

## ON GENERALIZED CUMULATIVE INFORMATION OF KULLBACK-LEIBLER TYPE

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**Abstract.** In this paper generalized versions of the empirical cumulative Kullback-Leibler information are introduced, together with their alternative representations. The generalization refers to both application of a weight function and the use of Tsallis extended logarithm in the original Kullback-Leibler information measure. The original measure was extensively studied in several papers. The first information measure proposed in this paper is the weighted version of the original Kullback-Leibler information, the second one implies the use of Tsallis extended logarithm, while the third one, combines the first and second. Some properties of the new measures are also discussed.

**Key words:** entropy, inaccuracy, cumulative Kullback-Leibler information, empirical cumulative Kullback-Leibler information, Tsallis logarithm.

### 1. INTRODUCTION

Along the years, many authors have proposed and studied entropy-related measures and used these concepts in applications. Among them, we underline the valuable results obtained by Di Crescenzo and Longobardi [1, 2, 3], M. Dumitrescu [4], M. Iosifescu [5], Kullback and Leibler [6], Park, Rao and Shin [8], V. Preda [9, 10], V. Preda, C. Balcau [11, 12], V. Preda, C. Balcau, D. Constantin and I.I. Panait [13], Rao, Chen and Vermuri [14].

The concept of differential entropy has been extended to the relative entropy, called Kullback-Leibler information [6], which represents a discrepancy between two distributions. S. Park, M. Rao, D. W. Shin [8] introduced the Kullback-Leibler cumulative information. A. Di Crescenzo and M. Longobardi [1] presented many properties of the cumulative and empirical cumulative Kullback-Leibler information. The authors used the above mentioned measures for different real-life applications.

In this paper, we aim to introduce generalized versions of the cumulative and empirical cumulative Kullback Leibler information. Section 2 presents the framework of the paper: general assumptions, definitions and notations that are needed to describe the newly proposed concepts. In Sections 3, 4 and 5, the weighted, Tsallis and Tsallis weighted version of cumulative and empirical cumulative Kullback-Leibler information are presented and discussed. Equivalent forms of the new empirical measures are derived.

### 2. GENERAL FRAMEWORK

We consider two absolutely continuous, non-negative, random variables  $X$  and  $Y$ , with distribution functions denoted by  $F$  and  $G$ . Let  $X_1, X_2, \dots, X_n$  sample variables, independently and identically distributed as  $X$ , and  $Y_1, Y_2, \dots, Y_n$  sample variables, independently and identically distributed as  $Y$ .

We denote the empirical cumulative distribution function of  $X$ , and respectively  $Y$  by

$$\hat{F}_n(x) = \frac{1}{n} \sum_{i=1}^n \mathbf{I}_{\{X_i \leq x\}}, \quad (1)$$

$$\hat{G}_m(y) = \frac{1}{m} \sum_{j=1}^m \mathbf{I}_{\{Y_j \leq y\}}, \quad (2)$$

where  $x, y \in \mathbf{R}$  and  $\mathbf{I}_{\{X \leq x\}}$  is the indicator function

$$\mathbf{I}_{\{Y_j \leq y\}} = \begin{cases} 1, & \text{if } X \in [0, x] \\ 0, & \text{if } X \notin [0, x] \end{cases}. \quad (3)$$

As usual,  $\bar{X}_n$  and  $\bar{Y}_m$  are the sample means,  $X_{(1)} \leq X_{(2)} \leq \dots \leq X_{(n)}$  and  $Y_{(1)} \leq Y_{(2)} \leq \dots \leq Y_{(m)}$  the order statistics of the two samples. We denote by  $\Delta X_{(i)} = X_{(i+1)} - X_{(i)}$ , for  $i = \overline{1, n-1}$ .

Using the notations from [1], let  $N_j, j = \overline{1, m}$  be the number of random variables of the first sample that are less than or equal to the  $j$ -th order statistic of the second sample, that is  $N_j = \sum_{i=1}^n \mathbf{I}_{\{X_i \leq Y_{(j)}\}}$ . The random variables of the first sample belonging to  $(Y_{(j)}, Y_{(j+1)}]$  are denoted by  $X_{j,1} \leq X_{j,2} \leq \dots \leq X_{N_{j+1}-N_j}$  (if there are any). Also as in [1], the left-hand point and right-hand point of a random variable  $T$  with cumulative distribution function  $F_T$  are  $l_T = \inf\{t \in \mathbf{R} | F_T(t) > 0\}$  and  $r_T = \sup\{t \in \mathbf{R} | F_T(t) < 0\}$ .

Moreover, let  $w: [0, \infty) \rightarrow [0, \infty)$  be the weight function and  $W: [0, \infty) \rightarrow [0, \infty)$  a primitive of it.

### 3. WEIGHTED EMPIRICAL CUMULATIVE KULLBACK-LEIBLER INFORMATION

*Definition 3.1.* The weighted empirical cumulative Kullback-Leibler information of the random variables  $X$  and  $Y$  is:

$$C_{KL}^w(\hat{F}_n, \hat{G}_m) = \int_0^\infty w(x) \left( \hat{F}_n(x) \ln \frac{\hat{F}_n(x)}{\hat{G}_m(x)} - \hat{F}_n(x) + \hat{G}_m(x) \right) dx. \quad (4)$$

*Remark 3.1.* If  $w(x) = 1$  for every  $x$ , we get  $C_{KL}^w(\hat{F}_n, \hat{G}_m)$  as defined in Di Crescenzo and Longobardi [1].

**THEOREM 3.1.** *The weighted empirical cumulative Kullback-Leibler information of the random variables  $X$  and  $Y$  is expressed as follows:*

$$C_{KL}^w(\hat{F}_n, \hat{G}_m) = \frac{1}{n} \sum_{j=1}^{m-1} \left[ \ln \frac{j}{m} \cdot \left( \sum_{r=1}^{N_{j+1}-N_j} W(X_{j,r}) + N_j \cdot W(Y_{(j)}) - N_{j+1} \cdot W(Y_{(j+1)}) \right) \right] + \sum_{i=1}^{n-1} \left( \frac{i}{n} \cdot \ln \frac{i}{n} \cdot \Delta W(X_{(i)}) \right) + \overline{W(X)}_n - \overline{W(Y)}_m, \quad (5)$$

where  $\overline{W(X)}_n$  and  $\overline{W(Y)}_m$  are the sample means of the samples  $(W(X_i))_{i=\overline{1, n}}$  and respectively  $(W(Y_j))_{j=\overline{1, m}}$ .

*Proof.* According to Definition 3.1, the weighted empirical cumulative Kullback-Leibler information could be written as follows:

$$C_{KL}^w(\hat{F}_n, \hat{G}_m) = - \int_0^\infty w(x) \hat{F}_n(x) \ln(\hat{G}_m(x)) dx + \int_0^\infty w(x) \hat{F}_n(x) \ln(\hat{F}_n(x)) dx + \int_0^\infty w(x) [-\hat{F}_n(x) + \hat{G}_m(x)] dx. \quad (6)$$

For the first integral, we get:

$$\begin{aligned} \int_0^{\infty} w(x) \hat{F}_n(x) \ln(\hat{G}_m(x)) dx &= \sum_{j=1}^{m-1} \left( \ln \frac{j}{m} \cdot \int_{Y_{(j)}}^{Y_{(j+1)}} w(x) \hat{F}_n(x) dx \right) = \\ &= -\frac{1}{n} \sum_{j=1}^{m-1} \left[ \ln \frac{j}{m} \cdot \left( \sum_{r=1}^{N_{j+1}-N_j} W(X_{j,r}) + N_j \cdot W(Y_{(j)}) - N_{j+1} \cdot W(Y_{(j+1)}) \right) \right], \end{aligned} \quad (7)$$

using the definitions and notations presented in Section 2.

The second integral in (6) is

$$\begin{aligned} \int_0^{\infty} w(x) \hat{F}_n(x) \ln(\hat{F}_n(x)) dx &= \sum_{i=1}^{n-1} \left( \frac{i}{n} \cdot \ln \frac{i}{n} \cdot \int_{X_{(i)}}^{X_{(i+1)}} w(x) dx \right) = \\ &= \sum_{i=1}^{n-1} \left( \frac{i}{n} \cdot \ln \frac{i}{n} \cdot \Delta W(X_{(i)}) \right). \end{aligned} \quad (8)$$

Finally, straightforward calculation leads to

$$\int_0^{\infty} w(x) [-\hat{F}_n(x) + \hat{G}_m(x)] dx = \frac{1}{n} \sum_{i=1}^{n-1} W(X_{(i)}) - \frac{1}{m} \sum_{j=1}^{m-1} W(Y_{(j)}) = \overline{W(X)}_n - \overline{W(Y)}_m. \quad (9)$$

Using relations (7)–(9) in (6), we get relation (5).

*Remark 3.2.* As expected, for  $w(x)=1$  for every  $x$ , we get the result obtained by Di Crescenzo and Longobardi in [1].

*Remark 3.3.* Defining weighted empirical cumulative inaccuracy by

$$K^w(\hat{F}_n, \hat{G}_m) = - \int_0^{\infty} w(x) \hat{F}_n(x) \ln(\hat{G}_m(x)) dx \quad (10)$$

and weighted empirical cumulative entropy as

$$CE^w(\hat{F}_n) = - \int_0^{\infty} w(x) \hat{F}_n(x) \ln(\hat{F}_n(x)) dx, \quad (11)$$

we obtain

$$C_{KL}^w(\hat{F}_n, \hat{G}_m) = K^w(\hat{F}_n, \hat{G}_m) - CE^w(\hat{F}_n) + \overline{W(X)}_n - \overline{W(Y)}_m. \quad (12)$$

The weighted versions of cumulative Kullback-Leibler information of random variables  $X$  and  $Y$ , cumulative inaccuracy and cumulative entropy are defined in what it follows.

*Definition 3.2.* Let  $X$  and  $Y$  be random variables with the same left-hand points  $l = l_X = l_Y$  and with  $E(W(X))$  and  $E(W(Y))$  finite. The weighted cumulative Kullback-Leibler information of  $X$  and  $Y$  is

$$C_{KL}^w(X, Y) = \int_l^{\max\{r_X, r_Y\}} w(x) \left( F(x) \ln \frac{F(x)}{G(x)} - F(x) + G(x) \right) dx. \quad (13)$$

*Remark 3.4.* For  $w(x)=1$ ,  $\forall x$  in (13), we get the cumulative Kullback-Leibler information as defined in Park et al [8].

*Definition 3.3.* For any pair of random variables  $X$  and  $Y$  having the same left-hand points  $l$ , the weighted cumulative inaccuracy is defined by

$$K^w(X, Y) = - \int_l^{\max\{r_X, r_Y\}} w(x)F(x) \ln G(x) dx, \quad (14)$$

provided that the integral is finite.

The weighted cumulative entropy of  $X$  is

$$CE^w(X) = - \int_0^{\infty} w(x)F(x) \ln F(x) dx. \quad (15)$$

*Remark 3.5.* Interesting results were obtained by F. Misagh in [7] for  $CE^w(X)$  and  $CE^w(\hat{F}_n)$  for the particular case  $w(x)=1$ .

*Remark 3.6.* Based on definitions 3.2 and 3.3, we get

$$C_{KL}^w(X, Y) = K^w(X, Y) - CE^w(X) + E(W(X)) - E(W(Y)). \quad (16)$$

*Numerical application.* Let  $X$  and  $Y$  be two continuous, nonnegative, random variables. The distributions taken into account for the variables, the weight functions considered and the theoretical weighted cumulative Kullback-Leibler information are presented in the table 3.1.

We conducted a simulation study for evaluating the weighted empirical cumulative Kullback-Leibler information (based on Theorem 3.1), considering a sample of size  $n = 1500$  for random variable  $X$  and  $m = 1000$  for random variable  $Y$ . The process was repeated 1000 times. The mean squared errors (MSEs) between average weighted empirical cumulative Kullback-Leibler information and its theoretical correspondent are also presented in Table 3.1.

Table 3.1

Weighted and empirical weighted cumulative Kullback-Leibler information

Distributions of $X$ and $Y$	Weight function $w(x)$	$C_{KL}^w(X, Y)$	Average $C_{KL}^w(\hat{F}_n, \hat{G}_m)$	Average MSE
$X$ : inverse Weibull ( $\theta = 1, \tau = 4$ ) $Y$ : inverse Weibull ( $\theta = 0.5, \tau = 4$ )	$w(x)=1$	0.3255	0.3257	0.000060
	$w(x)=x$	0.2493	0.2494	0.000046
$X$ : Power ( $\alpha = 6$ ) $Y$ : Power ( $\beta = 2$ )	$w(x)=1$	0.1088	0.1091	0.000046
	$w(x)=x$	0.0625	0.0627	0.000015
	$w(x)=1-x$	0.0463	0.0465	0.000009
$X$ : exponential ( $\lambda = 1$ ) $Y$ : exponential ( $\lambda = 2$ )	$w(x)=1$	0.0638	0.0641	0.000061
	$w(x)=x$	0.0572	0.0574	0.000050
	$w(x)=1 - e^{-x}$	0.0334	0.0336	0.000017

#### 4. TSALLIS EMPIRICAL CUMULATIVE KULLBACK-LEIBLER INFORMATION

The Tsallis extended logarithm is defined as follows for  $x \in \mathbb{R}_+^*$  and  $q \in \mathbb{R}$

$$\ln_q^t(x) = \begin{cases} \ln x & , \text{ if } x > 0 \text{ and } q = 1 \\ \frac{x^{1-q} - 1}{1-q} & , \text{ if } x > 0 \text{ and } q \neq 1 \end{cases} \quad (17)$$

*Remark 4.1.* The following property of Tsallis extended logarithm will be used for proving some of the

results in Sections 4 and 5:

$$\ln_q^t\left(\frac{x}{y}\right) = \ln_q^t(x) - x^{1-q} \cdot \ln_{2-q}^t(y). \quad (18)$$

*Definition 4.1.* The Tsallis empirical cumulative Kullback-Leibler information of the random variables  $X$  and  $Y$  is defined as follows:

$$C_{KL}^{t,q}(\hat{F}_n, \hat{G}_m) = \int_0^\infty \hat{F}_n(x) \cdot \ln_q^t\left(\frac{\hat{F}_n(x)}{\hat{G}_m(x)}\right) dx + \bar{X}_n - \bar{Y}_m. \quad (19)$$

*Remark 4.2.* The Tsallis empirical cumulative Kullback-Leibler information generalizes the empirical cumulative Kullback-Leibler information (obtained from the former for  $q=1$ ).

**THEOREM 4.1.** *The following relation holds for Tsallis empirical cumulative Kullback-Leibler information of the random variables  $X$  and  $Y$ :*

$$\begin{aligned} C_{KL}^{t,q}(\hat{F}_n, \hat{G}_m) &= \\ &= \frac{1}{n^{2-q}} \sum_{j=1}^{m-1} \ln_{2-q}^t\left(\frac{j}{m}\right) \left[ \sum_{r=1}^{N_{j+1}-N_j} \left( (N_j+r)^{2-q} - (N_j+r-1)^{2-q} \right) \cdot X_{j,r} + N_j^{2-q} \cdot Y_{(j)} - N_{j+1}^{2-q} \cdot Y_{(j+1)} \right] + \\ &\quad + \sum_{i=1}^{n-1} \frac{i}{n} \cdot \ln_q^t\left(\frac{i}{n}\right) \cdot \Delta X_{(i)} + \bar{X}_n - \bar{Y}_m. \end{aligned} \quad (20)$$

*Proof.* Using Remark 4.1 we get:

$$C_{KL}^{t,q}(\hat{F}_n, \hat{G}_m) = - \int_0^\infty (\hat{F}_n(x))^{2-q} \cdot \ln_{2-q}^t(\hat{G}_m(x)) dx + \int_0^\infty \hat{F}_n(x) \cdot \ln_q^t(\hat{F}_n(x)) dx + \bar{X}_n - \bar{Y}_m. \quad (21)$$

Since

$$\begin{aligned} &\int_0^\infty (\hat{F}_n(x))^{2-q} \cdot \ln_{2-q}^t(\hat{G}_m(x)) dx = \\ &= \sum_{j=1}^{m-1} \left[ \ln_{2-q}^t\left(\frac{j}{m}\right) \cdot \int_{Y_{(j)}}^{Y_{(j+1)}} (\hat{F}_n(x))^{2-q} dx \right] = \\ &= - \frac{1}{n^{2-q}} \sum_{j=1}^{m-1} \ln_{2-q}^t\left(\frac{j}{m}\right) \left[ \sum_{r=1}^{N_{j+1}-N_j} \left( (N_j+r)^{2-q} - (N_j+r-1)^{2-q} \right) \cdot X_{j,r} + N_j^{2-q} \cdot Y_{(j)} - N_{j+1}^{2-q} \cdot Y_{(j+1)} \right] \end{aligned} \quad (22)$$

and

$$\int_0^\infty \hat{F}_n(x) \cdot \ln_q^t(\hat{F}_n(x)) dx = \sum_{i=1}^{n-1} \left[ \int_{X_{(i)}}^{X_{(i+1)}} \frac{i}{n} \cdot \ln_q^t\left(\frac{i}{n}\right) dx \right] = \sum_{i=1}^{n-1} \frac{i}{n} \cdot \ln_q^t\left(\frac{i}{n}\right) \cdot \Delta X_{(i)}, \quad (23)$$

relation (20) is obtained by taking into account (22) and (23) in (21).

*Remark 4.3.* The Tsallis extended logarithm versions of empirical cumulative entropy and empirical cumulative inaccuracy are defined as follows:

– Tsallis empirical cumulative entropy of random variable  $X$

$$CE^{t,q}(\hat{F}_n) = - \int_0^{\infty} \hat{F}_n(x) \cdot \ln_q^t(\hat{F}_n(x)) dx \quad (24)$$

and respectively,

– Tsallis empirical cumulative inaccuracy of random variables  $X$  and  $Y$

$$K^{t,q}(\hat{F}_n, \hat{G}_m) = - \int_0^{\infty} (\hat{F}_n(x))^{2-q} \cdot \ln_{2-q}^t(\hat{G}_m(x)) dx. \quad (25)$$

From (21), (24) and (25), we get

$$C_{KL}^{t,q}(\hat{F}_n, \hat{G}_m) = K^{t,q}(\hat{F}_n, \hat{G}_m) - CE^{t,q}(\hat{F}_n) + \bar{X}_n - \bar{Y}_m. \quad (26)$$

The following definition generalizes the cumulative Kullback-Leibler information of random variables  $X$  and  $Y$ , cumulative inaccuracy and cumulative entropy.

*Definition 4.2.* Let  $X$  and  $Y$  be random variables with finite expectations and with  $l = l_X = l_Y$ .

The Tsallis cumulative Kullback-Leibler information of  $X$  and  $Y$  is given by

$$C_{KL}^{t,q}(X, Y) = \int_l^{\max\{r_X, r_Y\}} F(x) \cdot \ln_q^t\left(\frac{F(x)}{G(x)}\right) dx + E(X) - E(Y). \quad (27)$$

Tsallis cumulative inaccuracy of random variables  $X$  and  $Y$  is

$$K^{t,q}(X, Y) = - \int_l^{\max\{r_X, r_Y\}} (F(x))^{2-q} \ln_{2-q}^t(G(x)) dx, \quad (28)$$

provided that the integral is finite.

Similarly, Tsallis cumulative entropy is defined as

$$CE^{t,q}(X) = - \int_0^{\infty} F(x) \cdot \ln_q^t(F(x)) dx. \quad (29)$$

*Remark 4.4.* Using the property of Tsallis extended logarithm presented in Remark 4.1, one can derive

$$C_{KL}^{t,q}(X, Y) = K^{t,q}(X, Y) - CE^{t,q}(X) + E(X) - E(Y). \quad (30)$$

Numerical application of information measures discussed in this section (Tsallis cumulative Kullback-Leibler information and its empirical correspondent) will be presented at the end of Section 5, together with their weighted versions.

## 5. TSALLIS WEIGHTED EMPIRICAL CUMULATIVE KULLBACK-LEIBLER INFORMATION

In this section, the measures and results from Section 4 will be extended by considering the weighted cases. In the same time, Section 5 extends the results of Section 3 by considering Tsallis extended logarithm.

*Definition 5.1.* The Tsallis weighted empirical cumulative Kullback-Leibler information of the random variables  $X$  and  $Y$  is defined as follows:

$$C_{KL}^{w,t,q}(\hat{F}_n, \hat{G}_m) = \int_0^{\infty} w(x) \left[ \hat{F}_n(x) \cdot \ln_q^t\left(\frac{\hat{F}_n(x)}{\hat{G}_m(x)}\right) - \hat{F}_n(x) + \hat{G}_m(x) \right] dx. \quad (31)$$

Moreover, we define Tsallis weighted empirical cumulative inaccuracy as

$$K^{w,t,q}(\hat{F}_n, \hat{G}_m) = - \int_0^{\infty} w(x) (\hat{F}_n(x))^{2-q} \ln_{2-q}^t(\hat{G}_m(x)) dx \quad (32)$$

and Tsallis weighted empirical cumulative entropy

$$CE^{w,t,q}(\hat{F}_n) = - \int_0^{\infty} w(x) \cdot \hat{F}_n(x) \cdot \ln_q^t(\hat{F}_n(x)) dx. \quad (33)$$

*Remark 5.1.* Straightforward calculations lead to

$$C_{KL}^{w,t,q}(\hat{F}_n, \hat{G}_m) = K^{w,t,q}(\hat{F}_n, \hat{G}_m) - CE^{w,t,q}(\hat{F}_n) + \overline{W(X)}_n - \overline{W(Y)}_m. \quad (34)$$

**THEOREM 5.1.** *The following result holds true for Tsallis weighted empirical cumulative Kullback-Leibler information of the random variables  $X$  and  $Y$ :*

$$\begin{aligned} C_{KL}^{w,t,q}(\hat{F}_n, \hat{G}_m) &= \\ &= \frac{1}{n^{2-q}} \sum_{j=1}^{m-1} \ln_{2-q}^t\left(\frac{j}{m}\right) \left[ \sum_{r=1}^{\lceil N_{j+1} - N_j \rceil} ((N_j + r)^{2-q} - (N_j + r - 1)^{2-q}) \cdot W(X_{j,r}) + N_j^{2-q} \cdot W(Y_{(j)}) - N_{j+1}^{2-q} \cdot W(Y_{(j+1)}) \right] + \\ &\quad + \sum_{i=1}^{n-1} \frac{i}{n} \cdot \ln_q^t\left(\frac{i}{n}\right) \cdot \Delta W(X_{(i)}) + \overline{W(X)}_n - \overline{W(Y)}_m. \end{aligned} \quad (35)$$

*Proof.* Applying similar arguments used for proving Theorems 3.1 and 4.1, we get relation (35).

The extended (weighted) versions of Tsallis cumulative Kullback-Leibler information, Tsallis cumulative inaccuracy and Tsallis cumulative entropy are defined in what it follows.

*Definition 5.2.* Let  $X$  and  $Y$  be random variables having the same left-hand points  $l = l_X = l_Y$  and  $E(W(X))$  and  $E(W(Y))$  finite.

The Tsallis weighted cumulative Kullback-Leibler information of  $X$  and  $Y$  is given by

$$C_{KL}^{w,t,q}(X, Y) = \int_l^{\max\{r_X, r_Y\}} w(x) \cdot \left[ F(x) \cdot \ln_q^t\left(\frac{F(x)}{G(x)}\right) - F(x) + G(x) \right] dx. \quad (36)$$

The Tsallis weighted cumulative inaccuracy of random variables  $X$  and  $Y$  is

$$K^{w,t,q}(X, Y) = - \int_l^{\max\{r_X, r_Y\}} w(x) (F(x))^{2-q} \cdot \ln_{2-q}^t(G(x)) dx, \quad (37)$$

provided that the integral is finite, and Tsallis weighted cumulative entropy is defined as

$$CE^{w,t,q}(X) = - \int_0^{\infty} w(x) F(x) \cdot \ln_q^t(F(x)) dx. \quad (38)$$

*Remark 5.2.* Using relation (18), it can be shown that

$$C_{KL}^{w,t,q}(X, Y) = K^{w,t,q}(X, Y) - CE^{w,t,q}(X) + E(W(X)) - E(W(Y)). \quad (39)$$

*Numerical application.* As in numerical example from Section 3, we take  $X$  and  $Y$  two continuous, nonnegative, random variables with different distributions, for which we evaluate the Tsallis weighted and non-weighted cumulative Kullback-Leibler information and their empirical versions. For the empirical information, we considered a sample of size  $n = 1500$  for random variable  $X$  and  $m = 1000$  for random variable  $Y$  and repeated the calculation process 1000 times. The results are presented in Table 5.1 ( $q = 0.5$  in case of Power distribution and  $q = 1.5$  for exponential distribution).

Table 5.1

Tsallis weighted and empirical weighted cumulative Kullback-Leibler information

Distributions of $X$ and $Y$	Weight function $w(x)$	$C_{KL}^{w,t,q}(X, Y)$	Average $C_{KL}^{w,t,q}(\hat{F}_n, \hat{G}_m)$	AverageM SE
X: Power ( $\alpha = 6$ ) Y: Power ( $\beta = 4$ )	$w(x)=1$	0.0214	0.0215	0.000014
	$w(x)=x$	0.0139	0.0139	0.000006
	$w(x) = 1 - e^{-x}$	0.0101	0.0102	0.000003

Table 5.1 (continued)

X: exponential ( $\lambda = 1$ ) Y: exponential ( $\lambda = 2$ )	$w(x)=1$	0.0902	0.0905	0.000114
	$w(x)=x$	0.0822	0.0821	0.000095
	$w(x) = 1 - e^{-x}$	0.0477	0.0478	0.000031

## 6. CONCLUSIONS

In this paper we proposed some generalized versions of the cumulative Kullback Leibler information. Based on the work of A. di Crescenzo and M. Longobardi [1] we defined the weighted cumulative Kullback-Leibler information, the Tsallis cumulative Kullback-Leibler information and the Tsallis weighted cumulative Kullback-Leibler information and their empirical versions.

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