leq 2l + 5N\ell d c^{-1} c^{\frac{1}{4}} \leq 8A \ln(N+1) + 5c_0 (N+1)^{1+A \ln c}. \end{aligned} \quad (4.3)$$

We assume henceforth that  $-A \ln c = A |\ln c| > 2$  and set  $\bar{S}_{N,r} = S_{N,r} - \mathbb{E}S_{N,r}$ . For any two random variables  $X$  and  $Y$  defined on the same probability space we have  $|\mathbb{E}X^2 - \mathbb{E}Y^2| \leq \|X + Y\|_{L^2} \|X - Y\|_{L^2}$  and therefore by (4.2) and (2.8),

$$\begin{aligned} |\mathbb{E}S_N^2 - \mathbb{E}S_{N,r}^2| & \leq (2\|S_N\|_2 + c_0 (N+1)^{1+A \ln c}) c_0 (N+1)^{1+A \ln c} \\ & \leq 3c_0 (2 + c_0 + C_1 + D) (N+1)^{\frac{3}{2} + A \ln c} \end{aligned}$$

where we also used that  $A |\ln c| > 1$ . Next, by (4.3),

$$\begin{aligned} |\mathbb{E}S_{N,r}^2 - \mathbb{E}\bar{S}_{N,r}^2| & = |\mathbb{E}S_{N,r}^2 - \text{Var}S_{N,r}| = |\mathbb{E}S_{N,r}|^2 \\ & \leq 32A^2 \ln^2(N+1) + 25c_0^2 (N+1)^{2+2A \ln c} \end{aligned}$$

and together with the previous inequality and our assumption that  $A \ln c < -2$  we obtain that

$$|\mathbb{E}S_N^2 - \mathbb{E}\bar{S}_{N,r}^2| \leq c_2 \ln^2(N+1) \quad (4.4)$$

where  $c_2 = 32A^2 + 25c_0^2 + 3c_0(2 + c_0 + C_1 + D)$ . Combining this with (4.1), it follows that

$$\mathbb{E}\bar{S}_{N,r}^2 \geq D^2 N - N^{\frac{1}{2}} C_1 - c_2 \ln^2(N+1). \quad (4.5)$$

Let  $N$  be so large such that the right hand side in the previous inequality is positive. Then  $\mathbb{E}\bar{S}_{N,r}^2 > 0$ . Let  $\bar{s}_{N,r}$  be its positive square root and set

$$\bar{Z}_{N,r} = \frac{\bar{S}_{N,r}}{\bar{s}_{N,r}} = \sum_{n=1}^N Y_n$$



where for each  $n$ ,

$$Y_n = Y_{n,N,r} = \frac{F(\xi_{n,r}, \xi_{2n,r}, \dots, \xi_{\ell n,r}) - \mathbb{E}F(\xi_{n,r}, \xi_{2n,r}, \dots, \xi_{\ell n,r})}{\bar{s}_{N,r}}.$$

Observe now that

$$\begin{aligned} \|Z_N - \bar{Z}_{N,r}\|_{L^\infty} &\leq \|(s_N)^{-1}S_N - (\bar{s}_{N,r})^{-1}S_{N,r}\|_{L^\infty} + |(\bar{s}_{N,r})^{-1}\mathbb{E}S_{N,r}| \\ &\leq (s_N)^{-1}\|S_N - S_{N,r}\|_{L^\infty} + |(s_N)^{-1} - (\bar{s}_{N,r})^{-1}|\|S_{N,r}\|_{L^\infty} + |(\bar{s}_{N,r})^{-1}\mathbb{E}S_{N,r}|. \end{aligned}$$

The inequality  $|x^{-1} - y^{-1}| = |x^2 - y^2|(x^2y + y^2x)^{-1}$  holds true for any  $x, y > 0$  yielding that

$$|(s_N)^{-1} - (\bar{s}_{N,r})^{-1}| \leq \frac{c_2 \ln^2(N+1)}{(s_N + \bar{s}_{N,r})s_N\bar{s}_{N,r}} := e_1$$

where we used (4.4), and we conclude from (4.2), (4.3) and the above estimates that

$$\begin{aligned} \|Z_N - \bar{Z}_{N,r}\|_{L^\infty} &\leq (s_N)^{-1}c_0(N+1)^{1+A \ln c} + 2Ne_1 \\ &\quad + (\bar{s}_{N,r})^{-1}(8A \ln(N+1) + 5c_0(N+1)^{1+A \ln c}) \end{aligned} \quad (4.6)$$

where we used that  $\|S_{N,r}\|_{L^\infty} \leq 2N$  (recall our assumption that  $M = 1$ ). Next, using (4.1), (4.5) and that  $\ln(N+1) \leq N^{\frac{1}{2}}$  for any  $N \geq 1$  we derive that  $\min(s_N^2, \bar{s}_{N,r}^2) \geq \frac{1}{4}D^2N$  when  $3N^{\frac{1}{2}}D^2 \geq 8(C_1 + c_2)$  and in this case

$$\|Z_N - \bar{Z}_{N,r}\|_{L^\infty} \leq c_4 \max(D^{-1}, D^{-3})N^{-\frac{1}{2}} \ln^2(N+1) \quad (4.7)$$

where  $c_4 = C_4(1 + c_0 + c_2 + A)$ ,  $C_4 > 1$  is some absolute constant and we also used that  $N^{1+A \ln c} < 1$ .

Next, let  $N$  be sufficiently large so that  $3N^{\frac{1}{2}}D^2 \geq 8(C_1 + c_2)$ . Then by (2.4) and the above lower bound of  $\bar{s}_{N,r}^2$ ,

$$\|Y_n\|_{L^\infty} \leq 2(\bar{s}_{N,r})^{-1} \leq 4D^{-1}N^{-\frac{1}{2}}. \quad (4.8)$$

For any  $n = 1, 2, \dots, N$  set

$$A_n = A_{n,l,N} = \{1 \leq m \leq N : \min_{1 \leq i, j \leq \ell} |in - jm| \leq l\} = \{1 \leq m \leq N : d_\ell(n, m) \leq l\}$$

and for any  $k \geq 0$  set

$$\mathcal{A}_n(k) = \{1 \leq m \leq N : d_\ell(A_n, A_m) = k\}.$$

We claim that there exist constants  $K_1$  and  $K_2$  which depend only on  $\ell$  such that for any  $n = 1, 2, \dots, N$  and  $k \geq 0$ ,

$$|A_n| \leq K_1 l \quad \text{and} \quad |\mathcal{A}_n(k)| \leq K_2 l. \quad (4.9)$$

Indeed, since  $A_n$  is contained in a union of at most  $\ell^2$  intervals whose lengths do not exceed  $2l + 1$  we have  $|A_n| \leq \ell^2(2l + 1)$  and since  $l \geq 1$  we can take  $K_1 = 3\ell^2$ . To prove the second inequality, let  $n$  and  $m$  be such that  $d_\ell(A_n, A_m) = k \geq 0$ . Then there exist  $1 \leq i_s, j_s \leq \ell$ ,  $s = 1, 2, 3$  and  $1 \leq u, v \leq N$  such that  $|i_3u - j_3v| = k$ ,  $|i_1n - j_1u| \leq l$  and  $|i_2m - j_2v| \leq l$ . When  $j_3v - i_3u_3 = k$ , we deduce from the last two inequalities that

$$\left| m - \frac{j_2i_3i_1}{i_2j_3j_1}n - \frac{j_2}{j_3i_2}k \right| \leq l \left( \frac{1}{i_2} + \frac{j_2i_3}{i_2j_3j_1} \right)$$

and similar inequality holds when  $j_3v - i_3u_3 = -k$ . Thus, when  $n$  and  $k$  are fixed the set  $\mathcal{A}_n(k)$  is contained in a union of  $2\ell^6$  intervals whose lengths do not exceed  $2(\ell^2 + 1)l$ , and the choice of  $K_2 = 4\ell^6 \cdot (\ell^2 + 2)$  is sufficient.

Now, set  $\delta = \delta_{l,N} = \sum_{n=1}^N \sum_{m \in A_n} \mathbb{E} Y_n Y_m$ . Then

$$1 = \text{Var} \bar{Z}_{N,r} = \mathbb{E} \left( \sum_{n=1}^N Y_n \right)^2 = \delta + \gamma \quad (4.10)$$

where  $\gamma = \gamma_{l,N} = \sum_{n=1}^N \sum_{m \in \{1, \dots, N\} \setminus A_n} \mathbb{E} Y_n Y_m$ . Let  $1 \leq n, m \leq N$  be such that  $m \notin A_n$ . Consider the sets of indexes  $\Gamma_k = \{jn : 1 \leq j \leq \ell\}$  where  $k = n, m$  and set  $\Gamma_{n,m} = \Gamma_n \cup \Gamma_m$ . By the definition of the set  $A_n$  we have  $\text{dist}(\Gamma_n, \Gamma_m) = d_\ell(n, m) > l$ . Therefore, the set  $\Gamma_{n,m}$  can be represented in the form

$$\Gamma_{n,m} = \bigcup_{t=1}^L B_t, \quad B_1 < B_2 < \dots < B_L$$

where  $L \leq 2\ell$ , each  $B_t$  is either a subset of  $\Gamma_n$  or of  $\Gamma_m$  and  $\text{dist}(B_t, B_{t-1}) > l$ . Set

$$U_t = \{\xi_{s,r} : s \in B_t\}, \quad t = 1, \dots, L.$$

Since  $r \leq \frac{l}{4}$ , there exist numbers  $n_t$  and  $m_t$ ,  $t = 1, \dots, L$  such that  $n_{t-1} < m_t \leq n_t < m_{t+1} + \frac{l}{2}$ , where  $n_0 = -\infty, m_{L+1} := \infty$ , and each  $U_t$  is measurable with respect to  $\mathcal{F}_{m_t, n_t}$ . Set  $\mathcal{C}_1 = \{1 \leq t \leq L : B_t \subset \Gamma_n\}$  and  $\mathcal{C}_2 = \{1 \leq t \leq L : B_t \subset \Gamma_m\}$ . Then  $\{\mathcal{C}_1, \mathcal{C}_2\}$  is a partition of  $\{1, 2, \dots, L\}$ ,  $Y_n$  is measurable with respect to  $\sigma\{U_t : t \in \mathcal{C}_1\}$  and  $Y_m$  is measurable with respect to  $\sigma\{U_t : t \in \mathcal{C}_2\}$ . Therefore, by (3.10) and (4.8) and since  $\mathbb{E} Y_n = 0$ ,

$$|\mathbb{E} Y_n Y_m| \leq 64\ell N^{-1} D^{-2} \phi\left(\frac{l}{2}\right) \leq 64d\ell D^{-2} N^{-1} c^{\frac{1}{2}} \leq 64d\ell D^{-2} N^{2A \ln c - 1}$$

implying that

$$|\gamma| = |\delta - 1| = |\delta - \text{Var} \bar{Z}_{N,r}| \leq 64d\ell D^{-2} N^{1+2A \ln c}. \quad (4.11)$$

We assume now, in addition to the previous restriction on  $N$ , that  $64d\ell D^{-2} N^{-\frac{1}{2}} < \frac{1}{2}$ . Then  $\delta > \frac{1}{2}$  and so we can set  $\sigma = \sqrt{\delta}$  and  $W = \frac{\bar{Z}_{N,r}}{\sigma}$ . Then  $\sigma^2 \geq \frac{1}{2}$  and using (4.11) we obtain

$$\|W - \bar{Z}_{N,r}\|_{L^\infty} \leq \|\bar{Z}_{N,r}\|_{L^\infty} \left|1 - \frac{1}{\sigma}\right| \leq 4\|\bar{Z}_{N,r}\|_{L^\infty} |\delta - 1| \leq 16D^{-3} N^{\frac{3}{2} + 2A \ln c} \quad (4.12)$$

where we also used that  $\bar{s}_{N,r} \geq \frac{1}{2} DN^{\frac{1}{2}}$ . Since  $A|\ln c| > 1$  the above right hand side does not exceed  $16D^{-3} N^{-\frac{1}{2}}$  which together with (4.7) and Lemma 3.1 yields that

$$d_K(\mathcal{L}(Z_N), \mathcal{N}(0, 1)) \leq 3d_K(\mathcal{L}(W), \mathcal{N}(0, 1)) + c_5 \max(D^{-1}, D^{-3}) N^{-\frac{1}{2}} \ln^2(N+1) \quad (4.13)$$

where  $c_5 = 16c_4$ .

In order to approximate  $d_K(\mathcal{L}(W), \mathcal{N}(0, 1))$ , set  $X_n = \sigma^{-1} Y_n$ ,  $n = 1, 2, \dots, N$ . Then  $W = \sum_{n=1}^N X_n$  and by (4.8) we have  $\|X_n\|_{L^\infty} \leq R$ , where  $R = 4N^{-\frac{1}{2}} D^{-1} \sigma^{-1} \leq 8N^{-\frac{1}{2}} D^{-1}$ . Applying Theorem 2.1 in [13], using the equality (15) from there and taking into account (4.9) we obtain that

$$d_K(\mathcal{L}(W), \mathcal{N}(0, 1)) \leq R_1 + R_2 + R_3 + K_1 l R + 2K_1^2 l^2 N R^3$$

where

$$\begin{aligned} R_1 &= 4 \left\| \sum_{n=1}^N \sum_{m \in A_n} (X_n X_m - \mathbb{E} X_n X_m) \right\|_2, \\ R_2 &= \sqrt{2\pi} \sum_{n=1}^N \mathbb{E} |\mathbb{E}[X_n | X_m : m \notin A_n]|, \\ R_3 &= 2 \left\| \sum_{n=1}^N X_n \left( \sum_{m \in A_n} X_m \right)^2 \right\|_2 (\|W\|_2 + 5) \end{aligned}$$

and  $\|X\|_q^q = \mathbb{E}|X|^q = \|X\|_{L^q}^q$  for any random variable  $X$ . Now we estimate  $R_1, R_2$  and  $R_3$ . Set  $T_n = \sum_{m \in A_n} (X_n X_m - \mathbb{E}X_n X_m)$ ,  $n = 1, \dots, N$ . Then

$$R_1^2 = 16 \sum_{n_1, n_2=1}^N \mathbb{E}T_{n_1}T_{n_2}.$$

Let  $n_1$  and  $n_2$  be such that  $d_\ell(A_{n_1}, A_{n_2}) = k > 2r$ . Consider the sets  $\Gamma_s = \{jm : m \in A_{n_s}, 1 \leq j \leq \ell\}$ ,  $s = 1, 2$ . Then  $\text{dist}(\Gamma_1, \Gamma_2) = d_\ell(A_{n_1}, A_{n_2}) = k$ . Set  $\Gamma = \Gamma_1 \cup \Gamma_2$ . Both  $\Gamma_i$ 's are unions of at most  $\ell^3$  integer intervals (over the integers), and therefore there exist sets  $B_1, B_2, \dots, B_L$ ,  $L \leq 2\ell^3$  such that

$$\Gamma = \bigcup_{t=1}^L B_t, \quad B_1 < B_2 < \dots < B_L$$

where each  $B_t$  is either a subset of  $\Gamma_1$  or a subset of  $\Gamma_2$  and  $\text{dist}(B_t, B_{t-1}) \geq k$ ,  $t = 2, \dots, L$ . Set

$$U_t = \{\xi_{a,r} : a \in B_t\}, \quad t = 1, \dots, L.$$

Then there exist numbers  $m_t, n_t, t = 1, 2, \dots, L$  such that  $n_{t-1} < m_t \leq n_t \leq m_{t+1} + k - 2r$ ,  $n_0 = -\infty, m_{L+1} := \infty$  and each  $U_t$  is  $\mathcal{F}_{m_t, n_t}$ -measurable. Set  $\mathcal{C}_s = \{1 \leq t \leq L : B_t \subset \Gamma_s\}$ ,  $s = 1, 2$ . Then  $\{\mathcal{C}_1, \mathcal{C}_2\}$  is a partition of  $\{1, 2, \dots, L\}$  and  $T_{n_s}$ ,  $s = 1, 2$  is measurable with respect to  $\sigma\{U_t : t \in \mathcal{C}_s\}$ . Since  $\|X_n\|_{L^\infty} \leq R$  we have  $\|T_n\|_{L^\infty} \leq 2K_1 l R^2$  (recall (4.9)) and thus by (3.10),

$$|\mathbb{E}T_{n_1}T_{n_2}| \leq 16K_1^2 l^2 R^4 L^2 \phi(k - 2r) \leq 64\ell^6 K_1^2 l^2 R^4 d c^{k-2r}$$

where we used that  $\mathbb{E}T_n = 0$ . Given  $n_1$  and  $k > 2r$ , the number of  $n_2$ 's satisfying  $d_\ell(A_{n_1}, A_{n_2}) = k$  is at most  $K_2 l$  (recall (4.9)), while for any other  $n_2$  and  $k$  we can use the trivial upper bound  $|\mathbb{E}T_{n_1}T_{n_2}| \leq \|T_{n_1}\|_{L^\infty} \|T_{n_2}\|_{L^\infty} \leq 4K_1^2 l^2 R^4$ . Therefore, by the definitions of  $R$  and  $r$ ,

$$R_1^2 \leq 64\ell^4 K_1^2 l^2 R^4 N \left( K_2 l d \sum_{k=2r+1}^N c^{k-2r} + (2r+1)K_2 l \right) \leq C_0 l^4 N^{-1} D^{-4}$$

where  $C_0$  is a constant which depends only on  $c$  and  $d$  and  $\ell$ . In order to approximate  $R_2$ , let  $1 \leq n \leq N$  and set  $\mathcal{X}_n = \{X_m : m \notin A_n\}$ . Then,

$$\|\mathbb{E}[X_n | \mathcal{X}_n]\|_1^2 \leq \|\mathbb{E}[X_n | \mathcal{X}_n]\|_2^2 = \mathbb{E}X_n \mathbb{E}[X_n | \mathcal{X}_n]. \quad (4.14)$$

Consider the sets  $\tau_1 = \{n, 2n, \dots, \ell n\}$  and  $\tau_2 = \{jm : m \notin A_n, 1 \leq j \leq \ell\}$ . Then by the definition of  $A_n$  we have  $\text{dist}(\tau_1, \tau_2) > l$ . Thus, the union  $\tau_1 \cup \tau_2$  can be written as a union of at most  $2\ell + 1$  disjoint sets  $B_1, B_2, \dots, B_L$  such that  $B_1 < B_2 < \dots < B_L$ ,  $\text{dist}(B_t, B_{t+1}) > l$  and each  $B_t$  is either a subset of  $\tau_1$  or a subset of  $\tau_2$ . Consider the random vectors

$$U_t = \{\xi_{s,r} : s \in B_t\}, \quad t = 1, \dots, L$$

and the partition of  $\{1, 2, \dots, L\}$  into the sets  $\{\mathcal{C}_1, \mathcal{C}_2\}$ , where  $\mathcal{C}_s = \{1 \leq t \leq L : B_t \subset \tau_s\}$ ,  $s = 1, 2$ . Then  $X_n$  is measurable with respect to  $\sigma\{U_t : t \in \mathcal{C}_1\}$  and  $\mathbb{E}[X_n | \mathcal{X}_n]$  is measurable with respect to  $\sigma\{U_t : t \in \mathcal{C}_2\}$ . Therefore by (3.10) and (2.7),

$$|\mathbb{E}[X_n \mathbb{E}[X_n | \mathcal{X}_n]]| \leq 4(2\ell + 1)R^2 d c^{\frac{l}{2}} \quad (4.15)$$

where we used that  $r \leq \frac{l}{4}$ ,  $\mathbb{E}X_n = 0$  and that  $\|\mathbb{E}[X_n | \mathcal{X}_n]\|_{L^\infty} \leq \|X_n\|_{L^\infty} \leq R$ . We conclude from (4.14) and (4.15) that there exists a constant  $C'_0$  which depends only on  $\ell$  such that

$$R_2 \leq C'_0 N^{\frac{1}{2}} D^{-1} d^{\frac{1}{2}} c^{\frac{l}{4}}. \quad (4.16)$$

To estimate  $R_3$ , first observe that by the definition of  $W$  and by (4.10) we have  $\|W\|_2^2 = \delta^{-1} \|\bar{Z}_{N,r}\|_2^2 = 1 + \delta^{-1}\gamma$  and therefore  $\|W\|_2 \leq 2$ , since  $|\gamma| < \frac{1}{2}$  and  $\delta > \frac{1}{2}$ . The first factor in the definition of  $R_3$  is clearly bounded from above by  $2NK_1^2 l^2 R^3$  and we conclude that

$$R_3 \leq C_4 l^2 D^{-3} N^{-\frac{1}{2}}$$

for some constant  $C_4$  which depends only on  $\ell$ . The estimate (2.9) in Theorem 2.3 follows now by taking any  $A > \max(1, 2|\ln c|^{-1})$ , using (4.13) and the above estimates of  $R_i$ 's. Note that all the approximations in this section hold true only for  $N$ 's satisfying  $3N \geq 8D^{-2}(C_1 + c_2)$  and  $64d\ell D^{-2}N^{-\frac{1}{2}} < \frac{1}{2}$ , but inequality (2.9) follows for all other  $N$ 's from the basic estimate  $d_K(\mathcal{L}(Z_N), \mathcal{N}(0, 1)) \leq 1$ . We also remark that when  $\beta_\infty(r_0) = 0$  for some  $r_0$  then taking  $r \geq r_0$  we get  $S_{N,r} = S_N$  and so there is no need for (2.5) to hold true.

Now we derive (2.10) where again it is sufficient to consider the case when  $M = 1$ . Let  $0 < \epsilon < \frac{1}{4}$ . First for any  $b > 1$ ,

$$\begin{aligned} \|D^{-1}N^{-\frac{1}{2}}S_N - Z_N\|_{L^b} &= \|S_N\|_{L^b} |N^{-\frac{1}{2}}D^{-1} - (s_N)^{-1}| \\ &= \|S_N\|_{L^b} \left| \frac{\mathbb{E}S_N^2 - D^2N}{D^2Ns_N + D(s_N)^2N^{\frac{1}{2}}} \right| \end{aligned}$$

where in the second equality we used that  $|x^{-1} - y^{-1}| = |x^2 - y^2|(xy^2 + yx^2)^{-1}$  for any  $x, y > 0$ . By Lemma 5.2 in [7] for any  $b > 1$  there exists a constant  $M_b$  which depends only on  $c, d, b$  and  $\ell$  so that  $\|S_N\|_{L^b} \leq M_b N^{\frac{1}{2}}$ . Using the previous estimates, for any  $N$  so that  $3N^{\frac{1}{2}}D^2 \geq 8(C_1 + c_2)$  and  $64d\ell D^{-2}N^{-\frac{1}{2}} < \frac{1}{2}$  we have  $s_N \geq \frac{1}{2}D$ . Therefore,

$$\|D^{-1}N^{-\frac{1}{2}}S_N - Z_N\|_{L^b} \leq 8D^{-3}C_1M_bN^{-\frac{1}{2}}$$

where and we also used (2.8). Applying the second statement of Lemma 3.1 with  $b = \frac{1}{2\epsilon} - 1$  and using (2.9) completes the proof of (2.10).  $\square$

#### 4.1 Extensions and remarks

##### Unbounded functions

Let  $M, \iota > 0$ ,  $\kappa \in (0, 1]$  and  $F : (\mathbb{R}^\varphi)^\ell \rightarrow \mathbb{R}$  be a function satisfying

$$\begin{aligned} |F(x)| &\leq M(1 + \sum_{i=1}^\ell |x_i|^\iota) \quad \text{and} \\ |F(x) - F(y)| &\leq M(1 + \sum_{i=1}^\ell |x_i|^\iota + |y_i|^\iota) \sum_{i=1}^\ell |x_i - y_i|^\kappa \end{aligned} \quad (4.17)$$

for any  $x = (x_1, \dots, x_\ell)$  and  $y = (y_1, \dots, y_\ell)$  in  $(\mathbb{R}^\varphi)^\ell$ . For any  $R > 0$  set  $F_R(x) = F(x)\mathbb{I}(|F(x)| \leq R)$ . Then, assuming that for some  $p > \iota + 1$ ,

$$\gamma_p = \|\xi_1\|_{L^{p^*}} < \infty$$

we can first approximate  $F(\xi_n, \xi_{2n}, \dots, \xi_{\ell n})$  by  $F_R(\xi_{n,r}, \xi_{2n,r}, \dots, \xi_{\ell n,r})$  in the  $L^p$ -norm and then use Lemma 3.3 3.1. Applying Theorem 2.3 with the function  $F_R$  and taking  $R$  with an appropriate dependence on  $N$  we obtain convergence rate of the form  $CN^{-\frac{1}{2} + \varepsilon_p}$ , where  $\varepsilon_p$  depends on  $p$  and satisfies  $\lim_{p \rightarrow \infty} \varepsilon_p = 0$ . In fact, similar type of rates can be obtained assuming only that  $\phi(n) + \beta_q(n) \leq dn^{-\theta}$  for some  $q, d, \theta > 0$ , where  $\beta_q$  is defined similarly to  $\beta_\infty$ , but with the  $L^q$  norm.

##### Nonlinear indexes

Let  $q_i, i = 1, \dots, \ell$  be strictly increasing functions satisfying  $q_i(\mathbb{N}) \subset \mathbb{N}$  which are ordered so that

$$q_1(n) < q_2(n) < \dots < q_\ell(n) \quad \text{for any sufficiently large } n.$$

Consider the sums

$$S_N = \sum_{n=1}^N F(\xi_{q_1(n)}, \xi_{q_2(n)}, \dots, \xi_{q_\ell(n)}).$$

The proof of Theorem 2.3 proceeds essentially in the same way when all  $q_i$ 's are linear. For more general  $q_i$ 's, set

$$A_n = \{1 \leq m \leq N : \min_{1 \leq i, j \leq \ell} |q_i(n) - q_j(m)| \leq l\}.$$

The proof of Theorem 2.3 will proceed similarly for the sums  $S_N$  if we show that the limit  $D^2 = \lim_{N \rightarrow \infty} N^{-\frac{1}{2}} \mathbb{E} S_N^2$  exists, obtain convergence rate towards it and upper bounds similar to the ones in (4.9). Suppose that  $q_1, \dots, q_k$  are linear, for some  $k < \ell$  and that  $q_j, j \geq k$  are not. When all  $q_i$ 's are polynomials, existence of  $D^2$  is proved in [8]. Though the limit  $D^2$  does not exist in general, if  $q_{j+1}$  grows faster than  $q_j$  for  $j > k$  in the sense of (2.11) in [11], then existence of  $D^2$  follows from Theorem 2.3 in [11]. Convergence rate towards  $D^2$  when  $q_i$ 's are polynomials can be obtained by proceeding similarly to the proof of Proposition 5.3 in [8]. If, instead,  $q_{j+1}(n^\alpha) - q_j(n)$  converges to  $\infty$  as  $n \rightarrow \infty$  for some  $0 < \alpha < 1$  and all  $j \geq k$ , then convergence rate towards  $D^2$  with some dependence on  $\alpha$  follows from the arguments in [11].

Each  $q_i(n)$  grows at least as fast as linearly which implies that  $|A_n|$  is of order  $l$ . When all  $q_i$ 's are polynomials of the same degree then the limit  $\lim_{n \rightarrow \infty} q_i^{-1}(q_j(n))/n$  exists for any  $1 \leq i, j \leq \ell$  and therefore the proof of the second upper bound in (4.9) proceeds in a similar way but with  $\tilde{d}_\ell(a, b) = \min_{1 \leq i, j \leq \ell} |q_i(a) - q_j(b)|$  in place of  $d_\ell(a, b)$ . When  $q_i$ 's do not necessarily have the same degree then beginning the summation in the definition of  $S_N$  from  $cN^\gamma$  for appropriate  $\gamma < 1$  and  $c > 0$ , guarantees that  $|q_i(n) - q_j(m)| > CN$  when  $\deg q_i \neq \deg q_j$  and  $cN^\gamma \leq n, m \leq N$ . Similar to the latter inequality is satisfied when  $\max(i, j) > k$  and  $q_s$  grows faster than  $q_{s-1}$  for  $s = k+1, \dots, \ell$  and so an appropriate version of (4.9) follows in this situation, as well.

**Remark 4.1.** Using arguments similar to the ones in [3], we can obtain (2.9) with  $N^{-\frac{1}{2}} S_N$  in place of  $Z_N$  when  $F$  satisfies (4.17), and in particular when  $F$  is a bounded Hölder continuous function. In fact, Stein's method also yields (under appropriate conditions) the nonconventional functional CLT, which was proved in [11] using martingale approximation. These results require a relatively long presentation and their proofs are not short, so they will appear elsewhere.

## References

- [1] E. Bolthausen, *Exact convergence rates in some martingale central limit theorems*, Ann. Probab. 10 (1982), 672-688. MR-0659537
- [2] R.C. Bradley, *Introduction to Strong Mixing Conditions*, Volume 1, Kendrick Press, Heber City, 2007. MR-2325295
- [3] L.H.Y Chen and Q.M. Shao, *Normal approximation under local dependence*, Ann. Probab. 32 (2004), 1985-2028. MR-2073183
- [4] H. Furstenberg, *Nonconventional ergodic averages*, Proc. Symp. Pure Math. 50 (1990), 43-56. MR-1067751
- [5] M.I. Gordin, *On the central limit theorem for stationary processes*, Soviet Math. Dokl. 10 (1969), 1174-1176. MR-0251785
- [6] P. Hall, C.C. Heyde, *Rates of convergence in the martingale central limit theorem*, Ann. Probab. 9 (1981) 395-404. MR-0614625
- [7] Y. Hafouta and Yu. Kifer, *Berry-Esseen type estimates for nonconventional sums*, Stoch. Proc. Appl. 126 (2016), 2430-2464. MR-3505233
- [8] Y. Hafouta and Yu. Kifer, *Nonconventional polynomial CLT*, Stochastics, 89 (2017), 550-591. MR-3590435

- [9] Yu. Kifer, *Nonconventional limit theorems*, Probab. Th. Rel. Fields, 148 (2010), 71-106. MR-2653222
- [10] Yu. Kifer, *Strong approximations for nonconventional sums and almost sure limit theorems*, Stochastic Process. Appl., 123 (2013), 2286-2302. MR-3038505
- [11] Yu. Kifer and S.R.S Varadhan, *Nonconventional limit theorems in discrete and continuous time via martingales*, Ann. Probab., 42 (2014), 649-688. MR-3178470
- [12] D.L. McLeish, *Invariance principles for dependent variables*, Z.Wahrscheinlichkeitstheorie und Verw. Gebiete 32 (1975), 165-178. MR-0388483
- [13] Y. Rinott, *On normal approximation rates for certain sums of dependent random variables*, J. Comput. Appl. Math., 55 (1994), 135-143. MR-1327369
- [14] W. Rudin *Real and Complex Analysis*, McGraw-Hill, New York, 1987. MR-0924157
- [15] C. Stein, *A bound for the error in the normal approximation to the distribution of a sum of dependent random variables*, Proc. Sixth Berkeley Symp. Math. Statist. Probab, 2 (1972), 583-602. Univ. California Press, Berkeley. MR-0402873
- [16] C. Stein, *Approximation Computation of Expectations*, IMS, Hayward, CA (1986).
- [17] N. Shiryaev, *Probability*, Springer-Verlag, Berlin, 1995. MR-0737192

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